Miniconference on dependence and ecology Paris, IHP, 2023 03 15

# Scaling limits of nonlinear functions of random grain model with application to Burgers' equation

Donatas Surgailis (Vilnius University)

- Spatial long-range dependence (LRD) and limit theorems (scaling limits)
- Scaling of random grain (RG) model

- Spatial long-range dependence (LRD) and limit theorems (scaling limits)
- Scaling of random grain (RG) model
- Oharlier polynomials & Mehler's formula

- Spatial long-range dependence (LRD) and limit theorems (scaling limits)
- Scaling of random grain (RG) model
- Ocharlier polynomials & Mehler's formula
- Scaling of nonlinear functions of RG model

- Spatial long-range dependence (LRD) and limit theorems (scaling limits)
- Scaling of random grain (RG) model
- Ocharlier polynomials & Mehler's formula
- Scaling of nonlinear functions of RG model
- Application to Burgers' equation

- Spatial long-range dependence (LRD) and limit theorems (scaling limits)
- Scaling of random grain (RG) model
- Oharlier polynomials & Mehler's formula
- Scaling of nonlinear functions of RG model
- Application to Burgers' equation
- O Perspectives & open questions

Spatial process - stationary random field (RF) X = {X(t); t ∈ ℝ<sup>d</sup>} or X = {X(t); t ∈ ℝ<sup>d</sup>} with covariance r<sub>X</sub>(t) := Cov(X(0), X(t))

- Spatial process stationary random field (RF) X = {X(t); t ∈ ℝ<sup>d</sup>} or X = {X(t); t ∈ ℝ<sup>d</sup>} with covariance r<sub>X</sub>(t) := Cov(X(0), X(t))
- LRD:  $r_X$  nonintegrable (nonsummable):  $\int_{\mathbb{R}^d} |r_X(t)| dt = \infty$  or  $\sum_{t \in \mathbb{Z}^d} |r_X(t)| = \infty$

- Spatial process stationary random field (RF) X = {X(t); t ∈ ℝ<sup>d</sup>} or X = {X(t); t ∈ ℤ<sup>d</sup>} with covariance r<sub>X</sub>(t) := Cov(X(0), X(t))
- LRD:  $r_X$  nonintegrable (nonsummable):  $\int_{\mathbb{R}^d} |r_X(t)| dt = \infty$  or  $\sum_{t \in \mathbb{Z}^d} |r_X(t)| = \infty$
- Great variety of spatial LRD models, limit theorems, limit distributions (scaling limits)...

- Spatial process stationary random field (RF)  $X = \{X(t); t \in \mathbb{R}^d\}$  or  $X = \{X(t); t \in \mathbb{Z}^d\}$  with covariance  $r_X(t) := \text{Cov}(X(0), X(t))$
- LRD:  $r_X$  nonintegrable (nonsummable):  $\int_{\mathbb{R}^d} |r_X(t)| dt = \infty$  or  $\sum_{t \in \mathbb{Z}^d} |r_X(t)| = \infty$
- Great variety of spatial LRD models, limit theorems, limit distributions (scaling limits)...
- Scaling (zooming out): getting a distant view of the object

- Spatial process stationary random field (RF)  $X = \{X(t); t \in \mathbb{R}^d\}$  or  $X = \{X(t); t \in \mathbb{Z}^d\}$  with covariance  $r_X(t) := \text{Cov}(X(0), X(t))$
- LRD:  $r_X$  nonintegrable (nonsummable):  $\int_{\mathbb{R}^d} |r_X(t)| dt = \infty$  or  $\sum_{t \in \mathbb{Z}^d} |r_X(t)| = \infty$
- Great variety of spatial LRD models, limit theorems, limit distributions (scaling limits)...
- Scaling (zooming out): getting a distant view of the object
- At large scales, short-range details ('dependences', 'correlations') disappear but long-range 'correlations' may prevail

- Spatial process stationary random field (RF)  $X = \{X(t); t \in \mathbb{R}^d\}$  or  $X = \{X(t); t \in \mathbb{Z}^d\}$  with covariance  $r_X(t) := \text{Cov}(X(0), X(t))$
- LRD:  $r_X$  nonintegrable (nonsummable):  $\int_{\mathbb{R}^d} |r_X(t)| dt = \infty$  or  $\sum_{t \in \mathbb{Z}^d} |r_X(t)| = \infty$
- Great variety of spatial LRD models, limit theorems, limit distributions (scaling limits)...
- Scaling (zooming out): getting a distant view of the object
- At large scales, short-range details ('dependences', 'correlations') disappear but long-range 'correlations' may prevail
- Scaling (partial sums) limits of any *weakly dependent* 2nd order process X coincide with Brownian motion (Donsker's theorem)

- Spatial process stationary random field (RF)  $X = \{X(t); t \in \mathbb{R}^d\}$  or  $X = \{X(t); t \in \mathbb{Z}^d\}$  with covariance  $r_X(t) := \text{Cov}(X(0), X(t))$
- LRD:  $r_X$  nonintegrable (nonsummable):  $\int_{\mathbb{R}^d} |r_X(t)| dt = \infty$  or  $\sum_{t \in \mathbb{Z}^d} |r_X(t)| = \infty$
- Great variety of spatial LRD models, limit theorems, limit distributions (scaling limits)...
- Scaling (zooming out): getting a distant view of the object
- At large scales, short-range details ('dependences', 'correlations') disappear but long-range 'correlations' may prevail
- Scaling (partial sums) limits of any *weakly dependent* 2nd order process X coincide with Brownian motion (Donsker's theorem)

• Scaling limit of a stationary process X is self-similar (Lamperti, 1962) and provides a 'large-scale summary of dependence structure of X'

- Scaling limit of a stationary process X is self-similar (Lamperti, 1962) and provides a 'large-scale summary of dependence structure of X'
- We (and many other works on scaling limits) consider the limit distribution of integrals:

$$X_{\lambda}(\phi) := \int_{\mathbb{R}^d} X(t) \phi(t/\lambda) \mathrm{d}t, \quad \text{ as } \lambda \to \infty,$$
 (1)

(or respective sums in the discrete argument case), where  $X = \{X(t); t \in \mathbb{R}^d\}$  is a given stationary RF, for each  $\phi$  from a class of (test) functions  $\Phi = \{\phi : \mathbb{R}^d \to \mathbb{R}\}$ .

- Scaling limit of a stationary process X is self-similar (Lamperti, 1962) and provides a 'large-scale summary of dependence structure of X'
- We (and many other works on scaling limits) consider the limit distribution of integrals:

$$X_{\lambda}(\phi) := \int_{\mathbb{R}^d} X(t) \phi(t/\lambda) \mathrm{d}t, \quad \text{ as } \lambda \to \infty,$$
 (1)

(or respective sums in the discrete argument case), where  $X = \{X(t); t \in \mathbb{R}^d\}$  is a given stationary RF, for each  $\phi$  from a class of (test) functions  $\Phi = \{\phi : \mathbb{R}^d \to \mathbb{R}\}$ .

A suitably normalized limit

$$d_{\lambda}^{-1}(X_{\lambda}(\phi) - \mathbb{E}X_{\lambda}(\phi)) \stackrel{\mathrm{d}}{\longrightarrow} V(\phi), \quad \lambda \to \infty$$
 (2)

is a RF  $V(\phi)$  indexed by  $\phi \in \Phi$  is called the (isotropic) scaling limit of X

- Scaling limit of a stationary process X is self-similar (Lamperti, 1962) and provides a 'large-scale summary of dependence structure of X'
- We (and many other works on scaling limits) consider the limit distribution of integrals:

$$X_{\lambda}(\phi) := \int_{\mathbb{R}^d} X(t) \phi(t/\lambda) \mathrm{d}t, \quad \text{ as } \lambda \to \infty,$$
 (1)

(or respective sums in the discrete argument case), where  $X = \{X(t); t \in \mathbb{R}^d\}$  is a given stationary RF, for each  $\phi$  from a class of (test) functions  $\Phi = \{\phi : \mathbb{R}^d \to \mathbb{R}\}$ .

A suitably normalized limit

$$d_{\lambda}^{-1}(X_{\lambda}(\phi) - \mathbb{E}X_{\lambda}(\phi)) \stackrel{\mathrm{d}}{\longrightarrow} V(\phi), \quad \lambda \to \infty$$
 (2)

is a RF  $V(\phi)$  indexed by  $\phi \in \Phi$  is called the (isotropic) scaling limit of X

 The above approach is common in the theory of generalized RFs Gel'fand, I.M., Vilenkin, N.Ya. (1964) Generalized Functions - Vol.4: Applications of Harmonic Analysis Dobrushin, R.L. (1980) Automodel generalized random fields and their renormgroup. In: R.L. Dobrushin and Ya.G. Sinai (Eds.), Multicomponent Random Systems

• The limit in (2) strongly depends on the class  $\Phi$  of test functions

- The limit in (2) strongly depends on the class  $\Phi$  of test functions
- In the theory of generalized RFs ('random Schwartz distributions') Φ is Schwartz space D(R<sup>d</sup>) or S(R<sup>d</sup>) of very smooth (infinitely differentiable) functions, which is justified by applications in mathematical physics (quantum field theory)

- The limit in (2) strongly depends on the class  $\Phi$  of test functions
- In the theory of generalized RFs ('random Schwartz distributions') Φ is Schwartz space D(R<sup>d</sup>) or S(R<sup>d</sup>) of very smooth (infinitely differentiable) functions, which is justified by applications in mathematical physics (quantum field theory)
- In spatial statistics,  $\Phi = \{\phi\}$  may consist of indicator functions

 $\phi(\mathbf{t}) = \mathbb{I}(\mathbf{t} \in A), \qquad \mathbf{t} \in \mathbb{R}^d$ 

where A runs over a class  $\mathcal{A}$  of Borel subsets of  $\mathbb{R}^d$ 

- The limit in (2) strongly depends on the class  $\Phi$  of test functions
- In the theory of generalized RFs ('random Schwartz distributions') Φ is Schwartz space D(R<sup>d</sup>) or S(R<sup>d</sup>) of very smooth (infinitely differentiable) functions, which is justified by applications in mathematical physics (quantum field theory)
- In spatial statistics,  $\Phi = \{\phi\}$  may consist of indicator functions

$$\phi(t) = \mathbb{I}(t \in A), \qquad t \in \mathbb{R}^{d}$$

where A runs over a class  $\mathcal{A}$  of Borel subsets of  $\mathbb{R}^d$ 

• For above indicator function  $\phi$ ,

$$X_{\lambda}(\phi) = \int_{\boldsymbol{t} \in \lambda A} X(\boldsymbol{t}) \mathrm{d}\boldsymbol{t} \quad \text{or} \quad X_{\lambda}(\phi) = \sum_{\boldsymbol{t} \in \lambda A \cap \mathbb{Z}^d} X(\boldsymbol{t})$$
(3)

is the empirical mean of X (times  $Leb_d(\lambda A) = \lambda^d Leb_d(A)$ )

- The limit in (2) strongly depends on the class  $\Phi$  of test functions
- In the theory of generalized RFs ('random Schwartz distributions') Φ is Schwartz space D(ℝ<sup>d</sup>) or S(ℝ<sup>d</sup>) of very smooth (infinitely differentiable) functions, which is justified by applications in mathematical physics (quantum field theory)
- In spatial statistics,  $\Phi = \{\phi\}$  may consist of indicator functions

$$\phi(t) = \mathbb{I}(t \in A), \qquad t \in \mathbb{R}^{d}$$

where A runs over a class  $\mathcal A$  of Borel subsets of  $\mathbb R^d$ 

• For above indicator function  $\phi$ ,

$$X_{\lambda}(\phi) = \int_{\boldsymbol{t} \in \lambda A} X(\boldsymbol{t}) \mathrm{d}\boldsymbol{t} \quad \text{or} \quad X_{\lambda}(\phi) = \sum_{\boldsymbol{t} \in \lambda A \cap \mathbb{Z}^d} X(\boldsymbol{t})$$
(3)

is the empirical mean of X (times  $Leb_d(\lambda A) = \lambda^d Leb_d(A)$ ) given observations of X over 'inflated' set  $\lambda A$  whose limit distribution in (2) is of interest

- The limit in (2) strongly depends on the class  $\Phi$  of test functions
- In the theory of generalized RFs ('random Schwartz distributions') Φ is Schwartz space D(ℝ<sup>d</sup>) or S(ℝ<sup>d</sup>) of very smooth (infinitely differentiable) functions, which is justified by applications in mathematical physics (quantum field theory)
- In spatial statistics,  $\Phi = \{\phi\}$  may consist of indicator functions

$$\phi(t) = \mathbb{I}(t \in A), \qquad t \in \mathbb{R}^{d}$$

where A runs over a class  $\mathcal{A}$  of Borel subsets of  $\mathbb{R}^d$ 

• For above indicator function  $\phi$ ,

$$X_{\lambda}(\phi) = \int_{\boldsymbol{t} \in \lambda A} X(\boldsymbol{t}) \mathrm{d}\boldsymbol{t} \quad \text{or} \quad X_{\lambda}(\phi) = \sum_{\boldsymbol{t} \in \lambda A \cap \mathbb{Z}^d} X(\boldsymbol{t})$$
(3)

is the empirical mean of X (times  $Leb_d(\lambda A) = \lambda^d Leb_d(A)$ ) given observations of X over 'inflated' set  $\lambda A$  whose limit distribution in (2) is of interest

• The limit distribution of empirical mean in (3) may be difficult if A has irregular boundary ('edge effects')

Lahiri, S.N. and Robinson, P.M. (2016) Central limit theorems for long range dependent spatial linear processes.

Bernoulli 22, 345-375

Popular approach in limit theorems for RFs: integration/summation on rectangles:

$$\Phi_{\mathrm{rec},d} := \{ \mathbb{I}(oldsymbol{t} \in ]oldsymbol{0}, oldsymbol{s}]); oldsymbol{s} \in \mathbb{R}^d_+ \}, \qquad ]oldsymbol{0}, oldsymbol{s}] := \prod_{i=1}^a ]oldsymbol{0}, oldsymbol{s}_i]$$

Then  $X_{\lambda}(s) = \sum_{t \in ]0, \lambda s]} X(t)$  is a RF indexed by points  $s \in \mathbb{R}^{d}_{+}$ *d*-dimensional analog of the partial sums process of time series

Popular approach in limit theorems for RFs: integration/summation on rectangles:

$$\Phi_{\mathrm{rec},d}:=\{\mathbb{I}(oldsymbol{t}\in]oldsymbol{0},oldsymbol{s}]);oldsymbol{s}\in\mathbb{R}^d_+\},\qquad ]oldsymbol{0},oldsymbol{s}]:=\prod_{i=1}^d]0,oldsymbol{s}_i]$$

Then  $X_{\lambda}(s) = \sum_{t \in [0,\lambda s]} X(t)$  is a RF indexed by points  $s \in \mathbb{R}^d_+$ *d*-dimensional analog of the partial sums process of time series

 Weak dependent X under mixing conditions ('rectangles' = 'blocks'): Dedecker, J., Doukhan, P., Lang, G., León, J.R., Louhichi, S. and Prieur, C. (2007) Weak Dependendence. With Examples and Applications. Lecture Notes Statist. vol. 190. Springer

Popular approach in limit theorems for RFs: integration/summation on rectangles:

$$\Phi_{\mathrm{rec},d}:=\{\mathbb{I}(oldsymbol{t}\in]oldsymbol{0},oldsymbol{s}]);oldsymbol{s}\in\mathbb{R}^d_+\},\qquad ]oldsymbol{0},oldsymbol{s}]:=\prod_{i=1}^d]0,oldsymbol{s}_i]$$

Then  $X_{\lambda}(s) = \sum_{t \in [0,\lambda s]} X(t)$  is a RF indexed by points  $s \in \mathbb{R}^{d}_{+}$ *d*-dimensional analog of the partial sums process of time series

- Weak dependent X under mixing conditions ('rectangles' = 'blocks'): Dedecker, J., Doukhan, P., Lang, G., León, J.R., Louhichi, S. and Prieur, C. (2007) Weak Dependendence. With Examples and Applications. Lecture Notes Statist. vol. 190. Springer
- Functional convergence & tightness ignored in this talk

Popular approach in limit theorems for RFs: integration/summation on rectangles:

$$oldsymbol{\Phi}_{\mathrm{rec},d}:=\{\mathbb{I}(oldsymbol{t}\in]oldsymbol{0},oldsymbol{s}]);oldsymbol{s}\in\mathbb{R}^d_+\},\qquad ]oldsymbol{0},oldsymbol{s}]:=\prod_{i=1}^d]0,s_i]$$

Then  $X_{\lambda}(s) = \sum_{t \in [0,\lambda s]} X(t)$  is a RF indexed by points  $s \in \mathbb{R}^d_+$ *d*-dimensional analog of the partial sums process of time series

- Weak dependent X under mixing conditions ('rectangles' = 'blocks'): Dedecker, J., Doukhan, P., Lang, G., León, J.R., Louhichi, S. and Prieur, C. (2007) Weak Dependendence. With Examples and Applications. Lecture Notes Statist. vol. 190. Springer
- Functional convergence & tightness ignored in this talk
- Isotropic or uniform scaling t → t/λ in (1) can be replaced by anisotropic or operator scaling t → λ<sup>-Γ</sup>t where Γ is a d × d-matrix, particularly, a diagonal matrix

$$\Gamma = \operatorname{diag}(\gamma_1, \cdots, \gamma_d), \qquad (\gamma_1, \cdots, \gamma_d) = \gamma \in \mathbb{R}^d_+$$

Popular approach in limit theorems for RFs: integration/summation on rectangles:

$$oldsymbol{\Phi}_{\mathrm{rec},d}:=\{\mathbb{I}(oldsymbol{t}\in]oldsymbol{0},oldsymbol{s}]);oldsymbol{s}\in\mathbb{R}^d_+\},\qquad ]oldsymbol{0},oldsymbol{s}]:=\prod_{i=1}^d]0,s_i]$$

Then  $X_{\lambda}(s) = \sum_{t \in [0,\lambda s]} X(t)$  is a RF indexed by points  $s \in \mathbb{R}^d_+$ *d*-dimensional analog of the partial sums process of time series

- Weak dependent X under mixing conditions ('rectangles' = 'blocks'): Dedecker, J., Doukhan, P., Lang, G., León, J.R., Louhichi, S. and Prieur, C. (2007) Weak Dependendence. With Examples and Applications. Lecture Notes Statist. vol. 190. Springer
- Functional convergence & tightness ignored in this talk
- Isotropic or uniform scaling t → t/λ in (1) can be replaced by anisotropic or operator scaling t → λ<sup>-Γ</sup>t where Γ is a d × d-matrix, particularly, a diagonal matrix

$$\Gamma = \operatorname{diag}(\gamma_1, \cdots, \gamma_d), \qquad (\gamma_1, \cdots, \gamma_d) = \gamma \in \mathbb{R}^d_+$$

• Popular approach in limit theorems for RFs: integration/summation on rectangles:

$$oldsymbol{\Phi}_{\mathrm{rec},d}:=\{\mathbb{I}(oldsymbol{t}\in]oldsymbol{0},oldsymbol{s}]);oldsymbol{s}\in\mathbb{R}^d_+\},\qquad ]oldsymbol{0},oldsymbol{s}]:=\prod_{i=1}^d]0,s_i]$$

Then  $X_{\lambda}(s) = \sum_{t \in [0,\lambda s]} X(t)$  is a RF indexed by points  $s \in \mathbb{R}^d_+$ *d*-dimensional analog of the partial sums process of time series

- Weak dependent X under mixing conditions ('rectangles' = 'blocks'): Dedecker, J., Doukhan, P., Lang, G., León, J.R., Louhichi, S. and Prieur, C. (2007) Weak Dependendence. With Examples and Applications. Lecture Notes Statist. vol. 190. Springer
- Functional convergence & tightness ignored in this talk
- Isotropic or uniform scaling t → t/λ in (1) can be replaced by anisotropic or operator scaling t → λ<sup>-Γ</sup>t where Γ is a d × d-matrix, particularly, a diagonal matrix

$$\Gamma = \operatorname{diag}(\gamma_1, \cdots, \gamma_d), \qquad (\gamma_1, \cdots, \gamma_d) = \gamma \in \mathbb{R}^d_+,$$

Operator scaling RF (OSRF):

Biermé, H., Meerschaert, M.M. and Scheffler, H.P. (2007) Operator scaling stable random fields. Stoch. Process. Appl.

We can consider the limit distribution of RF X<sub>λ,Γ</sub>(φ) = ∑<sub>t∈Z<sup>d</sup></sub> X(t)φ(λ<sup>-Γ</sup>t) or the anisotropically rescaled partial sums RF:

$$X_{\lambda,\gamma}(\boldsymbol{s}) = \sum_{\boldsymbol{t} \in [\boldsymbol{0},\lambda^{\Gamma}\boldsymbol{s}]} X(\boldsymbol{t}), \qquad \lambda^{\Gamma}\boldsymbol{s} = (\lambda^{\gamma_1} s_1, \cdots, \lambda^{\gamma_d} s_d)$$
(4)
We can consider the limit distribution of RF X<sub>λ,Γ</sub>(φ) = ∑<sub>t∈Z<sup>d</sup></sub> X(t)φ(λ<sup>-Γ</sup>t) or the anisotropically rescaled partial sums RF:

$$X_{\lambda,\gamma}(\boldsymbol{s}) = \sum_{\boldsymbol{t} \in [\boldsymbol{0},\lambda^{\Gamma}\boldsymbol{s}]} X(\boldsymbol{t}), \qquad \lambda^{\Gamma}\boldsymbol{s} = (\lambda^{\gamma_1} s_1, \cdots, \lambda^{\gamma_d} s_d)$$
(4)

• Rectangle  $]\mathbf{0}, \lambda^{\Gamma} \mathbf{s}]$  grow at different rate  $\lambda^{\gamma_j}$  in *j*th direction

$$X_{\lambda,\gamma}(\boldsymbol{s}) = \sum_{\boldsymbol{t} \in [\boldsymbol{0},\lambda^{\Gamma}\boldsymbol{s}]} X(\boldsymbol{t}), \qquad \lambda^{\Gamma}\boldsymbol{s} = (\lambda^{\gamma_1} s_1, \cdots, \lambda^{\gamma_d} s_d)$$
(4)

- Rectangle  $]\mathbf{0}, \lambda^{\Gamma} s]$  grow at different rate  $\lambda^{\gamma_j}$  in *j*th direction
- For i.i.d. RF X the presence of γ<sub>j</sub> does not make any difference of the limit which is a stable sheet on R<sup>d</sup><sub>+</sub> (except for a change of normalization)

$$X_{\lambda,\gamma}(\boldsymbol{s}) = \sum_{\boldsymbol{t} \in [\boldsymbol{0},\lambda^{\Gamma}\boldsymbol{s}]} X(\boldsymbol{t}), \qquad \lambda^{\Gamma}\boldsymbol{s} = (\lambda^{\gamma_1} s_1, \cdots, \lambda^{\gamma_d} s_d)$$
(4)

- Rectangle  $]\mathbf{0}, \lambda^{\Gamma} \mathbf{s}]$  grow at different rate  $\lambda^{\gamma_j}$  in *j*th direction
- For i.i.d. RF X the presence of  $\gamma_j$  does not make any difference of the limit which is a stable sheet on  $\mathbb{R}^4_+$  (except for a change of normalization)
- The same indifference to  $\gamma_i$  of the limit in (4) is expected under weak dependence

$$X_{\lambda,\gamma}(\boldsymbol{s}) = \sum_{\boldsymbol{t} \in [\boldsymbol{0},\lambda^{\Gamma}\boldsymbol{s}]} X(\boldsymbol{t}), \qquad \lambda^{\Gamma}\boldsymbol{s} = (\lambda^{\gamma_1} s_1, \cdots, \lambda^{\gamma_d} s_d)$$
(4)

- Rectangle  $]\mathbf{0}, \lambda^{\Gamma} \mathbf{s}]$  grow at different rate  $\lambda^{\gamma_j}$  in *j*th direction
- For i.i.d. RF X the presence of  $\gamma_j$  does not make any difference of the limit which is a stable sheet on  $\mathbb{R}^4_+$  (except for a change of normalization)
- The same indifference to  $\gamma_i$  of the limit in (4) is expected under weak dependence
- Surprising: for a large class of *LRD X* scaling limits of (4) exist for any  $\gamma$  and depend on  $\gamma \in \mathbb{R}^d_+$ , moreover the number of different limits is finite.

$$X_{\lambda,\gamma}(\boldsymbol{s}) = \sum_{\boldsymbol{t} \in [0,\lambda^{\Gamma}\boldsymbol{s}]} X(\boldsymbol{t}), \qquad \lambda^{\Gamma}\boldsymbol{s} = (\lambda^{\gamma_1} s_1, \cdots, \lambda^{\gamma_d} s_d)$$
(4)

- Rectangle  $]\mathbf{0}, \lambda^{\Gamma} \mathbf{s}]$  grow at different rate  $\lambda^{\gamma_j}$  in *j*th direction
- For i.i.d. RF X the presence of  $\gamma_j$  does not make any difference of the limit which is a stable sheet on  $\mathbb{R}^4_+$  (except for a change of normalization)
- The same indifference to  $\gamma_i$  of the limit in (4) is expected under weak dependence
- Surprising: for a large class of LRD X scaling limits of (4) exist for any γ and depend on γ ∈ ℝ<sup>d</sup><sub>+</sub>, moreover the number of different limits is finite.
   In dimension d = 2 this number is 3: there exists γ<sub>0</sub> > 0 such that the limits do not depend on γ = (γ<sub>1</sub>, γ<sub>2</sub>) for <sup>γ<sub>2</sub></sup>/<sub>γ<sub>1</sub></sub> > γ<sub>0</sub> and <sup>γ<sub>2</sub></sup>/<sub>γ<sub>1</sub></sub> < γ<sub>0</sub>

$$X_{\lambda,\gamma}(\boldsymbol{s}) = \sum_{\boldsymbol{t} \in [0,\lambda^{\Gamma}\boldsymbol{s}]} X(\boldsymbol{t}), \qquad \lambda^{\Gamma}\boldsymbol{s} = (\lambda^{\gamma_1} s_1, \cdots, \lambda^{\gamma_d} s_d)$$
(4)

- Rectangle  $]\mathbf{0}, \lambda^{\Gamma} \mathbf{s}]$  grow at different rate  $\lambda^{\gamma_j}$  in *j*th direction
- For i.i.d. RF X the presence of γ<sub>j</sub> does not make any difference of the limit which is a stable sheet on R<sup>4</sup><sub>+</sub> (except for a change of normalization)
- The same indifference to  $\gamma_i$  of the limit in (4) is expected under weak dependence
- Surprising: for a large class of LRD X scaling limits of (4) exist for any γ and depend on γ ∈ ℝ<sup>d</sup><sub>+</sub>, moreover the number of different limits is finite. In dimension d = 2 this number is 3: there exists γ<sub>0</sub> > 0 such that the limits do not depend on γ = (γ<sub>1</sub>, γ<sub>2</sub>) for <sup>γ<sub>2</sub></sup>/<sub>γ<sub>1</sub></sub> > γ<sub>0</sub> and <sup>γ<sub>2</sub></sup>/<sub>γ<sub>1</sub></sub> < γ<sub>0</sub>
   We say that scaling transition occurs at critical <sup>γ<sub>2</sub></sup>/<sub>γ<sub>1</sub></sub> = γ<sub>0</sub> (ratio of scaling exponents on different axes of ℝ<sup>2</sup>)

### Some ref.:

Puplinskaitė, D. & S.D. (2015) Scaling transition for long-range dependent Gaussian random fields. Stoch. Proc. Appl. 125

Puplinskaitė, D. & S.D. (2016) Aggregation of autoregressive random fields and anisotropic long-range dependence. Bernoulli 22

Pilipauskaitė, V. & S.D. (2017) Scaling transition for nonlinear random fields with long-range dependence. Stoch. Proc. Appl. 127

Biermé, H., Durieu, O. & Wang, Y. (2017) Invariance principles for operator-scaling Gaussian random fields. Ann. Appl. Prob. 27

S.D. (2019) Anisotropic scaling limits of long-range dependent linear random fields on  $\mathbb{Z}^3$ . J. Math. Anal. Appl. 472

S.D. (2020) Scaling transition and edge effects for negatively dependent linear random fields on  $\mathbb{Z}^2.$  Stoch. Proc. Appl. 130

Damarackas, J. & Paulauskas, V. (2021) On Lamperti type limit theorem and scaling transition for random fields. J. Math. Anal. Appl. 497

S.D. (2022) Scaling transition for singular linear random fields on  $\mathbb{Z}^2$ : spectral approach. Preprint.

### • Some ref.:

Puplinskaitė, D. & S.D. (2015) Scaling transition for long-range dependent Gaussian random fields. Stoch. Proc. Appl. 125

Puplinskaitė, D. & S.D. (2016) Aggregation of autoregressive random fields and anisotropic long-range dependence. Bernoulli 22

Pilipauskaitė, V. & S.D. (2017) Scaling transition for nonlinear random fields with long-range dependence. Stoch. Proc. Appl. 127

Biermé, H., Durieu, O. & Wang, Y. (2017) Invariance principles for operator-scaling Gaussian random fields. Ann. Appl. Prob. 27

S.D. (2019) Anisotropic scaling limits of long-range dependent linear random fields on  $\mathbb{Z}^3$ . J. Math. Anal. Appl. 472

S.D. (2020) Scaling transition and edge effects for negatively dependent linear random fields on  $\mathbb{Z}^2.$  Stoch. Proc. Appl. 130

Damarackas, J. & Paulauskas, V. (2021) On Lamperti type limit theorem and scaling transition for random fields. J. Math. Anal. Appl. 497

S.D. (2022) Scaling transition for singular linear random fields on  $\mathbb{Z}^2$ : spectral approach. Preprint.

• Extension: (isotropic) *scaling with aggregation*: the limit distribution of a sum of *M independent* copies of (1):

$$X_{\lambda,M}(\phi) := \sum_{j=1}^{M} \int_{\mathbb{R}^d} X_j(t) \phi(t/\lambda) dt$$
 (5)

as  $M \to \infty$  and  $\lambda \to \infty$ 

### • Some ref.:

Puplinskaitė, D. & S.D. (2015) Scaling transition for long-range dependent Gaussian random fields. Stoch. Proc. Appl. 125

Puplinskaitė, D. & S.D. (2016) Aggregation of autoregressive random fields and anisotropic long-range dependence. Bernoulli 22

Pilipauskaitė, V. & S.D. (2017) Scaling transition for nonlinear random fields with long-range dependence. Stoch. Proc. Appl. 127

Biermé, H., Durieu, O. & Wang, Y. (2017) Invariance principles for operator-scaling Gaussian random fields. Ann. Appl. Prob. 27

S.D. (2019) Anisotropic scaling limits of long-range dependent linear random fields on  $\mathbb{Z}^3$ . J. Math. Anal. Appl. 472

S.D. (2020) Scaling transition and edge effects for negatively dependent linear random fields on  $\mathbb{Z}^2.$  Stoch. Proc. Appl. 130

Damarackas, J. & Paulauskas, V. (2021) On Lamperti type limit theorem and scaling transition for random fields. J. Math. Anal. Appl. 497

S.D. (2022) Scaling transition for singular linear random fields on  $\mathbb{Z}^2$ : spectral approach. Preprint.

• Extension: (isotropic) *scaling with aggregation*: the limit distribution of a sum of *M independent* copies of (1):

$$X_{\lambda,M}(\phi) := \sum_{j=1}^{M} \int_{\mathbb{R}^d} X_j(t) \phi(t/\lambda) dt$$
 (5)

as  $M \to \infty$  and  $\lambda \to \infty$ 

• For d = 1 and  $\phi(t) = \mathbb{I}(t \in ]0, s]), s \ge 0$  (5) represent the aggregated sum

$$X_{\lambda,M}(s) = \sum_{j=1}^{M} \int_{0}^{\lambda s} X_{j}(t) \mathrm{d}t, \quad s \ge 0$$
 (6)

of integrated independent and identically distributed 'input' processes  $X_j = \{X_j(t); t \in \mathbb{R}\}$  with LRD 'induced by heavy tails',

• For d = 1 and  $\phi(t) = \mathbb{I}(t \in ]0, s]), s \ge 0$  (5) represent the aggregated sum

$$X_{\lambda,M}(s) = \sum_{j=1}^{M} \int_{0}^{\lambda s} X_{j}(t) \mathrm{d}t, \quad s \ge 0$$
 (6)

of integrated independent and identically distributed 'input' processes  $X_j = \{X_j(t); t \in \mathbb{R}\}$  with LRD 'induced by heavy tails', M referred to as the 'connection rate'

• For d = 1 and  $\phi(t) = \mathbb{I}(t \in ]0, s]), s \ge 0$  (5) represent the aggregated sum

$$X_{\lambda,M}(s) = \sum_{j=1}^{M} \int_{0}^{\lambda s} X_j(t) \mathrm{d}t, \qquad s \ge 0$$
 (6)

of integrated independent and identically distributed 'input' processes  $X_j = \{X_j(t); t \in \mathbb{R}\}$  with LRD 'induced by heavy tails', M referred to as the 'connection rate'

Applications in telecommunications (X<sub>λ,M</sub>(s) 'aggregated workload from M independent sources') and econometrics (X<sub>λ,M</sub>(s) averaged panel data from M individual 'micro time series')

• For d = 1 and  $\phi(t) = \mathbb{I}(t \in ]0, s]), s \ge 0$  (5) represent the aggregated sum

$$X_{\lambda,M}(s) = \sum_{j=1}^{M} \int_{0}^{\lambda s} X_{j}(t) \mathrm{d}t, \qquad s \geq 0$$
 (6)

of integrated independent and identically distributed 'input' processes  $X_j = \{X_j(t); t \in \mathbb{R}\}$  with LRD 'induced by heavy tails', M referred to as the 'connection rate'

- Applications in telecommunications (X<sub>λ,M</sub>(s) 'aggregated workload from M independent sources') and econometrics (X<sub>λ,M</sub>(s) averaged panel data from M individual 'micro time series')
- 'Typical' result in the heavy-tailed aggregated traffic research says that there exists a critical 'connection rate'  $M_0 = M_0(\lambda) \to \infty \ (\lambda \to \infty)$  such that the (normalized) 'aggregated input'  $X_{\lambda,M}(s)$  tends to an  $\alpha$ -stable Lévy process or a Fractional Brownian Motion depending on whether  $M/M_0$  tends to 0 or  $\infty$ ; the critical growth  $M/M_0 \to c \in (0,\infty)$  results in a different 'intermediate' limit which is neither Gaussian nor stable

• For d = 1 and  $\phi(t) = \mathbb{I}(t \in ]0, s]), s \ge 0$  (5) represent the aggregated sum

$$X_{\lambda,M}(s) = \sum_{j=1}^{M} \int_{0}^{\lambda s} X_{j}(t) \mathrm{d}t, \qquad s \geq 0$$
 (6)

of integrated independent and identically distributed 'input' processes  $X_j = \{X_j(t); t \in \mathbb{R}\}$  with LRD 'induced by heavy tails', M referred to as the 'connection rate'

- Applications in telecommunications (X<sub>λ,M</sub>(s) 'aggregated workload from M independent sources') and econometrics (X<sub>λ,M</sub>(s) averaged panel data from M individual 'micro time series')
- 'Typical' result in the heavy-tailed aggregated traffic research says that there exists a critical 'connection rate'  $M_0 = M_0(\lambda) \to \infty \ (\lambda \to \infty)$  such that the (normalized) 'aggregated input'  $X_{\lambda,M}(s)$  tends to an  $\alpha$ -stable Lévy process or a Fractional Brownian Motion depending on whether  $M/M_0$  tends to 0 or  $\infty$ ; the critical growth  $M/M_0 \to c \in (0,\infty)$  results in a different 'intermediate' limit which is neither Gaussian nor stable

#### • 'Iterated limits': first $M \to \infty$ then $\lambda \to \infty$ or vice versa:

Willinger, W., Taqu, M.S., Leland, M. & Wilson, D. (1997) Self-similarity through high-variability: statistical analysis of Ethernet LAN traffic at the source level. IEEE/ACM Trans. Networking 5 Taquu, M.S., Willinger, W. & Sherman, R. (1997) Proof of a fundamental result in self-similar traffic modeling. Comput. Commun. Rev. 27 Lévy, J.B. & Taquu, M.S. (2000) Renewal reward processes with heavy-tailed interrenewal times and heavy-tailed rewards. Bernoulli 6 Zaffaroni, P. (2004) Contemporaneous aggregation of linear dynamic models in large economies. J. Econometrics 120 Puplinskaitė, D. and S.D. (2010) Aggregation of random coefficient AR1(1) process with infinite variance and idiosyncratic innovations. Adv. Appl. Probab. 42 Barczy, M., Nedényi, F. & Pap, G. (2017) Iterated scaling limits for aggregation of randomized INAR(1) processes with idiosyncratic Poisson innovations. J. Math. Anal. Appl. 451

#### • 'Iterated limits': first $M \to \infty$ then $\lambda \to \infty$ or vice versa:

Willinger, W., Taqu, M.S., Leland, M. & Wilson, D. (1997) Self-similarity through high-variability: statistical analysis of Ethernet LAN traffic at the source level. IEEE/ACM Trans. Networking 5 Taquu, M.S., Willinger, W. & Sherman, R. (1997) Proof of a fundamental result in self-similar traffic modeling. Comput. Commun. Rev. 27 Lévy, J.B. & Taquu, M.S. (2000) Renewal reward processes with heavy-tailed interrenewal times and heavy-tailed rewards. Bernoulli 6 Zaffaroni, P. (2004) Contemporaneous aggregation of linear dynamic models in large economies. J. Econometrics 120 Puplinskaitė, D. and S.D. (2010) Aggregation of random coefficient AR1(1) process with infinite variance and idiosyncratic innovations. Adv. Appl. Probab. 42 Barczy, M., Nedényi, F. & Pap, G. (2017) Iterated scaling limits for aggregation of randomized INAR(1) processes with idiosyncratic Poisson innovations. J. Math. Anal. Appl. 451

### • 'Joint limits': $M = M(\lambda) \rightarrow \infty$ together with $\lambda \rightarrow \infty$ :

Mikosch, T., Resnick, S., Rootzén, H. & Stegeman, A. (2002) Is network traffic approximated by stable Lévy motion or fractional Brownian motion? Ann. Appl. Probab. 12

Gaigalas, R. & Kaj, I. (2003) Convergence of scaled renewal processes and a packet arrival model. Bernoulli 9 Pipiras, V., Taqu, M.S. & Levy, L.B. (2004) Slow, fast, and arbitrary growth conditions for renewal reward processes when the renewals and the rewards are heavy-tailed. Bernoulli 10

Kaj, I. & Taqqu, M.S. (2008) Convergence to fractional Brownian motion and to the Telecom process: the integral representation approach. In: M.E. Vares and V. Sidoravicius (Eds.) In and Out of Equilibrium 2. Progress in Probability, vol. 60

Pilipauskaitė, V. & S.D. (2014) Joint temporal and contemporaneous aggregation of random-coefficient AR(1) processes. Stoch. Process. Appl. 124

Pilipauskaitė, V., Skorniakov, V. & S.D. (2020) Joint temporal and contemporaneous aggregation of random-coefficient AR(1) processes with infinite variance. Adv. Appl. Probab. 52

Leipus, R., Pilipauskaitė, V. & S.D. (2023) Aggregation of network traffic and anisotropic scaling of random fields. Th. Probab. Math. Statist.

#### • 'Iterated limits': first $M \to \infty$ then $\lambda \to \infty$ or vice versa:

Willinger, W., Taqu, M.S., Leland, M. & Wilson, D. (1997) Self-similarity through high-variability: statistical analysis of Ethernet LAN traffic at the source level. IEEE/ACM Trans. Networking 5 Taquu, M.S., Willinger, W. & Sherman, R. (1997) Proof of a fundamental result in self-similar traffic modeling. Comput. Commun. Rev. 27 Lévy, J.B. & Taquu, M.S. (2000) Renewal reward processes with heavy-tailed interrenewal times and heavy-tailed rewards. Bernoulli 6 Zaffaroni, P. (2004) Contemporaneous aggregation of linear dynamic models in large economies. J. Econometrics 120 Puplinskaitė, D. and S.D. (2010) Aggregation of random coefficient AR1(1) process with infinite variance and idiosyncratic innovations. Adv. Appl. Probab. 42 Barczy, M., Nedényi, F. & Pap, G. (2017) Iterated scaling limits for aggregation of randomized INAR(1) processes with idiosyncratic Poisson innovations. J. Math. Anal. Appl. 451

### • 'Joint limits': $M = M(\lambda) \rightarrow \infty$ together with $\lambda \rightarrow \infty$ :

Mikosch, T., Resnick, S., Rootzén, H. & Stegeman, A. (2002) Is network traffic approximated by stable Lévy motion or fractional Brownian motion? Ann. Appl. Probab. 12

Gaigalas, R. & Kaj, I. (2003) Convergence of scaled renewal processes and a packet arrival model. Bernoulli 9 Pipiras, V., Taqu, M.S. & Levy, L.B. (2004) Slow, fast, and arbitrary growth conditions for renewal reward processes when the renewals and the rewards are heavy-tailed. Bernoulli 10

Kaj, I. & Taqqu, M.S. (2008) Convergence to fractional Brownian motion and to the Telecom process: the integral representation approach. In: M.E. Vares and V. Sidoravicius (Eds.) In and Out of Equilibrium 2. Progress in Probability, vol. 60

Pilipauskaitė, V. & S.D. (2014) Joint temporal and contemporaneous aggregation of random-coefficient AR(1) processes. Stoch. Process. Appl. 124

Pilipauskaitė, V., Skorniakov, V. & S.D. (2020) Joint temporal and contemporaneous aggregation of random-coefficient AR(1) processes with infinite variance. Adv. Appl. Probab. 52

Leipus, R., Pilipauskaitė, V. & S.D. (2023) Aggregation of network traffic and anisotropic scaling of random fields. Th. Probab. Math. Statist.

• In these works, the critical 'connection rate'  $M_0 = M_0(\lambda)$  separating Gaussian ('fast connection rate') and stable ('slow connection rate') limits grows as  $M_0 = O(\lambda^{\gamma_0})$  with some  $\gamma_0 > 0$  (up to slowly varying factor).

• In these works, the critical 'connection rate'  $M_0 = M_0(\lambda)$  separating Gaussian ('fast connection rate') and stable ('slow connection rate') limits grows as  $M_0 = O(\lambda^{\gamma_0})$  with some  $\gamma_0 > 0$  (up to slowly varying factor). We argue that these works fit into the previous set-up of 'scaling limit with aggregation' for RF  $X'(t, j) := X_j(t)$  on  $(t, j) \in \mathbb{R} \times \mathbb{Z}$ .

In these works, the critical 'connection rate' M<sub>0</sub> = M<sub>0</sub>(λ) separating Gaussian ('fast connection rate') and stable ('slow connection rate') limits grows as M<sub>0</sub> = O(λ<sup>γ0</sup>) with some γ<sub>0</sub> > 0 (up to slowly varying factor). We argue that these works fit into the previous set-up of 'scaling limit with aggregation' for RF X'(t, j) := X<sub>j</sub>(t) on (t, j) ∈ ℝ × ℤ. The trichotomy of scaling limits can be interpreted as scaling transition for RF X'.

- In these works, the critical 'connection rate' M<sub>0</sub> = M<sub>0</sub>(λ) separating Gaussian ('fast connection rate') and stable ('slow connection rate') limits grows as M<sub>0</sub> = O(λ<sup>γ0</sup>) with some γ<sub>0</sub> > 0 (up to slowly varying factor). We argue that these works fit into the previous set-up of 'scaling limit with aggregation' for RF X'(t, j) := X<sub>j</sub>(t) on (t, j) ∈ ℝ × ℤ. The trichotomy of scaling limits can be interpreted as scaling transition for RF X'.
- Unsurprisingly, most existing results on scaling limits of LRD RFs ( $d \ge 2$ ) (with or without aggregation) apply to linear models.

- In these works, the critical 'connection rate' M<sub>0</sub> = M<sub>0</sub>(λ) separating Gaussian ('fast connection rate') and stable ('slow connection rate') limits grows as M<sub>0</sub> = O(λ<sup>γ0</sup>) with some γ<sub>0</sub> > 0 (up to slowly varying factor). We argue that these works fit into the previous set-up of 'scaling limit with aggregation' for RF X'(t, j) := X<sub>j</sub>(t) on (t, j) ∈ ℝ × ℤ. The trichotomy of scaling limits can be interpreted as scaling transition for RF X'.
- Unsurprisingly, most existing results on scaling limits of LRD RFs (d ≥ 2) (with or without aggregation) apply to linear models.
   A notable exception is *Gaussian subordinated RFs* (written as a nonlinear function G(X(t)) of a Gaussian LRD RF X) treated via Hermite expansion 'Dobrushin-Major-Taqqu [DMT] theory'.

- In these works, the critical 'connection rate' M<sub>0</sub> = M<sub>0</sub>(λ) separating Gaussian ('fast connection rate') and stable ('slow connection rate') limits grows as M<sub>0</sub> = O(λ<sup>γ0</sup>) with some γ<sub>0</sub> > 0 (up to slowly varying factor). We argue that these works fit into the previous set-up of 'scaling limit with aggregation' for RF X'(t, j) := X<sub>j</sub>(t) on (t, j) ∈ ℝ × ℤ. The trichotomy of scaling limits can be interpreted as scaling transition for RF X'.
- Unsurprisingly, most existing results on scaling limits of LRD RFs (d ≥ 2) (with or without aggregation) apply to linear models.

A notable exception is *Gaussian subordinated RFs* (written as a nonlinear function G(X(t)) of a Gaussian LRD RF X) treated via Hermite expansion 'Dobrushin-Major-Taggu [DMT] theory'.

This is in contrast to the one-dimensional case d = 1, where the martingale approach developed in Ho, H.-C. & Hsing, T. (1997) Limit theorems for functionals of moving averages. Ann. Probab. 25 is applicable to nonlinear functions and statistics of *causal LRD moving averages*.

- In these works, the critical 'connection rate' M<sub>0</sub> = M<sub>0</sub>(λ) separating Gaussian ('fast connection rate') and stable ('slow connection rate') limits grows as M<sub>0</sub> = O(λ<sup>γ0</sup>) with some γ<sub>0</sub> > 0 (up to slowly varying factor). We argue that these works fit into the previous set-up of 'scaling limit with aggregation' for RF X'(t, j) := X<sub>j</sub>(t) on (t, j) ∈ ℝ × ℤ. The trichotomy of scaling limits can be interpreted as scaling transition for RF X'.
- Unsurprisingly, most existing results on scaling limits of LRD RFs ( $d \ge 2$ ) (with or without aggregation) apply to linear models.

A notable exception is *Gaussian subordinated RFs* (written as a nonlinear function G(X(t)) of a Gaussian LRD RF X) treated via Hermite expansion 'Dobrushin-Major-Taggu [DMT] theory'.

This is in contrast to the one-dimensional case d = 1, where the martingale approach developed in Ho, H.-C. & Hsing, T. (1997) Limit theorems for functionals of moving averages. Ann. Probab. 25 is applicable to nonlinear functions and statistics of *causal LRD moving averages*.

Nonlinear models are important since most statistics are nonlinear.

- In these works, the critical 'connection rate' M<sub>0</sub> = M<sub>0</sub>(λ) separating Gaussian ('fast connection rate') and stable ('slow connection rate') limits grows as M<sub>0</sub> = O(λ<sup>γ0</sup>) with some γ<sub>0</sub> > 0 (up to slowly varying factor). We argue that these works fit into the previous set-up of 'scaling limit with aggregation' for RF X'(t, j) := X<sub>j</sub>(t) on (t, j) ∈ ℝ × ℤ. The trichotomy of scaling limits can be interpreted as scaling transition for RF X'.
- Unsurprisingly, most existing results on scaling limits of LRD RFs ( $d \ge 2$ ) (with or without aggregation) apply to linear models.

A notable exception is *Gaussian subordinated RFs* (written as a nonlinear function G(X(t)) of a Gaussian LRD RF X) treated via Hermite expansion 'Dobrushin-Major-Taggu [DMT] theory'.

This is in contrast to the one-dimensional case d = 1, where the martingale approach developed in Ho, H.-C. & Hsing, T. (1997) Limit theorems for functionals of moving averages. Ann. Probab. 25 is applicable to nonlinear functions and statistics of *causal LRD moving averages*.

Nonlinear models are important since most statistics are nonlinear.

• Causality not very natural in spatial context & HH (1997) method hard to adopt in noncausal case.

• Causality not very natural in spatial context & HH (1997) method hard to adopt in noncausal case. But:

Doukhan, P., Lang, G. & S.D. (2002) Asymptotics of weighted empirical processes of linear random fields with long range dependence. Ann. Inst. H. Poincaré 38 for  $G(x) = \mathbb{I}(x \leq y)$ .

• Causality not very natural in spatial context & HH (1997) method hard to adopt in noncausal case. But:

Doukhan, P., Lang, G. & S.D. (2002) Asymptotics of weighted empirical processes of linear random fields with long range dependence. Ann. Inst. H. Poincaré 38 for  $G(x) = \mathbb{I}(x \leq y)$ . This result was applied to goodness-of-fit testing

Koul, H.L., Mimoto, N. & S.D. (2016) Goodness-of-fit tests for marginal distribution of linear random fields with long memory. Metrika 79

• Causality not very natural in spatial context & HH (1997) method hard to adopt in noncausal case. But:

Doukhan, P., Lang, G. & S.D. (2002) Asymptotics of weighted empirical processes of linear random fields with long range dependence. Ann. Inst. H. Poincaré 38 for  $G(x) = \mathbb{I}(x \leq y)$ . This result was applied to goodness-of-fit testing

Koul, H.L., Mimoto, N. & S.D. (2016) Goodness-of-fit tests for marginal distribution of linear random fields with long memory. Metrika 79

# 2. Random grain model

 Causality not very natural in spatial context & HH (1997) method hard to adopt in noncausal case. But:

Doukhan, P., Lang, G. & S.D. (2002) Asymptotics of weighted empirical processes of linear random fields with long range dependence. Ann. Inst. H. Poincaré 38 for  $G(x) = \mathbb{I}(x \leq y)$ . This result was applied to goodness-of-fit testing

Koul, H.L., Mimoto, N. & S.D. (2016) Goodness-of-fit tests for marginal distribution of linear random fields with long memory. Metrika 79

## 2. Random grain model

A 'superposition of uniformly scattered in  $\mathbb{R}^{d}$ ' and randomly dilated grains:

$$X(\boldsymbol{t}) := \sum_{j=1}^{\infty} \mathbb{I}(\boldsymbol{t} \in (\boldsymbol{u}_j + \Xi_j)), \qquad \boldsymbol{t} \in \mathbb{R}^d.$$
(7)

where:

•  $\{u_j\} \subset \mathbb{R}^d$ : Poisson process of 'centers' or 'germs' with uniform intensity du

 Causality not very natural in spatial context & HH (1997) method hard to adopt in noncausal case. But:

Doukhan, P., Lang, G. & S.D. (2002) Asymptotics of weighted empirical processes of linear random fields with long range dependence. Ann. Inst. H. Poincaré 38 for  $G(x) = \mathbb{I}(x \leq y)$ . This result was applied to goodness-of-fit testing

Koul, H.L., Mimoto, N. & S.D. (2016) Goodness-of-fit tests for marginal distribution of linear random fields with long memory. Metrika 79

# 2. Random grain model

A 'superposition of uniformly scattered in  $\mathbb{R}^{d}$ ' and randomly dilated grains:

$$X(\boldsymbol{t}) := \sum_{j=1}^{\infty} \mathbb{I}(\boldsymbol{t} \in (\boldsymbol{u}_j + \Xi_j)), \qquad \boldsymbol{t} \in \mathbb{R}^d.$$
(7)

where:

- $\{u_j\} \subset \mathbb{R}^d$ : Poisson process of 'centers' or 'germs' with uniform intensity du
- $\Xi_j = R_j^{1/d} \Xi^0$ ,  $\{R, R_j > 0\}$  i.i.d. with  $F(dr) := P(R \in dr)$  independent of  $\{u_j\}$ ,  $\Xi^0$  ('generic grain'): a deterministic bounded Borel subset of  $\mathbb{R}^d$
Causality not very natural in spatial context & HH (1997) method hard to adopt in noncausal case. But:

Doukhan, P., Lang, G. & S.D. (2002) Asymptotics of weighted empirical processes of linear random fields with long range dependence. Ann. Inst. H. Poincaré 38 for  $G(x) = \mathbb{I}(x \leq y)$ . This result was applied to goodness-of-fit testing

Koul, H.L., Mimoto, N. & S.D. (2016) Goodness-of-fit tests for marginal distribution of linear random fields with long memory. Metrika 79

# 2. Random grain model

A 'superposition of uniformly scattered in  $\mathbb{R}^{d}$ ' and randomly dilated grains:

$$X(\boldsymbol{t}) := \sum_{j=1}^{\infty} \mathbb{I}(\boldsymbol{t} \in (\boldsymbol{u}_j + \Xi_j)), \qquad \boldsymbol{t} \in \mathbb{R}^d.$$
(7)

where:

- $\{u_j\} \subset \mathbb{R}^d$ : Poisson process of 'centers' or 'germs' with uniform intensity du
- $\Xi_j = R_j^{1/d} \Xi^0$ ,  $\{R, R_j > 0\}$  i.i.d. with  $F(dr) := P(R \in dr)$  independent of  $\{u_j\}$ ,  $\Xi^0$  ('generic grain'): a deterministic bounded Borel subset of  $\mathbb{R}^d$

•  $Leb_d(\Xi_j) = R_j Leb_d(\Xi^0)$ : dilates  $Leb_d(\Xi^0)$  by random factor  $R_j$ 

- $Leb_d(\Xi_j) = R_j Leb_d(\Xi^0)$ : dilates  $Leb_d(\Xi^0)$  by random factor  $R_j$
- X(t) in (7) counts the number of random grains which cover  $t \in \mathbb{R}^d$

- $Leb_d(\Xi_j) = R_j Leb_d(\Xi^0)$ : dilates  $Leb_d(\Xi^0)$  by random factor  $R_j$
- X(t) in (7) counts the number of random grains which cover  $t \in \mathbb{R}^d$
- X(t) has marginal Poisson distribution with mean

$$\mu = \mathrm{E}X(t) = \int_{\mathbb{R}^d} \mathrm{P}(t - u \in R^{1/d} \Xi^0) \mathrm{d}u = Leb_d(\Xi^0) \mathrm{E}R < \infty$$

- $Leb_d(\Xi_j) = R_j Leb_d(\Xi^0)$ : dilates  $Leb_d(\Xi^0)$  by random factor  $R_j$
- X(t) in (7) counts the number of random grains which cover  $t \in \mathbb{R}^d$
- X(t) has marginal Poisson distribution with mean

$$\mu = \mathrm{E}X(t) = \int_{\mathbb{R}^d} \mathrm{P}(t - u \in R^{1/d} \Xi^0) \mathrm{d}u = Leb_d(\Xi^0) \mathrm{E}R < \infty$$

and stochastic integral representation

$$X(t) = \int_{\mathbb{R}^d \times \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}(\mathrm{d} u, \mathrm{d} r), \quad t \in \mathbb{R}^d,$$
(8)

 $\mathcal{N}(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r})$ : Poisson random measure with  $\mathrm{E}\mathcal{N}(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = \mathrm{d}\boldsymbol{u}F(\mathrm{d}\boldsymbol{r})$ 

- Leb<sub>d</sub>(Ξ<sub>j</sub>) = R<sub>j</sub>Leb<sub>d</sub>(Ξ<sup>0</sup>): dilates Leb<sub>d</sub>(Ξ<sup>0</sup>) by random factor R<sub>j</sub>
- X(t) in (7) counts the number of random grains which cover  $t \in \mathbb{R}^d$
- X(t) has marginal Poisson distribution with mean

$$\mu = \mathrm{E}X(t) = \int_{\mathbb{R}^d} \mathrm{P}(t - u \in R^{1/d} \Xi^0) \mathrm{d}u = Leb_d(\Xi^0) \mathrm{E}R < \infty$$

and stochastic integral representation

$$X(t) = \int_{\mathbb{R}^d \times \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}(\mathrm{d} u, \mathrm{d} r), \quad t \in \mathbb{R}^d,$$
(8)

 $\mathcal{N}(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r})$ : Poisson random measure with  $\mathrm{E}\mathcal{N}(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = \mathrm{d}\boldsymbol{u}F(\mathrm{d}\boldsymbol{r})$ 

Closely related object: the (random) Boolean set:

$$\mathcal{X} := \bigcup_{j=1}^{\infty} (\boldsymbol{u}_j + R_j^{1/d} \Xi^0) \quad \subset \mathbb{R}^d.$$
(9)

- Leb<sub>d</sub>(Ξ<sub>j</sub>) = R<sub>j</sub>Leb<sub>d</sub>(Ξ<sup>0</sup>): dilates Leb<sub>d</sub>(Ξ<sup>0</sup>) by random factor R<sub>j</sub>
- X(t) in (7) counts the number of random grains which cover  $t \in \mathbb{R}^d$
- X(t) has marginal Poisson distribution with mean

$$\mu = \mathrm{E}X(t) = \int_{\mathbb{R}^d} \mathrm{P}(t - u \in R^{1/d} \Xi^0) \mathrm{d}u = Leb_d(\Xi^0) \mathrm{E}R < \infty$$

and stochastic integral representation

$$X(t) = \int_{\mathbb{R}^d \times \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}(\mathrm{d} u, \mathrm{d} r), \quad t \in \mathbb{R}^d,$$
(8)

 $\mathcal{N}(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r})$ : Poisson random measure with  $\mathrm{E}\mathcal{N}(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = \mathrm{d}\boldsymbol{u}\boldsymbol{F}(\mathrm{d}\boldsymbol{r})$ 

Closely related object: the (random) Boolean set:

$$\mathcal{X} := \bigcup_{j=1}^{\infty} (\boldsymbol{u}_j + R_j^{1/d} \Xi^0) \quad \subset \mathbb{R}^d.$$
(9)

Boolean model is basic in stochastic geometry and stereology

• 
$$\Xi^0 = \{ \| \boldsymbol{t} \| \le 1 \}$$
 unit ball: random ball model

- $\Xi^0 = \{ \| \boldsymbol{t} \| \leq 1 \}$  unit ball: random ball model
- Trajectories of RG model very different from Gaussian:



Isotropically scaled random ball model,  $\gamma = 1, \alpha = 3/2$ . Left:  $\lambda = 5$ , right:  $\lambda = 10$ 



Same anisotropically (  $\gamma=$  3) scaled random ball model. Left:  $\lambda=$  5, right:  $\lambda=$  10



Same anisotropically ( $\gamma = 3$ ) scaled random ball model. Left:  $\lambda = 5$ , right:  $\lambda = 10$ 

 Isotropic scaling with aggregation of RG model was discussed in important works: [KLNS] Kaj, I., Leskelä, L., Norros, I. & Schmidt, V. (2007) Scaling limits for random fields with long-range dependence. Ann. Probab. 35
 [BEK] Biermé, H., Estrade, A. & Kaj, I. (2010) Self-similar random fields and rescaled random balls models. J. Theoret. Probab. 23



Same anisotropically ( $\gamma = 3$ ) scaled random ball model. Left:  $\lambda = 5$ , right:  $\lambda = 10$ 

 Isotropic scaling with aggregation of RG model was discussed in important works: [KLNS] Kaj, I., Leskelä, L., Norros, I. & Schmidt, V. (2007) Scaling limits for random fields with long-range dependence. Ann. Probab. 35
 [BEK] Biermé, H., Estrade, A. & Kaj, I. (2010) Self-similar random fields and rescaled random balls models. J. Theoret. Probab. 23
 Anisotropic scaling without aggregation of RG model (d = 2):

Anisotropic scaling without aggregation of RG model (d = 2): Pilipauskaitė, V. & S.D. (2016) Anisotropic scaling of random grain model with application to network traffic. J. Appl. Probab. 53



Same anisotropically ( $\gamma = 3$ ) scaled random ball model. Left:  $\lambda = 5$ , right:  $\lambda = 10$ 

 Isotropic scaling with aggregation of RG model was discussed in important works: [KLNS] Kaj, I., Leskelä, L., Norros, I. & Schmidt, V. (2007) Scaling limits for random fields with long-range dependence. Ann. Probab. 35
 [BEK] Biermé, H., Estrade, A. & Kaj, I. (2010) Self-similar random fields and rescaled random balls models. J. Theoret. Probab. 23
 Anisotropic scaling without aggregation of RG model (d = 2):

Anisotropic scaling without aggregation of RG model (d = 2): Pilipauskaitė, V. & S.D. (2016) Anisotropic scaling of random grain model with application to network traffic. J. Appl. Probab. 53

Covariance  $r_X(t) = Cov(X(0), X(t))$  of RG model writes as

$$r_X(\mathbf{t}) = \int_0^\infty Leb_d(\Xi^0 \cap (\Xi^0 - r^{-1/d}\mathbf{t})) r F(\mathrm{d}r).$$

Well-known: RG is LRD [= nonintegrable covariance] if  $P(R > r) = F(r, \infty)$  varies regularly at  $\infty$  with exponent  $\alpha \in (1, 2)$ .

Covariance  $r_X(t) = Cov(X(0), X(t))$  of RG model writes as

$$r_X(\mathbf{t}) = \int_0^\infty Leb_d(\Xi^0 \cap (\Xi^0 - r^{-1/d}\mathbf{t})) r F(\mathrm{d}r).$$

Well-known: RG is LRD [= nonintegrable covariance] if  $P(R > r) = F(r, \infty)$  varies regularly at  $\infty$  with exponent  $\alpha \in (1, 2)$ .

Assumption LRD  $\Xi^0 \subset \mathbb{R}^d$  is a bounded Borel set whereas F(dr) = f(r)dr has density function s.t.

$$f(r) \sim c_f r^{-1-\alpha}, \qquad r \to \infty \quad (\exists c_f > 0, \quad \alpha \in (1,2)).$$
 (10)

Moreover,  $(r, z) \mapsto Leb_d \left( \Xi^0 \cap (\Xi^0 - r^{-1/d} z) \right)$  is continuous on  $(r, z) \in \mathbb{R}_+ \times \{ \|z\| = 1 \}$ .

• mild regularity of boundary  $\partial \Xi^0$ 

Covariance  $r_X(t) = Cov(X(0), X(t))$  of RG model writes as

$$r_X(\mathbf{t}) = \int_0^\infty Leb_d(\Xi^0 \cap (\Xi^0 - r^{-1/d}\mathbf{t})) r F(\mathrm{d}r).$$

Well-known: RG is LRD [= nonintegrable covariance] if  $P(R > r) = F(r, \infty)$  varies regularly at  $\infty$  with exponent  $\alpha \in (1, 2)$ .

Assumption LRD  $\Xi^0 \subset \mathbb{R}^d$  is a bounded Borel set whereas F(dr) = f(r)dr has density function s.t.

$$f(r) \sim c_f r^{-1-\alpha}, \qquad r \to \infty \quad (\exists c_f > 0, \quad \alpha \in (1,2)).$$
 (10)

Moreover,  $(r, z) \mapsto Leb_d \left( \Xi^0 \cap (\Xi^0 - r^{-1/d} z) \right)$  is continuous on  $(r, z) \in \mathbb{R}_+ \times \{ \|z\| = 1 \}$ .

- mild regularity of boundary  $\partial \Xi^0$
- Under Assumption LRD

$$r_X(t) \sim \|t\|^{-d(\alpha-1)}\ell(rac{t}{\|t\|}), \quad |t| \to \infty, \quad 1 < lpha < 2,$$
 (11)

where  $\ell(z)$ , ||z|| = 1 is a bdd cont. (angular) function

$$\ell(\boldsymbol{z}) := c_f \int_0^\infty Leb_d \big( \Xi^0 \cap (\Xi^0 - r^{-1/d} \, \boldsymbol{z}) \big) r^{-\alpha} \mathrm{d} r.$$

• LRD property holds for  $1 < \alpha < 2$  and does not hold for  $\alpha > 2$ ,  $\alpha > 1$  necessary for  $EX(t) < \infty$  and existence of X

- LRD property holds for 1 < α < 2 and does not hold for α > 2, α > 1 necessary for EX(t) < ∞ and existence of X</li>
- For  $d = 1, \Xi^0 = [0, 1]$ ,  $X = \{X(t); t \in \mathbb{R}\}$  is stationary  $M/G/\infty$  queue:

$$X(t) = \sum_{u_j \leq t} \mathbb{I}(t - u_j \leq R_j)$$

- LRD property holds for 1 < α < 2 and does not hold for α > 2, α > 1 necessary for EX(t) < ∞ and existence of X</li>
- For  $d = 1, \Xi^0 = [0, 1]$ ,  $X = \{X(t); t \in \mathbb{R}\}$  is stationary  $M/G/\infty$  queue:

$$X(t) = \sum_{u_j \leq t} \mathbb{I}(t - u_j \leq R_j)$$

counting the number of customers at time t at a queueing system with standard Poisson arrivals  $u_i$ , service times  $R_i$  and infinite waiting room,

- LRD property holds for 1 < α < 2 and does not hold for α > 2, α > 1 necessary for EX(t) < ∞ and existence of X</li>
- For  $d = 1, \Xi^0 = [0, 1]$ ,  $X = \{X(t); t \in \mathbb{R}\}$  is stationary  $M/G/\infty$  queue:

$$X(t) = \sum_{u_j \leq t} \mathbb{I}(t - u_j \leq R_j)$$

counting the number of customers at time t at a queueing system with standard Poisson arrivals  $u_j$ , service times  $R_j$  and infinite waiting room, also called the *infinite source Poisson model*.

- LRD property holds for 1 < α < 2 and does not hold for α > 2, α > 1 necessary for EX(t) < ∞ and existence of X</li>
- For  $d = 1, \Xi^0 = [0, 1]$ ,  $X = \{X(t); t \in \mathbb{R}\}$  is stationary  $M/G/\infty$  queue:

$$X(t) = \sum_{u_j \leq t} \mathbb{I}(t - u_j \leq R_j)$$

counting the number of customers at time t at a queueing system with standard Poisson arrivals  $u_j$ , service times  $R_j$  and infinite waiting room,

also called the *infinite source Poisson model*. Then  $r_X(t) = \int_t^{\infty} P(R > r) dr = O(t^{-(\alpha-1)})$  is LRD for  $\alpha \in (1,2)$ 

- LRD property holds for 1 < α < 2 and does not hold for α > 2, α > 1 necessary for EX(t) < ∞ and existence of X</li>
- For  $d = 1, \Xi^0 = [0, 1]$ ,  $X = \{X(t); t \in \mathbb{R}\}$  is stationary  $M/G/\infty$  queue:

$$X(t) = \sum_{u_j \leq t} \mathbb{I}(t - u_j \leq R_j)$$

counting the number of customers at time t at a queueing system with standard Poisson arrivals  $u_j$ , service times  $R_j$  and infinite waiting room, also called the *infinite source Poisson model*. Then  $r_X(t) = \int_{-\infty}^{\infty} P(R > r) dr = O(t^{-(\alpha-1)})$  is LRD for  $\alpha \in (1, 2)$ 

(11) implies the asymptotics of the variance of X<sub>λ</sub>(φ) = ∫<sub>ℝ<sup>d</sup></sub> X(t)φ(t/λ)dt: for any φ ∈ Φ = L<sup>1</sup>(ℝ<sup>d</sup>) ∩ L<sup>∞</sup>(ℝ<sup>d</sup>)

$$\operatorname{Var}(X_{\lambda}(\phi)) \sim \lambda^{d(3-\alpha)} c(\phi), \quad \lambda \to \infty,$$
 (12)

- LRD property holds for 1 < α < 2 and does not hold for α > 2, α > 1 necessary for EX(t) < ∞ and existence of X</li>
- For  $d = 1, \Xi^0 = [0, 1]$ ,  $X = \{X(t); t \in \mathbb{R}\}$  is stationary  $M/G/\infty$  queue:

$$X(t) = \sum_{u_j \leq t} \mathbb{I}(t - u_j \leq R_j)$$

counting the number of customers at time t at a queueing system with standard Poisson arrivals  $u_j$ , service times  $R_j$  and infinite waiting room, also called the *infinite source Poisson model*. Then  $r_X(t) = \int_{-\infty}^{\infty} P(R > r) dr = O(t^{-(\alpha-1)})$  is LRD for  $\alpha \in (1, 2)$ 

(11) implies the asymptotics of the variance of X<sub>λ</sub>(φ) = ∫<sub>ℝ<sup>d</sup></sub> X(t)φ(t/λ)dt: for any φ ∈ Φ = L<sup>1</sup>(ℝ<sup>d</sup>) ∩ L<sup>∞</sup>(ℝ<sup>d</sup>)

$$\operatorname{Var}(X_{\lambda}(\phi)) \sim \lambda^{d(3-\alpha)} c(\phi), \quad \lambda \to \infty,$$
 (12)

where

$$c(\phi) := \int_{\mathbb{R}^{2d}} \phi(\boldsymbol{t}_1) \phi(\boldsymbol{t}_2) \ell\left(\frac{\boldsymbol{t}_1 - \boldsymbol{t}_2}{\|\boldsymbol{t}_1 - \boldsymbol{t}_2\|}\right) \frac{\mathrm{d}\boldsymbol{t}_1 \mathrm{d}\boldsymbol{t}_2}{\|\boldsymbol{t}_1 - \boldsymbol{t}_2\|^{d(\alpha-1)}}$$

• Scaling limits with aggregation for sums  $X_{\lambda,M}(\phi) := \sum_{j=1}^{M} \int_{\mathbb{R}^d} X_j(t) \phi(t/\lambda) dt$  of independent RG models.

• Scaling limits with aggregation for sums  $X_{\lambda,M}(\phi) := \sum_{j=1}^{M} \int_{\mathbb{R}^d} X_j(t) \phi(t/\lambda) dt$  of independent RG models. These can be identified as integral

$$egin{aligned} X_{\lambda,M}(\phi) &= \int_{\mathbb{R}^d} X_M(oldsymbol{t}) \phi(oldsymbol{t}/\lambda) \mathrm{d}oldsymbol{t}, & ext{where} \ X_M(oldsymbol{t}) &= \int_{\mathbb{R}^d imes \mathbb{R}^+} \mathbb{I}(oldsymbol{t} - oldsymbol{u} \in r^{1/d} \Xi^0) \mathcal{N}_M(\mathrm{d}oldsymbol{u}, \mathrm{d}oldsymbol{r}) \end{aligned}$$

w.r.t. Poisson measure with  $E\mathcal{N}_M(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = M\mathrm{d}\boldsymbol{u}F(\mathrm{d}\boldsymbol{r})$ 

Scaling limits with aggregation for sums X<sub>λ,M</sub>(φ) := Σ<sup>M</sup><sub>j=1</sub> ∫<sub>ℝ<sup>d</sup></sub> X<sub>j</sub>(t)φ(t/λ)dt of independent RG models. These can be identified as integral

$$egin{aligned} X_{\lambda,M}(\phi) &= \int_{\mathbb{R}^d} X_M(oldsymbol{t}) \phi(oldsymbol{t}/\lambda) \mathrm{d}oldsymbol{t}, & ext{where} \ X_M(oldsymbol{t}) &= \int_{\mathbb{R}^d imes \mathbb{R}^+} \mathbb{I}(oldsymbol{t} - oldsymbol{u} \in r^{1/d} \Xi^0) \mathcal{N}_M(\mathrm{d}oldsymbol{u}, \mathrm{d}oldsymbol{r}) \end{aligned}$$

w.r.t. Poisson measure with  $E\mathcal{N}_M(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = M\mathrm{d}\boldsymbol{u}F(\mathrm{d}\boldsymbol{r})$ 

• For Poisson based models, aggregation amounts to multiplication of intensity

• Scaling limits with aggregation for sums  $X_{\lambda,M}(\phi) := \sum_{j=1}^{M} \int_{\mathbb{R}^d} X_j(t) \phi(t/\lambda) dt$  of independent RG models. These can be identified as integral

$$egin{aligned} X_{\lambda,\mathcal{M}}(\phi) &= \int_{\mathbb{R}^d} X_{\mathcal{M}}(oldsymbol{t}) \phi(oldsymbol{t}/\lambda) \mathrm{d}oldsymbol{t}, & ext{where} \ X_{\mathcal{M}}(oldsymbol{t}) &= \int_{\mathbb{R}^d imes \mathbb{R}^+} \mathbb{I}(oldsymbol{t} - oldsymbol{u} \in r^{1/d} \Xi^0) \mathcal{N}_{\mathcal{M}}(\mathrm{d}oldsymbol{u}, \mathrm{d}oldsymbol{r}) \end{aligned}$$

w.r.t. Poisson measure with  $E\mathcal{N}_M(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = M\mathrm{d}\boldsymbol{u}F(\mathrm{d}\boldsymbol{r})$ 

- For Poisson based models, aggregation amounts to multiplication of intensity
- [KLNS] discuss scaling limit of  $X_{\lambda,M}(\phi)$  indexed by signed (Riesz) measures  $\phi$ .

Scaling limits with aggregation for sums X<sub>λ,M</sub>(φ) := Σ<sup>M</sup><sub>j=1</sub> ∫<sub>ℝ<sup>d</sup></sub> X<sub>j</sub>(t)φ(t/λ)dt of independent RG models. These can be identified as integral

$$egin{aligned} X_{\lambda,\mathcal{M}}(\phi) &= \int_{\mathbb{R}^d} X_{\mathcal{M}}(oldsymbol{t}) \phi(oldsymbol{t}/\lambda) \mathrm{d}oldsymbol{t}, & ext{where} \ X_{\mathcal{M}}(oldsymbol{t}) &= \int_{\mathbb{R}^d imes \mathbb{R}^+} \mathbb{I}(oldsymbol{t} - oldsymbol{u} \in r^{1/d} \Xi^0) \mathcal{N}_{\mathcal{M}}(\mathrm{d}oldsymbol{u}, \mathrm{d}oldsymbol{r}) \end{aligned}$$

w.r.t. Poisson measure with  $E\mathcal{N}_M(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = M\mathrm{d}\boldsymbol{u}F(\mathrm{d}\boldsymbol{r})$ 

- For Poisson based models, aggregation amounts to multiplication of intensity
- [KLNS] discuss scaling limit of  $X_{\lambda,M}(\phi)$  indexed by signed (Riesz) measures  $\phi$ . In this talk:

$$\phi \in \Phi = L^1(\mathbb{R}^d) \cap L^\infty(\mathbb{R}^d).$$

• Scaling limits with aggregation for sums  $X_{\lambda,M}(\phi) := \sum_{j=1}^{M} \int_{\mathbb{R}^d} X_j(t) \phi(t/\lambda) dt$  of independent RG models. These can be identified as integral

$$egin{aligned} X_{\lambda,\mathcal{M}}(\phi) &= \int_{\mathbb{R}^d} X_{\mathcal{M}}(oldsymbol{t}) \phi(oldsymbol{t}/\lambda) \mathrm{d}oldsymbol{t}, & ext{where} \ X_{\mathcal{M}}(oldsymbol{t}) &= \int_{\mathbb{R}^d imes \mathbb{R}^+} \mathbb{I}(oldsymbol{t} - oldsymbol{u} \in r^{1/d} \Xi^0) \mathcal{N}_{\mathcal{M}}(\mathrm{d}oldsymbol{u}, \mathrm{d}oldsymbol{r}) \end{aligned}$$

w.r.t. Poisson measure with  $E\mathcal{N}_M(\mathrm{d} u, \mathrm{d} r) = M\mathrm{d} uF(\mathrm{d} r)$ 

- For Poisson based models, aggregation amounts to multiplication of intensity
- [KLNS] discuss scaling limit of  $X_{\lambda,M}(\phi)$  indexed by signed (Riesz) measures  $\phi$ . In this talk:

$$\phi \in \Phi = L^1(\mathbb{R}^d) \cap L^\infty(\mathbb{R}^d).$$

• What happens when  $M \to \infty$  together with  $\lambda \to \infty$ ?

Scaling limits with aggregation for sums X<sub>λ,M</sub>(φ) := Σ<sup>M</sup><sub>j=1</sub> ∫<sub>ℝ<sup>d</sup></sub> X<sub>j</sub>(t)φ(t/λ)dt of independent RG models. These can be identified as integral

$$egin{aligned} X_{\lambda,\mathcal{M}}(\phi) &= \int_{\mathbb{R}^d} X_{\mathcal{M}}(oldsymbol{t}) \phi(oldsymbol{t}/\lambda) \mathrm{d}oldsymbol{t}, & ext{where} \ X_{\mathcal{M}}(oldsymbol{t}) &= \int_{\mathbb{R}^d imes \mathbb{R}^+} \mathbb{I}(oldsymbol{t} - oldsymbol{u} \in r^{1/d} \Xi^0) \mathcal{N}_{\mathcal{M}}(\mathrm{d}oldsymbol{u}, \mathrm{d}oldsymbol{r}) \end{aligned}$$

w.r.t. Poisson measure with  $E\mathcal{N}_M(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = M\mathrm{d}\boldsymbol{u}F(\mathrm{d}\boldsymbol{r})$ 

- For Poisson based models, aggregation amounts to multiplication of intensity
- [KLNS] discuss scaling limit of  $X_{\lambda,M}(\phi)$  indexed by signed (Riesz) measures  $\phi$ . In this talk:

$$\phi \in \Phi = L^1(\mathbb{R}^d) \cap L^\infty(\mathbb{R}^d).$$

• What happens when  $M \to \infty$  together with  $\lambda \to \infty$ ? For fixed  $t \in \mathbb{R}^d$  clearly  $(X_M(t) - EX_M(t))/M^{1/2} \stackrel{d}{\longrightarrow} N(0,\mu) \ (M \to \infty)$  with  $\mu = EX(t) = Var(X(t))$  by CLT
# Theorem (1)

Let Assumption LRD hold,  $M = \lambda^{\gamma}$  ( $\gamma > 0$ ). Then for any  $\phi \in \Phi$ 

$$\lambda^{-H(\gamma)}(X_{\lambda,M}(\phi) - \mathrm{E}X_{\lambda,M}(\phi)) \xrightarrow{\mathrm{d}} \begin{cases} B_{\alpha}(\phi), \quad \gamma > d(\alpha - 1), \ H(\gamma) = \frac{\gamma + (3 - \alpha)d}{2}, \\ L_{\alpha}(\phi), \quad \gamma < d(\alpha - 1), \ H(\gamma) = \frac{\gamma + d}{\alpha}, \\ J_{\alpha}(\phi), \quad \gamma = d(\alpha - 1), \ H(\gamma) = d. \end{cases}$$

• Thm essentially due to [KLNS]

# Theorem (1)

$$\lambda^{-H(\gamma)}(X_{\lambda,M}(\phi) - \mathcal{E}X_{\lambda,M}(\phi)) \xrightarrow{\mathrm{d}} \begin{cases} B_{\alpha}(\phi), \quad \gamma > d(\alpha - 1), \ H(\gamma) = \frac{\gamma + (3 - \alpha)d}{2}, \\ L_{\alpha}(\phi), \quad \gamma < d(\alpha - 1), \ H(\gamma) = \frac{\gamma + d}{\alpha}, \\ J_{\alpha}(\phi), \quad \gamma = d(\alpha - 1), \ H(\gamma) = d. \end{cases}$$

- Thm essentially due to [KLNS]
- $B_{\alpha}(\phi)$ : Gaussian RF with  $\operatorname{Var}(B_{\lambda}(\phi)) = c(\phi)$  in (12).

# Theorem (1)

$$\lambda^{-H(\gamma)}(X_{\lambda,M}(\phi) - \mathcal{E}X_{\lambda,M}(\phi)) \xrightarrow{\mathrm{d}} \begin{cases} B_{\alpha}(\phi), \quad \gamma > d(\alpha - 1), \ H(\gamma) = \frac{\gamma + (3 - \alpha)d}{2}, \\ L_{\alpha}(\phi), \quad \gamma < d(\alpha - 1), \ H(\gamma) = \frac{\gamma + d}{\alpha}, \\ J_{\alpha}(\phi), \quad \gamma = d(\alpha - 1), \ H(\gamma) = d. \end{cases}$$

- Thm essentially due to [KLNS]
- $B_{\alpha}(\phi)$ : Gaussian RF with  $\operatorname{Var}(B_{\lambda}(\phi)) = c(\phi)$  in (12). It is represented as integral  $B_{\alpha}(\phi) = \int_{\mathbb{R}^{d} \times \mathbb{R}_{+}} W_{\alpha}(\mathrm{d}\boldsymbol{u}, \mathrm{d}r) \int_{\mathbb{R}^{d}} \phi(\boldsymbol{t}) \mathbb{I}(\boldsymbol{t} \boldsymbol{u} \in r^{1/d} \Xi^{0}) \mathrm{d}\boldsymbol{t}$  w.r.t. Gaussian noise with  $\mathrm{E}W_{\alpha}(\mathrm{d}\boldsymbol{u}, \mathrm{d}r)^{2} = c_{f}r^{-1-\alpha}\mathrm{d}r\mathrm{d}\boldsymbol{u}$

# Theorem (1)

$$\lambda^{-H(\gamma)}(X_{\lambda,M}(\phi) - \mathrm{E}X_{\lambda,M}(\phi)) \xrightarrow{\mathrm{d}} \begin{cases} B_{\alpha}(\phi), \quad \gamma > d(\alpha - 1), \ H(\gamma) = \frac{\gamma + (3 - \alpha)d}{2}, \\ L_{\alpha}(\phi), \quad \gamma < d(\alpha - 1), \ H(\gamma) = \frac{\gamma + d}{\alpha}, \\ J_{\alpha}(\phi), \quad \gamma = d(\alpha - 1), \ H(\gamma) = d. \end{cases}$$

- Thm essentially due to [KLNS]
- $B_{\alpha}(\phi)$ : Gaussian RF with  $\operatorname{Var}(B_{\lambda}(\phi)) = c(\phi)$  in (12). It is represented as integral  $B_{\alpha}(\phi) = \int_{\mathbb{R}^{d} \times \mathbb{R}_{+}} W_{\alpha}(\mathrm{d}\boldsymbol{u}, \mathrm{d}r) \int_{\mathbb{R}^{d}} \phi(\boldsymbol{t}) \mathbb{I}(\boldsymbol{t} \boldsymbol{u} \in r^{1/d} \Xi^{0}) \mathrm{d}\boldsymbol{t}$  w.r.t. Gaussian noise with  $\mathrm{E}W_{\alpha}(\mathrm{d}\boldsymbol{u}, \mathrm{d}r)^{2} = c_{f}r^{-1-\alpha}\mathrm{d}r\mathrm{d}\boldsymbol{u}$
- $L_{\alpha}(\phi)$ :  $\alpha$ -stable RF written as stochastic integral  $L_{\alpha}(\phi) = \int_{\mathbb{R}^d} \phi(t) L_{\alpha}(dt)$  w.r.t.  $\alpha$ -stable random measure  $L_{\alpha}$

# Theorem (1)

$$\lambda^{-H(\gamma)}(X_{\lambda,M}(\phi) - \mathrm{E}X_{\lambda,M}(\phi)) \xrightarrow{\mathrm{d}} \begin{cases} B_{\alpha}(\phi), \quad \gamma > d(\alpha - 1), \ H(\gamma) = \frac{\gamma + (3 - \alpha)d}{2}, \\ L_{\alpha}(\phi), \quad \gamma < d(\alpha - 1), \ H(\gamma) = \frac{\gamma + d}{\alpha}, \\ J_{\alpha}(\phi), \quad \gamma = d(\alpha - 1), \ H(\gamma) = d. \end{cases}$$

- Thm essentially due to [KLNS]
- $B_{\alpha}(\phi)$ : Gaussian RF with  $\operatorname{Var}(B_{\lambda}(\phi)) = c(\phi)$  in (12). It is represented as integral  $B_{\alpha}(\phi) = \int_{\mathbb{R}^d \times \mathbb{R}_+} W_{\alpha}(\mathrm{d} \boldsymbol{u}, \mathrm{d} r) \int_{\mathbb{R}^d} \phi(\boldsymbol{t}) \mathbb{I}(\boldsymbol{t} \boldsymbol{u} \in r^{1/d} \Xi^0) \mathrm{d} \boldsymbol{t}$  w.r.t. Gaussian noise with  $\mathrm{E}W_{\alpha}(\mathrm{d} \boldsymbol{u}, \mathrm{d} r)^2 = c_f r^{-1-\alpha} \mathrm{d} r \mathrm{d} \boldsymbol{u}$
- L<sub>α</sub>(φ): α-stable RF written as stochastic integral L<sub>α</sub>(φ) = ∫<sub>ℝ<sup>d</sup></sub> φ(t)L<sub>α</sub>(dt) w.r.t. α-stable random measure L<sub>α</sub>
- $J_{\alpha}(\phi)$ : 'intermediate Poisson' RF written as stochastic integral  $J_{\alpha}(\phi) = \int_{\mathbb{R}^d \times \mathbb{R}_+} \tilde{N}_{\alpha}(\mathrm{d}\boldsymbol{u},\mathrm{d}r) \int_{\mathbb{R}^d} \phi(\boldsymbol{t}) \mathbb{I}(\boldsymbol{t} - \boldsymbol{u} \in r^{1/d} \Xi^0) \mathrm{d}\boldsymbol{t}$  w.r.t. centered Poisson random measure with variance  $\mathbb{E}\tilde{N}_{\alpha}(\mathrm{d}\boldsymbol{u},\mathrm{d}r)^2 = c_f r^{-1-\alpha} \mathrm{d}r \mathrm{d}\boldsymbol{u}$  (the same as  $W_{\alpha}$ )

Stochastic representations of B<sub>α</sub>(φ), L<sub>α</sub>(φ), J<sub>α</sub>(φ) mimic the structure of the original RG model except for a change of random measure

- Stochastic representations of  $B_{\alpha}(\phi), L_{\alpha}(\phi), J_{\alpha}(\phi)$  mimic the structure of the original RG model except for a change of random measure
- Thm obtains trichotomy of scaling limits of X<sub>M</sub>(φ) at γ<sub>0</sub> = d(α 1) (scaling transition)

- Stochastic representations of  $B_{\alpha}(\phi), L_{\alpha}(\phi), J_{\alpha}(\phi)$  mimic the structure of the original RG model except for a change of random measure
- Thm obtains trichotomy of scaling limits of X<sub>M</sub>(φ) at γ<sub>0</sub> = d(α 1) (scaling transition)
- For  $d = 1, \Xi^0 = [0, 1]$  (infinite source Poisson model) Thm agrees with:

Mikosch, T., Resnick, S., Rootzén, H. & Stegeman, A. (2002) Is network traffic approximated by stable Lévy motion or fractional Brownian motion? Ann. Appl. Probab. 12

Kaj, I. & Taqqu, M.S. (2008) Convergence to fractional Brownian motion and to the Telecom process: the integral representation approach. In: M.E. Vares and V. Sidoravicius (Eds.) In and Out of Equilibrium 2. Progress in Probability, vol. 60

Leipus, R., Pilipauskaitė, V. & S.D. (2023) Aggregation of network traffic and anisotropic scaling of random fields.

Th. Probab. Math. Statist.

- Stochastic representations of  $B_{\alpha}(\phi), L_{\alpha}(\phi), J_{\alpha}(\phi)$  mimic the structure of the original RG model except for a change of random measure
- Thm obtains trichotomy of scaling limits of X<sub>M</sub>(φ) at γ<sub>0</sub> = d(α 1) (scaling transition)
- For  $d = 1, \Xi^0 = [0, 1]$  (infinite source Poisson model) Thm agrees with:

Mikosch, T., Resnick, S., Rootzén, H. & Stegeman, A. (2002) Is network traffic approximated by stable Lévy motion or fractional Brownian motion? Ann. Appl. Probab. 12

Kaj, I. & Taqqu, M.S. (2008) Convergence to fractional Brownian motion and to the Telecom process: the integral representation approach. In: M.E. Vares and V. Sidoravicius (Eds.) In and Out of Equilibrium 2. Progress in Probability, vol. 60

Leipus, R., Pilipauskaitė, V. & S.D. (2023) Aggregation of network traffic and anisotropic scaling of random fields.

Th. Probab. Math. Statist.

Rather short proof using characteristic function

- Stochastic representations of  $B_{\alpha}(\phi), L_{\alpha}(\phi), J_{\alpha}(\phi)$  mimic the structure of the original RG model except for a change of random measure
- Thm obtains trichotomy of scaling limits of X<sub>M</sub>(φ) at γ<sub>0</sub> = d(α 1) (scaling transition)
- For  $d = 1, \Xi^0 = [0, 1]$  (infinite source Poisson model) Thm agrees with:

Mikosch, T., Resnick, S., Rootzén, H. & Stegeman, A. (2002) Is network traffic approximated by stable Lévy motion or fractional Brownian motion? Ann. Appl. Probab. 12

Kaj, I. & Taqqu, M.S. (2008) Convergence to fractional Brownian motion and to the Telecom process: the integral representation approach. In: M.E. Vares and V. Sidoravicius (Eds.) In and Out of Equilibrium 2. Progress in Probability, vol. 60

Leipus, R., Pilipauskaitė, V. & S.D. (2023) Aggregation of network traffic and anisotropic scaling of random fields.

Th. Probab. Math. Statist.

- Rather short proof using characteristic function
- Intuitive explanation of different limits for  $\gamma > \gamma_0$  and  $\gamma < \gamma_0$ :

- Stochastic representations of B<sub>α</sub>(φ), L<sub>α</sub>(φ), J<sub>α</sub>(φ) mimic the structure of the original RG model except for a change of random measure
- Thm obtains trichotomy of scaling limits of X<sub>M</sub>(φ) at γ<sub>0</sub> = d(α 1) (scaling transition)
- For  $d = 1, \Xi^0 = [0, 1]$  (infinite source Poisson model) Thm agrees with:

Mikosch, T., Resnick, S., Rootzén, H. & Stegeman, A. (2002) Is network traffic approximated by stable Lévy motion or fractional Brownian motion? Ann. Appl. Probab. 12

Kaj, I. & Taqqu, M.S. (2008) Convergence to fractional Brownian motion and to the Telecom process: the integral representation approach. In: M.E. Vares and V. Sidoravicius (Eds.) In and Out of Equilibrium 2. Progress in Probability, vol. 60

Leipus, R., Pilipauskaitė, V. & S.D. (2023) Aggregation of network traffic and anisotropic scaling of random fields.

Th. Probab. Math. Statist.

- Rather short proof using characteristic function
- Intuitive explanation of different limits for  $\gamma > \gamma_0$  and  $\gamma < \gamma_0$ : For  $\phi(t) = \mathbb{I}(t \in ]0, 1]$ ,  $M = \lambda^{\gamma}$

$$X_M(\phi) \approx \sum_j Leb_d((\boldsymbol{u}_{j,M} + R_j \Xi^0) \cap ]0, \lambda]^d),$$

where  $\{\boldsymbol{u}_{j,M}\}$  is Poisson process with intensity  $M d\boldsymbol{u} = \lambda^{\gamma} d\boldsymbol{u}$ 

• Intuition (ctnd): If random grain  $\boldsymbol{u}_{j,M} + R_j \Xi^0 \subset ]0, \lambda]^d$  then the corresponding  $Leb_d((\boldsymbol{u}_{j,M} + R_j^{1/d} \Xi^0) \cap ]0, \lambda]^d) = Leb_d(R_j^{1/d} \Xi^0) \propto R_j$  which are independent and  $\alpha$ -tailed r.v.

• Intuition (ctnd): If random grain  $\boldsymbol{u}_{j,M} + R_j \Xi^0 \subset ]0, \lambda]^d$  then the corresponding  $Leb_d((\boldsymbol{u}_{j,M} + R_j^{1/d} \Xi^0) \cap ]0, \lambda]^d) = Leb_d(R_j^{1/d} \Xi^0) \propto R_j$  which are independent and  $\alpha$ -tailed r.v.

The number of Poisson points  $u_{j,M} \subset ]0, \lambda]^d$  grows as  $M\lambda^d = \lambda^{\gamma+d}$ 

Intuition (ctnd): If random grain u<sub>j,M</sub> + R<sub>j</sub>Ξ<sup>0</sup> ⊂]0, λ]<sup>d</sup> then the corresponding Leb<sub>d</sub>((u<sub>j,M</sub> + R<sub>j</sub><sup>1/d</sup>Ξ<sup>0</sup>)∩]0, λ]<sup>d</sup>) = Leb<sub>d</sub>(R<sub>j</sub><sup>1/d</sup>Ξ<sup>0</sup>) ∝ R<sub>j</sub> which are independent and α-tailed r.v.

The number of Poisson points  $oldsymbol{u}_{j,M}\subset ]0,\lambda]^d$  grows as  $M\lambda^d=\lambda^{\gamma+d}$ 

Therefore,  $X_M(\phi) \approx \sum_{j=1}^{\lambda^{d+\gamma}} R_j \wedge \lambda^d$  behaves as a sum of  $\alpha$ -tailed i.i.d. r.v.s 'truncated' at level  $\lambda^d$ 

Intuition (ctnd): If random grain u<sub>j,M</sub> + R<sub>j</sub>Ξ<sup>0</sup> ⊂]0, λ]<sup>d</sup> then the corresponding Leb<sub>d</sub>((u<sub>j,M</sub> + R<sub>j</sub><sup>1/d</sup>Ξ<sup>0</sup>)∩]0, λ]<sup>d</sup>) = Leb<sub>d</sub>(R<sub>j</sub><sup>1/d</sup>Ξ<sup>0</sup>) ∝ R<sub>j</sub> which are independent and α-tailed r.v.

The number of Poisson points  $oldsymbol{u}_{j,M}\subset ]0,\lambda]^d$  grows as  $M\lambda^d=\lambda^{\gamma+d}$ 

Therefore,  $X_M(\phi) \approx \sum_{j=1}^{\lambda^{d+\gamma}} R_j \wedge \lambda^d$  behaves as a sum of  $\alpha$ -tailed i.i.d. r.v.s 'truncated' at level  $\lambda^d$ 

Chakrabarty, A. and Samorodnitsky, G. (2012) Tails in a bounded world or, is a truncated heavy tail heavy or not? Stoch. Models 28

discussed limit distribution of sums of *n*  $\alpha$ -tailed r.v.s truncated at level  $c_n \rightarrow \infty$ 

Intuition (ctnd): If random grain u<sub>j,M</sub> + R<sub>j</sub>Ξ<sup>0</sup> ⊂]0, λ]<sup>d</sup> then the corresponding Leb<sub>d</sub>((u<sub>j,M</sub> + R<sub>j</sub><sup>1/d</sup>Ξ<sup>0</sup>)∩]0, λ]<sup>d</sup>) = Leb<sub>d</sub>(R<sub>j</sub><sup>1/d</sup>Ξ<sup>0</sup>) ∝ R<sub>j</sub> which are independent and α-tailed r.v.

The number of Poisson points  $oldsymbol{u}_{j,M}\subset ]0,\lambda]^d$  grows as  $M\lambda^d=\lambda^{\gamma+d}$ 

Therefore,  $X_M(\phi) \approx \sum_{j=1}^{\lambda^{d+\gamma}} R_j \wedge \lambda^d$  behaves as a sum of  $\alpha$ -tailed i.i.d. r.v.s 'truncated' at level  $\lambda^d$ 

Chakrabarty, A. and Samorodnitsky, G. (2012) Tails in a bounded world or, is a truncated heavy tail heavy or not? Stoch. Models 28

discussed limit distribution of sums of *n* lpha-tailed r.v.s truncated at level  $c_n 
ightarrow \infty$ 

• They show that the limit distribution of such sums is Gaussian for when  $c_n/n^{1/\alpha} \to 0$  and  $\alpha$ -stable when  $c_n/n^{1/\alpha} \to \infty$ 

Intuition (ctnd): If random grain u<sub>j,M</sub> + R<sub>j</sub>Ξ<sup>0</sup> ⊂]0, λ]<sup>d</sup> then the corresponding Leb<sub>d</sub>((u<sub>j,M</sub> + R<sub>j</sub><sup>1/d</sup>Ξ<sup>0</sup>)∩]0, λ]<sup>d</sup>) = Leb<sub>d</sub>(R<sub>j</sub><sup>1/d</sup>Ξ<sup>0</sup>) ∝ R<sub>j</sub> which are independent and α-tailed r.v.

The number of Poisson points  $oldsymbol{u}_{j,M}\subset ]0,\lambda]^d$  grows as  $M\lambda^d=\lambda^{\gamma+d}$ 

Therefore,  $X_M(\phi) \approx \sum_{j=1}^{\lambda^{d+\gamma}} R_j \wedge \lambda^d$  behaves as a sum of  $\alpha$ -tailed i.i.d. r.v.s 'truncated' at level  $\lambda^d$ 

Chakrabarty, A. and Samorodnitsky, G. (2012) Tails in a bounded world or, is a truncated heavy tail heavy or not? Stoch. Models 28

discussed limit distribution of sums of *n* lpha-tailed r.v.s truncated at level  $c_n 
ightarrow \infty$ 

• They show that the limit distribution of such sums is Gaussian for when  $c_n/n^{1/\alpha} \to 0$  and  $\alpha$ -stable when  $c_n/n^{1/\alpha} \to \infty$ The boundary truncation level  $c_n = n^{1/\alpha}$  results in 'intermediate' infinitely divisible limit

Intuition (ctnd): If random grain u<sub>j,M</sub> + R<sub>j</sub>Ξ<sup>0</sup> ⊂]0, λ]<sup>d</sup> then the corresponding Leb<sub>d</sub>((u<sub>j,M</sub> + R<sub>j</sub><sup>1/d</sup>Ξ<sup>0</sup>)∩]0, λ]<sup>d</sup>) = Leb<sub>d</sub>(R<sub>j</sub><sup>1/d</sup>Ξ<sup>0</sup>) ∝ R<sub>j</sub> which are independent and α-tailed r.v.

The number of Poisson points  $oldsymbol{u}_{j,M} \subset ]0,\lambda]^d$  grows as  $M\lambda^d = \lambda^{\gamma+d}$ 

Therefore,  $X_M(\phi) \approx \sum_{j=1}^{\lambda^{d+\gamma}} R_j \wedge \lambda^d$  behaves as a sum of  $\alpha$ -tailed i.i.d. r.v.s 'truncated' at level  $\lambda^d$ 

Chakrabarty, A. and Samorodnitsky, G. (2012) Tails in a bounded world or, is a truncated heavy tail heavy or not? Stoch. Models 28

discussed limit distribution of sums of *n* lpha-tailed r.v.s truncated at level  $c_n 
ightarrow \infty$ 

- They show that the limit distribution of such sums is Gaussian for when  $c_n/n^{1/\alpha} \to 0$  and  $\alpha$ -stable when  $c_n/n^{1/\alpha} \to \infty$ The boundary truncation level  $c_n = n^{1/\alpha}$  results in 'intermediate' infinitely divisible limit
- In our case  $n = \lambda^{d+\gamma}$ ,  $c_n = \lambda^d$ ,  $c_n/n^{1/\alpha} = \lambda^{d-\frac{d+\gamma}{\alpha}}$  and  $d \frac{d+\gamma}{\alpha} = 0$  is equivalent to  $\gamma = \gamma_0 = d(\alpha 1)$  exactly as in the above theorem.

3.1. (Classics:) Gaussian distribution, Hermite polynomials and Mehler's formula.

- 3.1. (Classics:) Gaussian distribution, Hermite polynomials and Mehler's formula.
  - Hermite polynomials H<sub>k</sub>(x), x ∈ ℝ, k ∈ ℕ related Z ∼ N(0, 1) are defined by power series

$$\mathrm{e}^{\mathrm{i}ux+u^2/2} = \sum_{k=0}^{\infty} \frac{(\mathrm{i}u)^k}{k!} H_k(x), \qquad x, u \in \mathbb{R}.$$

- 3.1. (Classics:) Gaussian distribution, Hermite polynomials and Mehler's formula.
  - Hermite polynomials H<sub>k</sub>(x), x ∈ ℝ, k ∈ ℕ related Z ∼ N(0, 1) are defined by power series

$$e^{iux+u^2/2} = \sum_{k=0}^{\infty} \frac{(iu)^k}{k!} H_k(x), \qquad x, u \in \mathbb{R}.$$
 (13)

The r.v.s  $H_k(Z)$ ,  $k \ge N$  are orthogonal:

 $\mathrm{E}H_k(Z)=0, \qquad \mathrm{E}H_k(Z)^2=k!, \quad \mathrm{E}H_k(Z)H_\ell(Z)=0, \quad k\neq \ell=0,1,\cdots.$ 

- 3.1. (Classics:) Gaussian distribution, Hermite polynomials and Mehler's formula.
  - Hermite polynomials H<sub>k</sub>(x), x ∈ ℝ, k ∈ ℕ related Z ∼ N(0, 1) are defined by power series

$$e^{iux+u^2/2} = \sum_{k=0}^{\infty} \frac{(iu)^k}{k!} H_k(x), \qquad x, u \in \mathbb{R}.$$
 (13)

The r.v.s  $H_k(Z)$ ,  $k \ge N$  are orthogonal:

 $EH_k(Z) = 0,$   $EH_k(Z)^2 = k!,$   $EH_k(Z)H_\ell(Z) = 0,$   $k \neq \ell = 0, 1, \cdots$ . Any  $G \in L^2$  can be expanded in Hermite polynomials:

 $G(x) = \sum_{k=0}^{\infty} \frac{h_G(k)}{k!} H_k(x), \qquad h_G(k) := \operatorname{E} G(Z) H_k(Z), \quad j = 0, 1, \cdots.$ 

- 3.1. (Classics:) Gaussian distribution, Hermite polynomials and Mehler's formula.
  - Hermite polynomials H<sub>k</sub>(x), x ∈ ℝ, k ∈ ℕ related Z ~ N(0, 1) are defined by power series

$$e^{iux+u^2/2} = \sum_{k=0}^{\infty} \frac{(iu)^k}{k!} H_k(x), \qquad x, u \in \mathbb{R}.$$
 (13)

The r.v.s  $H_k(Z)$ ,  $k \ge N$  are orthogonal:

 $EH_k(Z) = 0,$   $EH_k(Z)^2 = k!,$   $EH_k(Z)H_\ell(Z) = 0,$   $k \neq \ell = 0, 1, \cdots$ . Any  $G \in L^2$  can be expanded in Hermite polynomials:

 $G(x) = \sum_{k=0}^{\infty} \frac{h_G(k)}{k!} H_k(x), \qquad h_G(k) := EG(Z)H_k(Z), \quad j = 0, 1, \cdots.$  $h_G(k) \text{ are called Hermite coefficients of } G.$ 

- 3.1. (Classics:) Gaussian distribution, Hermite polynomials and Mehler's formula.
  - Hermite polynomials H<sub>k</sub>(x), x ∈ ℝ, k ∈ ℕ related Z ∼ N(0, 1) are defined by power series

$$e^{iux+u^2/2} = \sum_{k=0}^{\infty} \frac{(iu)^k}{k!} H_k(x), \qquad x, u \in \mathbb{R}.$$
 (13)

The r.v.s  $H_k(Z)$ ,  $k \ge N$  are orthogonal:

 $EH_k(Z) = 0,$   $EH_k(Z)^2 = k!,$   $EH_k(Z)H_\ell(Z) = 0,$   $k \neq \ell = 0, 1, \cdots$ . Any  $G \in L^2$  can be expanded in Hermite polynomials:

 $G(x) = \sum_{k=0}^{\infty} \frac{h_G(k)}{k!} H_k(x), \qquad h_G(k) := EG(Z)H_k(Z), \quad j = 0, 1, \cdots.$ 

 $h_G(k)$  are called Hermite coefficients of G. Note  $h_G(0) = EG(Z)$  and  $EG(Z)^2 = \sum_{k=0}^{\infty} \frac{h_G^2(k)}{k!}$ .

- 3.1. (Classics:) Gaussian distribution, Hermite polynomials and Mehler's formula.
  - Hermite polynomials H<sub>k</sub>(x), x ∈ ℝ, k ∈ ℕ related Z ∼ N(0, 1) are defined by power series

$$e^{iux+u^2/2} = \sum_{k=0}^{\infty} \frac{(iu)^k}{k!} H_k(x), \qquad x, u \in \mathbb{R}.$$
 (13)

The r.v.s  $H_k(Z)$ ,  $k \ge N$  are orthogonal:

 $EH_k(Z) = 0,$   $EH_k(Z)^2 = k!,$   $EH_k(Z)H_\ell(Z) = 0,$   $k \neq \ell = 0, 1, \cdots$ . Any  $G \in L^2$  can be expanded in Hermite polynomials:

 $G(x) = \sum_{k=0}^{\infty} \frac{h_G(k)}{k!} H_k(x), \qquad h_G(k) := EG(Z)H_k(Z), \quad j = 0, 1, \cdots.$  $h_G(k) \text{ are called Hermite coefficients of } G. \text{ Note } h_G(0) = EG(Z) \text{ and}$  $EG(Z)^2 = \sum_{k=0}^{\infty} \frac{h_G^2(k)}{k!}.$ 

 Let (Z<sub>1</sub>, Z<sub>2</sub>) have bivariate normal distribution with mean zero, unit variances and correlation coefficient ρ ∈ (−1, 1), with the joint density

$$\phi(x,y) = (2\pi\sqrt{1-
ho^2})^{-1} \exp\left\{-\frac{1}{2(1-
ho^2)}(x^2+y^2-2
ho xy)
ight\} \quad x,y \in \mathbb{R}.$$

and marginal density  $\phi(x) = (2\pi)^{-1/2} e^{-x^2/2}$ . Then:

- 3.1. (Classics:) Gaussian distribution, Hermite polynomials and Mehler's formula.
  - Hermite polynomials H<sub>k</sub>(x), x ∈ ℝ, k ∈ ℕ related Z ∼ N(0, 1) are defined by power series

$$e^{iux+u^2/2} = \sum_{k=0}^{\infty} \frac{(iu)^k}{k!} H_k(x), \qquad x, u \in \mathbb{R}.$$
 (13)

The r.v.s  $H_k(Z)$ ,  $k \ge N$  are orthogonal:

 $EH_k(Z) = 0,$   $EH_k(Z)^2 = k!,$   $EH_k(Z)H_\ell(Z) = 0,$   $k \neq \ell = 0, 1, \cdots$ . Any  $G \in L^2$  can be expanded in Hermite polynomials:

 $G(x) = \sum_{k=0}^{\infty} \frac{h_G(k)}{k!} H_k(x), \qquad h_G(k) := EG(Z)H_k(Z), \quad j = 0, 1, \cdots.$  $h_G(k) \text{ are called Hermite coefficients of } G. \text{ Note } h_G(0) = EG(Z) \text{ and}$  $EG(Z)^2 = \sum_{k=0}^{\infty} \frac{h_G^2(k)}{k!}.$ 

 Let (Z<sub>1</sub>, Z<sub>2</sub>) have bivariate normal distribution with mean zero, unit variances and correlation coefficient ρ ∈ (−1, 1), with the joint density

$$\phi(x,y) = (2\pi\sqrt{1-
ho^2})^{-1} \exp\left\{-\frac{1}{2(1-
ho^2)}(x^2+y^2-2
ho xy)
ight\} \quad x,y \in \mathbb{R}.$$

and marginal density  $\phi(x) = (2\pi)^{-1/2} e^{-x^2/2}$ . Then:

(i) (Orthogonality property): For any  $k, \ell \in \mathbb{N}$ 

$$\mathbf{E}H_k(Z_1)H_\ell(Z_2) = \begin{cases} 0, & k \neq \ell, \\ \rho^k k!, & k = \ell, \end{cases}$$

(i) (Orthogonality property): For any  $k, \ell \in \mathbb{N}$ 

$$\mathbf{E}H_k(Z_1)H_\ell(Z_2) = \begin{cases} 0, & k \neq \ell, \\ \rho^k k!, & k = \ell, \end{cases}$$

(ii) Let  $G_i = G_i(x), x \in \mathbb{R}, i = 1, 2$  be given functions,  $EG_i^2(Z_i) < \infty, i = 1, 2$ . Then  $EG_1(Z_1)G_2(Z_2) = \sum_{k=0}^{\infty} \frac{h_{G_1}(k)h_{G_2}(k)}{k!}\rho^k.$  (14)

(iii) (Mehler's formula):

$$\phi(x,y) = \sum_{k=0}^{\infty} \frac{\rho^k}{k!} \phi^{(k)}(x) \phi^{(k)}(y) = \phi(x) \phi(y) \sum_{k=0}^{\infty} \frac{\rho^k}{k!} H_k(x) H_k(y).$$

(i) (Orthogonality property): For any  $k, \ell \in \mathbb{N}$ 

$$\mathbf{E}H_k(Z_1)H_\ell(Z_2) = \begin{cases} 0, & k \neq \ell, \\ \rho^k k!, & k = \ell, \end{cases}$$

(ii) Let  $G_i = G_i(x), x \in \mathbb{R}, i = 1, 2$  be given functions,  $\mathrm{E}G_i^2(Z_i) < \infty, i = 1, 2$ . Then

$$EG_1(Z_1)G_2(Z_2) = \sum_{k=0}^{\infty} \frac{h_{G_1}(k)h_{G_2}(k)}{k!} \rho^k.$$
 (14)

(iii) (Mehler's formula):

$$\phi(x,y) = \sum_{k=0}^{\infty} \frac{\rho^k}{k!} \phi^{(k)}(x) \phi^{(k)}(y) = \phi(x) \phi(y) \sum_{k=0}^{\infty} \frac{\rho^k}{k!} H_k(x) H_k(y).$$

**Proof of (i):** Multiply generating functions  $e^{iuZ_1+u^2/2}$  and  $e^{ivZ_2+v^2/2}$  and take expectation to obtain

$$\begin{split} \sum_{k,\ell=0}^{\infty} \frac{(\mathrm{i}u)^{k}(\mathrm{i}v)^{\ell}}{k!\ell!} \mathrm{E}H_{k}(Z_{1})H_{\ell}(Z_{2}) &= \mathrm{Ee}^{\mathrm{i}(uZ_{1}+vZ_{2})}\mathrm{e}^{(u^{2}+v^{2})/2} \\ &= \mathrm{e}^{-\rho uv} = \sum_{k=0}^{\infty} \frac{(-1)^{k}}{k!}\rho^{k}u^{k}v^{k}. \end{split}$$

(i) follows from this equality by comparing coefficients of powers  $u^k v^\ell$  on both sides.

(i) (Orthogonality property): For any  $k, \ell \in \mathbb{N}$ 

$$\mathbf{E}H_k(Z_1)H_\ell(Z_2) = \begin{cases} 0, & k \neq \ell, \\ \rho^k k!, & k = \ell, \end{cases}$$

(ii) Let  $G_i = G_i(x), x \in \mathbb{R}, i = 1, 2$  be given functions,  $\mathrm{E}G_i^2(Z_i) < \infty, i = 1, 2$ . Then

$$EG_1(Z_1)G_2(Z_2) = \sum_{k=0}^{\infty} \frac{h_{G_1}(k)h_{G_2}(k)}{k!} \rho^k.$$
 (14)

(iii) (Mehler's formula):

$$\phi(x,y) = \sum_{k=0}^{\infty} \frac{\rho^k}{k!} \phi^{(k)}(x) \phi^{(k)}(y) = \phi(x) \phi(y) \sum_{k=0}^{\infty} \frac{\rho^k}{k!} H_k(x) H_k(y).$$

**Proof of (i):** Multiply generating functions  $e^{iuZ_1+u^2/2}$  and  $e^{ivZ_2+v^2/2}$  and take expectation to obtain

$$\begin{split} \sum_{k,\ell=0}^{\infty} \frac{(\mathrm{i}u)^{k}(\mathrm{i}v)^{\ell}}{k!\ell!} \mathrm{E}H_{k}(Z_{1})H_{\ell}(Z_{2}) &= \mathrm{Ee}^{\mathrm{i}(uZ_{1}+vZ_{2})}\mathrm{e}^{(u^{2}+v^{2})/2} \\ &= \mathrm{e}^{-\rho uv} = \sum_{k=0}^{\infty} \frac{(-1)^{k}}{k!}\rho^{k}u^{k}v^{k}. \end{split}$$

(i) follows from this equality by comparing coefficients of powers  $u^k v^\ell$  on both sides.
**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ .

**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ . **Proof of (iii):** bivariate ch.f. of  $(Z_1, Z_2)$ :

$$\int_{\mathbb{R}^2} e^{i(xu+yv)} \phi(x,y) dx dy = e^{-(u^2 - 2\rho uv + v^2)/2}$$

**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ . **Proof of (iii):** bivariate ch.f. of  $(Z_1, Z_2)$ :

$$\int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \phi(x,y) \mathrm{d}x \mathrm{d}y = \mathrm{e}^{-(u^2 - 2\rho uv + v^2)/2}.$$

Show this equality remains valid with  $\phi(x, y)$  replaced by the r.h.s. of Mehler's formula, denoted by  $\tilde{\phi}(x, y)$ .

**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ . **Proof of (iii):** bivariate ch.f. of  $(Z_1, Z_2)$ :

$$\int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \phi(x,y) \mathrm{d}x \mathrm{d}y = \mathrm{e}^{-(u^2 - 2\rho uv + v^2)/2}.$$

Show this equality remains valid with  $\phi(x, y)$  replaced by the r.h.s. of Mehler's formula, denoted by  $\tilde{\phi}(x, y)$ . We have

 $I := \int_{\mathbb{R}^2} e^{i(xu+yv)} \tilde{\phi}(x,y) dx dy$ 

**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ . **Proof of (iii):** bivariate ch.f. of  $(Z_1, Z_2)$ :

$$\int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \phi(x,y) \mathrm{d}x \mathrm{d}y = \mathrm{e}^{-(u^2 - 2\rho uv + v^2)/2}.$$

Show this equality remains valid with  $\phi(x, y)$  replaced by the r.h.s. of Mehler's formula, denoted by  $\tilde{\phi}(x, y)$ . We have

$$I := \int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \tilde{\phi}(x,y) \mathrm{d}x \mathrm{d}y = \sum_{k=0}^{\infty} \frac{\rho^k}{k!} \int_{\mathbb{R}} \mathrm{e}^{\mathrm{i}xu} \phi^{(k)}(x) \mathrm{d}x \int_{\mathbb{R}} \mathrm{e}^{\mathrm{i}yv} \phi^{(k)}(y) \mathrm{d}y.$$

**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ . **Proof of (iii):** bivariate ch.f. of  $(Z_1, Z_2)$ :

$$\int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \phi(x,y) \mathrm{d}x \mathrm{d}y = \mathrm{e}^{-(u^2 - 2\rho uv + v^2)/2}.$$

Show this equality remains valid with  $\phi(x, y)$  replaced by the r.h.s. of Mehler's formula, denoted by  $\tilde{\phi}(x, y)$ . We have

$$I := \int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \tilde{\phi}(x,y) \mathrm{d}x \mathrm{d}y = \sum_{k=0}^{\infty} \frac{\rho^k}{k!} \int_{\mathbb{R}} \mathrm{e}^{\mathrm{i}xu} \phi^{(k)}(x) \mathrm{d}x \int_{\mathbb{R}} \mathrm{e}^{\mathrm{i}yv} \phi^{(k)}(y) \mathrm{d}y.$$

Integrating by parts,  $\int_{\mathbb{R}} e^{ixu} \phi^{(k)}(x) dx = (iu)^k \int_{\mathbb{R}} e^{ixu} \phi(x) dx = (iu)^k e^{-u^2/2}$ , hence

$$I = e^{-(u^2 + v^2)/2} \sum_{k=0}^{\infty} \frac{\rho^k}{k!} (iu)^k (iv)^k = e^{-(u^2 - 2\rho uv + v^2)/2}$$

completing the proof of (iii).

**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ . **Proof of (iii):** bivariate ch.f. of  $(Z_1, Z_2)$ :

$$\int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \phi(x,y) \mathrm{d}x \mathrm{d}y = \mathrm{e}^{-(u^2 - 2\rho uv + v^2)/2}.$$

Show this equality remains valid with  $\phi(x, y)$  replaced by the r.h.s. of Mehler's formula, denoted by  $\tilde{\phi}(x, y)$ . We have

$$I := \int_{\mathbb{R}^2} e^{i(xu+yv)} \tilde{\phi}(x,y) dx dy = \sum_{k=0}^{\infty} \frac{\rho^k}{k!} \int_{\mathbb{R}} e^{ixu} \phi^{(k)}(x) dx \int_{\mathbb{R}} e^{iyv} \phi^{(k)}(y) dy$$

Integrating by parts,  $\int_{\mathbb{R}} e^{ixu} \phi^{(k)}(x) dx = (iu)^k \int_{\mathbb{R}} e^{ixu} \phi(x) dx = (iu)^k e^{-u^2/2}$ , hence

$$I = e^{-(u^2 + v^2)/2} \sum_{k=0}^{\infty} \frac{\rho^k}{k!} (iu)^k (iv)^k = e^{-(u^2 - 2\rho uv + v^2)/2}$$

completing the proof of (iii).

 Mehler's formula: classical tool in mathematics & physics (O-U evolution, harmonic oscillators ...)

**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ . **Proof of (iii):** bivariate ch.f. of  $(Z_1, Z_2)$ :

$$\int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \phi(x,y) \mathrm{d}x \mathrm{d}y = \mathrm{e}^{-(u^2 - 2\rho uv + v^2)/2}.$$

Show this equality remains valid with  $\phi(x, y)$  replaced by the r.h.s. of Mehler's formula, denoted by  $\tilde{\phi}(x, y)$ . We have

$$I := \int_{\mathbb{R}^2} e^{i(xu+yv)} \tilde{\phi}(x,y) dx dy = \sum_{k=0}^{\infty} \frac{\rho^k}{k!} \int_{\mathbb{R}} e^{ixu} \phi^{(k)}(x) dx \int_{\mathbb{R}} e^{iyv} \phi^{(k)}(y) dy$$

Integrating by parts,  $\int_{\mathbb{R}} e^{ixu} \phi^{(k)}(x) dx = (iu)^k \int_{\mathbb{R}} e^{ixu} \phi(x) dx = (iu)^k e^{-u^2/2}$ , hence

$$I = e^{-(u^2 + v^2)/2} \sum_{k=0}^{\infty} \frac{\rho^k}{k!} (iu)^k (iv)^k = e^{-(u^2 - 2\rho uv + v^2)/2}$$

completing the proof of (iii).

 Mehler's formula: classical tool in mathematics & physics (O-U evolution, harmonic oscillators ...) Google search: > 300,000

**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ . **Proof of (iii):** bivariate ch.f. of  $(Z_1, Z_2)$ :

$$\int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \phi(x,y) \mathrm{d}x \mathrm{d}y = \mathrm{e}^{-(u^2 - 2\rho uv + v^2)/2}.$$

Show this equality remains valid with  $\phi(x, y)$  replaced by the r.h.s. of Mehler's formula, denoted by  $\tilde{\phi}(x, y)$ . We have

$$I := \int_{\mathbb{R}^2} e^{i(xu+yv)} \tilde{\phi}(x,y) dx dy = \sum_{k=0}^{\infty} \frac{\rho^k}{k!} \int_{\mathbb{R}} e^{ixu} \phi^{(k)}(x) dx \int_{\mathbb{R}} e^{iyv} \phi^{(k)}(y) dy$$

Integrating by parts,  $\int_{\mathbb{R}} e^{ixu} \phi^{(k)}(x) dx = (iu)^k \int_{\mathbb{R}} e^{ixu} \phi(x) dx = (iu)^k e^{-u^2/2}$ , hence

$$I = e^{-(u^2 + v^2)/2} \sum_{k=0}^{\infty} \frac{\rho^k}{k!} (iu)^k (iv)^k = e^{-(u^2 - 2\rho uv + v^2)/2}$$

completing the proof of (iii).

 Mehler's formula: classical tool in mathematics & physics (O-U evolution, harmonic oscillators ...) Google search: > 300,000

**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ . **Proof of (iii):** bivariate ch.f. of  $(Z_1, Z_2)$ :

$$\int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \phi(x,y) \mathrm{d}x \mathrm{d}y = \mathrm{e}^{-(u^2 - 2\rho uv + v^2)/2}.$$

Show this equality remains valid with  $\phi(x, y)$  replaced by the r.h.s. of Mehler's formula, denoted by  $\tilde{\phi}(x, y)$ . We have

$$I := \int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \tilde{\phi}(x,y) \mathrm{d}x \mathrm{d}y = \sum_{k=0}^{\infty} \frac{\rho^k}{k!} \int_{\mathbb{R}} \mathrm{e}^{\mathrm{i}xu} \phi^{(k)}(x) \mathrm{d}x \int_{\mathbb{R}} \mathrm{e}^{\mathrm{i}yv} \phi^{(k)}(y) \mathrm{d}y.$$

Integrating by parts,  $\int_{\mathbb{R}} e^{ixu} \phi^{(k)}(x) dx = (iu)^k \int_{\mathbb{R}} e^{ixu} \phi(x) dx = (iu)^k e^{-u^2/2}$ , hence

$$I = e^{-(u^2 + v^2)/2} \sum_{k=0}^{\infty} \frac{\rho^k}{k!} (iu)^k (iv)^k = e^{-(u^2 - 2\rho uv + v^2)/2}$$

completing the proof of (iii).

- Mehler's formula: classical tool in mathematics & physics (O-U evolution, harmonic oscillators ...) Google search: > 300,000
- Hermite rank  $k_H(G)$  of an G: the index of the first non-zero coefficient  $h_G(k)$  in the Hermite expansion (14):  $G(x) EG(Z) = \sum_{k=k_H(G)}^{\infty} h_G(k)H_k(x)/k!$

**Proof of (ii):** immediate from (i) and  $G_i(Z_i) = \sum_{j=0}^{\infty} \frac{h_{G_i}(j)}{j!} H_j(Z_i)$ . **Proof of (iii):** bivariate ch.f. of  $(Z_1, Z_2)$ :

$$\int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \phi(x,y) \mathrm{d}x \mathrm{d}y = \mathrm{e}^{-(u^2 - 2\rho uv + v^2)/2}.$$

Show this equality remains valid with  $\phi(x, y)$  replaced by the r.h.s. of Mehler's formula, denoted by  $\tilde{\phi}(x, y)$ . We have

$$I := \int_{\mathbb{R}^2} \mathrm{e}^{\mathrm{i}(xu+yv)} \tilde{\phi}(x,y) \mathrm{d}x \mathrm{d}y = \sum_{k=0}^{\infty} \frac{\rho^k}{k!} \int_{\mathbb{R}} \mathrm{e}^{\mathrm{i}xu} \phi^{(k)}(x) \mathrm{d}x \int_{\mathbb{R}} \mathrm{e}^{\mathrm{i}yv} \phi^{(k)}(y) \mathrm{d}y.$$

Integrating by parts,  $\int_{\mathbb{R}} e^{ixu} \phi^{(k)}(x) dx = (iu)^k \int_{\mathbb{R}} e^{ixu} \phi(x) dx = (iu)^k e^{-u^2/2}$ , hence

$$I = e^{-(u^2 + v^2)/2} \sum_{k=0}^{\infty} \frac{\rho^k}{k!} (iu)^k (iv)^k = e^{-(u^2 - 2\rho uv + v^2)/2}$$

completing the proof of (iii).

- Mehler's formula: classical tool in mathematics & physics (O-U evolution, harmonic oscillators ...) Google search: > 300,000
- Hermite rank  $k_H(G)$  of an G: the index of the first non-zero coefficient  $h_G(k)$  in the Hermite expansion (14):  $G(x) EG(Z) = \sum_{k=k_H(G)}^{\infty} h_G(k)H_k(x)/k!$

• (ii) and Cauchy-Schwarz imply

 $\operatorname{Cov}(G_1(Z_1), G_2(Z_2)) \le |\rho|^{k_H^*} \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2} \le |\rho| \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2}$ 

where  $k_H^* := k_H(G_1) \vee k_H(G_2) \ge 1$ ,  $\rho = \mathbb{E}Z_1Z_2 = \text{correlation coefficient}$ 

• (ii) and Cauchy-Schwarz imply

 $\operatorname{Cov}(G_1(Z_1), G_2(Z_2)) \le |\rho|^{k_H^*} \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2} \le |\rho| \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2}$ 

where  $k_H^* := k_H(G_1) \lor k_H(G_2) \ge 1$ ,  $\rho = \mathbb{E}Z_1Z_2 = \text{correlation coefficient}$ 

Arcones, M.A. (1994) Limit theorems for nonlinear functionals of a stationary Gaussian sequence of vectors, Ann. Probab. 22, 2242–2274.

Gebelein, H. (1941) Das statistische Problem der Korrelation als Variations- und Eigenwertproblem und sein Zusammenhang mit der Ausgleichsrechnung. Z. Angew. Math. Mech. 21.

• (ii) and Cauchy-Schwarz imply

 $\operatorname{Cov}(G_1(Z_1), G_2(Z_2)) \le |\rho|^{k_H^*} \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2} \le |\rho| \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2}$ 

where  $k_H^* := k_H(G_1) \lor k_H(G_2) \ge 1$ ,  $\rho = \mathbb{E}Z_1Z_2 = \text{correlation coefficient}$ 

Arcones, M.A. (1994) Limit theorems for nonlinear functionals of a stationary Gaussian sequence of vectors, Ann. Probab. 22, 2242–2274.

Gebelein, H. (1941) Das statistische Problem der Korrelation als Variations- und Eigenwertproblem und sein Zusammenhang mit der Ausgleichsrechnung, Z. Angew. Math. Mech. 21.

• If *G* has Hermite rank  $k_H(G) = 1$  then (linear) Gaussian RF  $\{G(X(t)), t \in \mathbb{R}^d\}$ and (nonlinear) Gaussian subordinated RF  $Y(t) := G(X(t)), t \in \mathbb{R}^d$  have the same LRD properties and scaling limits:

• (ii) and Cauchy-Schwarz imply

 $\operatorname{Cov}(G_1(Z_1), G_2(Z_2)) \le |\rho|^{k_H^*} \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2} \le |\rho| \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2}$ 

where  $k_H^* := k_H(G_1) \lor k_H(G_2) \ge 1$ ,  $\rho = \mathbb{E}Z_1Z_2 = \text{correlation coefficient}$ 

Arcones, M.A. (1994) Limit theorems for nonlinear functionals of a stationary Gaussian sequence of vectors, Ann. Probab. 22, 2242–2274.

Gebelein, H. (1941) Das statistische Problem der Korrelation als Variations- und Eigenwertproblem und sein Zusammenhang mit der Ausgleichsrechnung. Z. Angew. Math. Mech. 21.

If G has Hermite rank k<sub>H</sub>(G) = 1 then (linear) Gaussian RF {G(X(t)), t ∈ ℝ<sup>d</sup>} and (nonlinear) Gaussian subordinated RF Y(t) := G(X(t)), t ∈ ℝ<sup>d</sup> have the same LRD properties and scaling limits:
 Y(t) = h<sub>G</sub>(1)X(t) + Y<sup>\*</sup>(t) where 'remainder' Y<sup>\*</sup>(t) is negligible

• (ii) and Cauchy-Schwarz imply

 $\operatorname{Cov}(G_1(Z_1), G_2(Z_2)) \le |\rho|^{k_{\mathcal{H}}^*} \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2} \le |\rho| \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2}$ 

where  $k_H^* := k_H(G_1) \lor k_H(G_2) \ge 1$ ,  $\rho = \mathbb{E}Z_1Z_2 = \text{correlation coefficient}$ 

Arcones, M.A. (1994) Limit theorems for nonlinear functionals of a stationary Gaussian sequence of vectors, Ann. Probab. 22, 2242–2274.

Gebelein, H. (1941) Das statistische Problem der Korrelation als Variations- und Eigenwertproblem und sein Zusammenhang mit der Ausgleichsrechnung, Z. Angew. Math. Mech. 21.

- If G has Hermite rank  $k_H(G) = 1$  then (linear) Gaussian RF  $\{G(X(t)), t \in \mathbb{R}^d\}$ and (nonlinear) Gaussian subordinated RF  $Y(t) := G(X(t)), t \in \mathbb{R}^d$  have the same LRD properties and scaling limits:  $Y(t) = h_G(1)X(t) + Y^*(t)$  where 'remainder'  $Y^*(t)$  is negligible
- Dobrushin-Major-Taqqu theory treats the general case of LRD Gaussian subordinated RF Y(t) = G(X(t)), t ∈ ℝ<sup>d</sup> of arbitrary Hermite rank k<sub>H</sub>(G) ≥ 1

• (ii) and Cauchy-Schwarz imply

 $\operatorname{Cov}(G_1(Z_1), G_2(Z_2)) \le |\rho|^{k_{\mathcal{H}}^*} \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2} \le |\rho| \prod_{i=1}^2 \operatorname{Var}(G(Z_i))^{1/2}$ 

where  $k_H^* := k_H(G_1) \lor k_H(G_2) \ge 1$ ,  $\rho = \mathbb{E}Z_1Z_2 = \text{correlation coefficient}$ 

Arcones, M.A. (1994) Limit theorems for nonlinear functionals of a stationary Gaussian sequence of vectors, Ann. Probab. 22, 2242–2274.

Gebelein, H. (1941) Das statistische Problem der Korrelation als Variations- und Eigenwertproblem und sein Zusammenhang mit der Ausgleichsrechnung, Z. Angew. Math. Mech. 21.

- If G has Hermite rank k<sub>H</sub>(G) = 1 then (linear) Gaussian RF {G(X(t)), t ∈ ℝ<sup>d</sup>} and (nonlinear) Gaussian subordinated RF Y(t) := G(X(t)), t ∈ ℝ<sup>d</sup> have the same LRD properties and scaling limits:
   Y(t) = h<sub>G</sub>(1)X(t) + Y<sup>\*</sup>(t) where 'remainder' Y<sup>\*</sup>(t) is negligible
- Dobrushin-Major-Taqqu theory treats the general case of LRD Gaussian subordinated RF Y(t) = G(X(t)), t ∈ ℝ<sup>d</sup> of arbitrary Hermite rank k<sub>H</sub>(G) ≥ 1 (scaling limits written through Gaussian polynomial chaos or multiple Wiener-Itô integrals)

3.2. Poisson distribution & Charlier polynomials

### 3.2. Poisson distribution & Charlier polynomials

• N = Poisson r.v. with mean  $\mu = EN$  and distribution  $p(x; \mu) = e^{-\mu \frac{\mu^x}{x!}}, x \in \mathbb{N}$ 

#### 3.2. Poisson distribution & Charlier polynomials

- N = Poisson r.v. with mean  $\mu = EN$  and distribution  $p(x; \mu) = e^{-\mu \frac{\mu^{x}}{x!}}, x \in \mathbb{N}$
- Charlier polynomials P<sub>k</sub>(x; μ) of discrete variable x ∈ N defined through generating function:

$$\mathcal{P}(u; x, \mu) := \sum_{k=0}^{\infty} \frac{u^k}{k!} P_k(x; \mu) = (1+u)^x \mathrm{e}^{-u\mu},$$

#### 3.2. Poisson distribution & Charlier polynomials

- N = Poisson r.v. with mean  $\mu = EN$  and distribution  $p(x; \mu) = e^{-\mu} \frac{\mu^x}{x!}, x \in \mathbb{N}$
- Charlier polynomials P<sub>k</sub>(x; μ) of discrete variable x ∈ N defined through generating function:

$$\mathcal{P}(u; x, \mu) := \sum_{k=0}^{\infty} \frac{u^k}{k!} P_k(x; \mu) = (1+u)^x e^{-u\mu},$$
(15)

 $P_0(x;\mu) = 1, P_1(x;\mu) = x - \mu, P_2(x;\mu) = x^2 - (2\mu + 1)x + \mu^2,$ 

#### 3.2. Poisson distribution & Charlier polynomials

- N = Poisson r.v. with mean  $\mu = EN$  and distribution  $p(x; \mu) = e^{-\mu \frac{\mu^x}{x!}}, x \in \mathbb{N}$
- Charlier polynomials P<sub>k</sub>(x; μ) of discrete variable x ∈ N defined through generating function:

$$\mathcal{P}(u; x, \mu) := \sum_{k=0}^{\infty} \frac{u^k}{k!} P_k(x; \mu) = (1+u)^x \mathrm{e}^{-u\mu}, \tag{15}$$

 $P_{0}(x;\mu) = 1, P_{1}(x;\mu) = x - \mu, P_{2}(x;\mu) = x^{2} - (2\mu + 1)x + \mu^{2},$  $P_{k}(x;\mu) = (-1)^{k} \mu^{k} p(x;\mu)^{-1} D_{-}^{k} p(x;\mu), \quad k \in \mathbb{N}$ (16)

where  $D_{-}^{k} := D_{-}D_{-}^{k-1}$  is the backward difference operator,  $D_{-}G(x) := G(x) - G(x-1)\mathbb{I}(x \ge 1), D_{-}^{0}G(x) = G(x)$ 

Orthogonality relations for Charlier polynomials

$$\mathrm{E}P_k(N;\mu) = 0, \quad \mathrm{E}P_k(N)^2 = k!\mu^k, \quad \mathrm{E}P_k(N;\mu)P_\ell(N;\mu) = 0, \quad k \neq \ell$$

follow from multiplying the series in (15) at the points u and v and taking the expectation of the product:

$$\sum_{k,\ell=0}^{\infty} \frac{u^{k} v^{\ell}}{k!\ell!} \mathbf{E} P_{k}(N; \mu) P_{\ell}(N; \mu) = e^{-(u+v)\mu} \mathbf{E}[((1+u)(1+v))^{N}]$$

#### 3.2. Poisson distribution & Charlier polynomials

- N = Poisson r.v. with mean  $\mu = EN$  and distribution  $p(x; \mu) = e^{-\mu \frac{\mu^x}{x!}}, x \in \mathbb{N}$
- Charlier polynomials P<sub>k</sub>(x; μ) of discrete variable x ∈ N defined through generating function:

$$\mathcal{P}(u; x, \mu) := \sum_{k=0}^{\infty} \frac{u^k}{k!} P_k(x; \mu) = (1+u)^x \mathrm{e}^{-u\mu}, \tag{15}$$

$$P_{0}(x;\mu) = 1, P_{1}(x;\mu) = x - \mu, P_{2}(x;\mu) = x^{2} - (2\mu + 1)x + \mu^{2},$$
  

$$P_{k}(x;\mu) = (-1)^{k} \mu^{k} p(x;\mu)^{-1} D_{-}^{k} p(x;\mu), \quad k \in \mathbb{N}$$
(16)

where  $D_{-}^{k} := D_{-}D_{-}^{k-1}$  is the backward difference operator,  $D_{-}G(x) := G(x) - G(x-1)\mathbb{I}(x \ge 1), D_{-}^{0}G(x) = G(x)$ 

Orthogonality relations for Charlier polynomials

$$\mathrm{E}P_k(N;\mu) = 0, \quad \mathrm{E}P_k(N)^2 = k!\mu^k, \quad \mathrm{E}P_k(N;\mu)P_\ell(N;\mu) = 0, \quad k \neq \ell$$

follow from multiplying the series in (15) at the points u and v and taking the expectation of the product:

$$\sum_{k,\ell=0}^{\infty} \frac{u^k v^\ell}{k!\ell!} \operatorname{E} P_k(N;\mu) P_\ell(N;\mu) = \operatorname{e}^{-(u+\nu)\mu} \operatorname{E}[((1+u)(1+\nu))^N]$$
$$= \operatorname{e}^{\mu u \nu} = \sum_{k=0}^{\infty} \frac{(\mu u \nu)^k}{k!}$$

and equating the coefficients of  $u^k v^{\ell}$ ,  $k, \ell \in \mathbb{N}$  of the power series.

• Any  $G = G(x), x \in \mathbb{N}$  with  $\mathrm{E}G^2(N) < \infty$  can be expanded in Charlier polynomials

$$G(x) = \sum_{k=0}^{\infty} \frac{c_G(k)}{k!} P_k(x;\mu), \quad x \in \mathbb{N}$$
(17)

where

$$c_G(k) := \mu^{-k} \mathbb{E}G(N) P_k(N; \mu), \quad k \in \mathbb{N}$$
(18)

are Charlier coefficients of G in (17).

• Any  $G = G(x), x \in \mathbb{N}$  with  $\mathrm{E}G^2(N) < \infty$  can be expanded in Charlier polynomials

$$G(x) = \sum_{k=0}^{\infty} \frac{c_G(k)}{k!} P_k(x;\mu), \quad x \in \mathbb{N}$$
(17)

where

$$c_G(k) := \mu^{-k} \mathbb{E}G(N) P_k(N; \mu), \quad k \in \mathbb{N}$$
(18)

are Charlier coefficients of G in (17). Summation by parts yields

$$c_G(k) = \mathrm{E} D^k_+ G(N), \quad k \in \mathbb{N}, \tag{19}$$

where  $D_+G(x) := G(x+1) - G(x)$  is the forward difference.

• Any  $G = G(x), x \in \mathbb{N}$  with  $\mathrm{E}G^2(N) < \infty$  can be expanded in Charlier polynomials

$$G(x) = \sum_{k=0}^{\infty} \frac{c_G(k)}{k!} P_k(x;\mu), \quad x \in \mathbb{N}$$
(17)

where

$$c_G(k) := \mu^{-k} \mathbb{E}G(N) P_k(N; \mu), \quad k \in \mathbb{N}$$
(18)

are Charlier coefficients of G in (17). Summation by parts yields

$$c_G(k) = ED^k_+ G(N), \quad k \in \mathbb{N},$$
(19)

where  $D_+G(x) := G(x+1) - G(x)$  is the forward difference. (18) and (17) yield the bound

$$|c_G(k)| \le \mu^{-k} \sqrt{\mathbb{E}[G^2(N)]\mathbb{E}[P_k^2(N;\mu)]} = C(k!/\mu^k)^{1/2}, \quad C = \sqrt{\mathbb{E}G(N)^2}.$$
 (20)

• Any  $G = G(x), x \in \mathbb{N}$  with  $\mathbb{E}G^2(N) < \infty$  can be expanded in Charlier polynomials

$$G(x) = \sum_{k=0}^{\infty} \frac{c_G(k)}{k!} P_k(x;\mu), \quad x \in \mathbb{N}$$
(17)

where

$$c_G(k) := \mu^{-k} \mathbb{E}G(N) P_k(N; \mu), \quad k \in \mathbb{N}$$
(18)

are Charlier coefficients of G in (17). Summation by parts yields

$$c_G(k) = ED^k_+ G(N), \quad k \in \mathbb{N},$$
(19)

where  $D_+G(x) := G(x+1) - G(x)$  is the forward difference. (18) and (17) yield the bound

$$|c_G(k)| \le \mu^{-k} \sqrt{\mathbb{E}[G^2(N)]\mathbb{E}[P_k^2(N;\mu)]} = C(k!/\mu^k)^{1/2}, \quad C = \sqrt{\mathbb{E}G(N)^2}.$$
 (20)

 Charlier rank k<sub>C</sub>(G) of G: the index of the first non-zero coefficient c<sub>G</sub>(k), k ≥ 1 in the Charlier expansion (17)

• What is bivariate (multivariate) Poisson distribution with *dependent* components?

 What is bivariate (multivariate) Poisson distribution with *dependent* components? Or, when random vector (N<sub>1</sub>, N<sub>2</sub>) taking values in N<sup>2</sup> has joint Poisson distribution?

 What is bivariate (multivariate) Poisson distribution with *dependent* components? Or, when random vector (N<sub>1</sub>, N<sub>2</sub>) taking values in N<sup>2</sup> has joint Poisson distribution? Possible answer:

$$N_1 = M_1 + M_3, \quad N_1 = M_+ M_3$$
 (21)

where  $M_i$ , i = 1, 2, 3 are *independent* Poisson r.v.s with  $EM_i = \mu_i$ .

 What is bivariate (multivariate) Poisson distribution with *dependent* components? Or, when random vector (N<sub>1</sub>, N<sub>2</sub>) taking values in N<sup>2</sup> has joint Poisson distribution? Possible answer:

$$N_1 = M_1 + M_3, \quad N_1 = M_+ M_3$$
 (21)

where  $M_i$ , i = 1, 2, 3 are *independent* Poisson r.v.s with  $EM_i = \mu_i$ .

• Multivariate Poisson  $(N_1, \cdots, N_p) \in \mathbb{N}^p$ :

$$N_i = M(A_i), \quad i = 1, \cdots, p,$$

where M(dx) is Poisson random measure on measurable space  $(\mathcal{X}, \mu)$  and  $A_i \subset \mathcal{X}, \mu(A_i) < \infty$  are any subsets.
What is bivariate (multivariate) Poisson distribution with *dependent* components? Or, when random vector (N<sub>1</sub>, N<sub>2</sub>) taking values in N<sup>2</sup> has joint Poisson distribution? Possible answer:

$$N_1 = M_1 + M_3, \quad N_1 = M_+ M_3$$
 (21)

where  $M_i$ , i = 1, 2, 3 are *independent* Poisson r.v.s with  $EM_i = \mu_i$ .

• Multivariate Poisson  $(N_1, \cdots, N_p) \in \mathbb{N}^p$ :

$$N_i = M(A_i), \quad i = 1, \cdots, p,$$

where M(dx) is Poisson random measure on measurable space  $(\mathcal{X}, \mu)$  and  $A_i \subset \mathcal{X}, \mu(A_i) < \infty$  are any subsets. (nonconstructive?)

 What is bivariate (multivariate) Poisson distribution with *dependent* components? Or, when random vector (N<sub>1</sub>, N<sub>2</sub>) taking values in N<sup>2</sup> has joint Poisson distribution? Possible answer:

$$N_1 = M_1 + M_3, \quad N_1 = M_+ M_3$$
 (21)

where  $M_i$ , i = 1, 2, 3 are *independent* Poisson r.v.s with  $EM_i = \mu_i$ .

• Multivariate Poisson  $(N_1, \cdots, N_p) \in \mathbb{N}^p$ :

$$N_i = M(A_i), \quad i = 1, \cdots, p,$$

where M(dx) is Poisson random measure on measurable space  $(\mathcal{X}, \mu)$  and  $A_i \subset \mathcal{X}, \mu(A_i) < \infty$  are any subsets. (nonconstructive?)

Examples of random processes with multivariate Poisson distribution: random grain model, *trawl process with Poisson seed*:

Barndorff-Nielsen, O.E., Lunde, A., Shepard, N. & Veraart, A.E.D. (2014) Integer-valued trawl processes: a class of stationary infinitely divisible processes. Scand. J. Statist. 41, 693–724.

Doukhan, P., Jakubowski, A., Lopes, S.R.C. & S.D. (2019) Discrete-time trawl processes. Stoch. Proc. Appl. 129

Let  $M_i$ , i = 1, 2, 3, be *independent* Poisson r.v.s,  $\mu_i := EM_i$ , i = 1, 2, 3,  $\mu_1 = \mu_2$  and  $N_1 := M_1 + M_3$ ,  $N_2 := M_2 + M_3$ ,  $\mu := \mu_1 + \mu_3 = EN_1 = \mu_2 + \mu_3 = EN_2$ ,

Let  $M_i$ , i = 1, 2, 3, be *independent* Poisson r.v.s,  $\mu_i := EM_i$ , i = 1, 2, 3,  $\mu_1 = \mu_2$  and  $N_1 := M_1 + M_3$ ,  $N_2 := M_2 + M_3$ ,  $\mu := \mu_1 + \mu_3 = EN_1 = \mu_2 + \mu_3 = EN_2$ ,

$$p(x,y;\mu) := \mathrm{P}(N_1 = x, N_2 = y), \quad (x,y) \in \mathbb{N}^2$$

- joint distribution of  $(N_1, N_2)$ ;

Let  $M_i$ , i = 1, 2, 3, be *independent* Poisson r.v.s,  $\mu_i := EM_i$ , i = 1, 2, 3,  $\mu_1 = \mu_2$  and  $N_1 := M_1 + M_3$ ,  $N_2 := M_2 + M_3$ ,  $\mu := \mu_1 + \mu_3 = EN_1 = \mu_2 + \mu_3 = EN_2$ ,

$$p(x,y;\mu) := \mathrm{P}(N_1 = x, N_2 = y), \quad (x,y) \in \mathbb{N}^2$$

- joint distribution of  $(N_1, N_2)$ ;

 $\operatorname{Var}(N_1) = \operatorname{Var}(N_2) = \mu$ ,  $\operatorname{Cov}(N_1, N_2) = \mu_3$ ,  $\operatorname{Corr}(N_1, N_2) = \frac{\mu_3}{\mu}$ ;

Let  $M_i$ , i = 1, 2, 3, be *independent* Poisson r.v.s,  $\mu_i := EM_i$ , i = 1, 2, 3,  $\mu_1 = \mu_2$  and  $N_1 := M_1 + M_3$ ,  $N_2 := M_2 + M_3$ ,  $\mu := \mu_1 + \mu_3 = EN_1 = \mu_2 + \mu_3 = EN_2$ ,

$$p(x,y;\mu) := \mathrm{P}(N_1 = x, N_2 = y), \quad (x,y) \in \mathbb{N}^2$$

- joint distribution of  $(N_1, N_2)$ ;

 $\operatorname{Var}(N_1) = \operatorname{Var}(N_2) = \mu, \quad \operatorname{Cov}(N_1, N_2) = \mu_3, \quad \operatorname{Corr}(N_1, N_2) = \frac{\mu_3}{\mu};$ 

 $P_k(x; \mu), k \in \mathbb{N}$ : Charlier polynomials

Let  $M_i$ , i = 1, 2, 3, be *independent* Poisson r.v.s,  $\mu_i := EM_i$ , i = 1, 2, 3,  $\mu_1 = \mu_2$  and  $N_1 := M_1 + M_3$ ,  $N_2 := M_2 + M_3$ ,  $\mu := \mu_1 + \mu_3 = EN_1 = \mu_2 + \mu_3 = EN_2$ ,

$$p(x,y;\mu) := \mathrm{P}(N_1 = x, N_2 = y), \quad (x,y) \in \mathbb{N}^2$$

- joint distribution of  $(N_1, N_2)$ ;

 $\operatorname{Var}(N_1) = \operatorname{Var}(N_2) = \mu$ ,  $\operatorname{Cov}(N_1, N_2) = \mu_3$ ,  $\operatorname{Corr}(N_1, N_2) = \frac{\mu_3}{\mu}$ ;

 $P_k(x; \mu), k \in \mathbb{N}$ : Charlier polynomials

### Lemma (1)

(i) (orthogonality): For any  $k, \ell \in \mathbb{N}$ 

$$\mathbf{E} P_k(N_1;\mu) P_\ell(N_2;\mu) = \begin{cases} 0, & k \neq \ell, \\ \mu_3^k k!, & k = \ell, \end{cases}$$

Let  $M_i$ , i = 1, 2, 3, be *independent* Poisson r.v.s,  $\mu_i := \mathbb{E}M_i$ , i = 1, 2, 3,  $\mu_1 = \mu_2$  and  $N_1 := M_1 + M_3$ ,  $N_2 := M_2 + M_3$ ,  $\mu := \mu_1 + \mu_3 = \mathbb{E}N_1 = \mu_2 + \mu_3 = \mathbb{E}N_2$ ,

$$p(x,y;\mu) := \mathrm{P}(N_1 = x, N_2 = y), \quad (x,y) \in \mathbb{N}^2$$

- joint distribution of  $(N_1, N_2)$ ;

 $\operatorname{Var}(N_1) = \operatorname{Var}(N_2) = \mu$ ,  $\operatorname{Cov}(N_1, N_2) = \mu_3$ ,  $\operatorname{Corr}(N_1, N_2) = \frac{\mu_3}{\mu}$ ;

 $P_k(x; \mu), k \in \mathbb{N}$ : Charlier polynomials

### Lemma (1)

(i) (orthogonality): For any  $k, \ell \in \mathbb{N}$ 

$$\mathbb{E} P_k(N_1;\mu) P_\ell(N_2;\mu) = \begin{cases} 0, & k \neq \ell, \\ \mu_3^k k!, & k = \ell, \end{cases}$$

(ii) Let  $G_i = G_i(x), x \in \mathbb{N}, i = 1, 2, EG_i(N_i)^2 < \infty, i = 1, 2.$ 

Let  $M_i$ , i = 1, 2, 3, be *independent* Poisson r.v.s,  $\mu_i := EM_i$ , i = 1, 2, 3,  $\mu_1 = \mu_2$  and  $N_1 := M_1 + M_3$ ,  $N_2 := M_2 + M_3$ ,  $\mu := \mu_1 + \mu_3 = EN_1 = \mu_2 + \mu_3 = EN_2$ ,

$$p(x,y;\mu) := \mathrm{P}(N_1 = x, N_2 = y), \quad (x,y) \in \mathbb{N}^2$$

- joint distribution of  $(N_1, N_2)$ ;

 $\operatorname{Var}(N_1) = \operatorname{Var}(N_2) = \mu$ ,  $\operatorname{Cov}(N_1, N_2) = \mu_3$ ,  $\operatorname{Corr}(N_1, N_2) = \frac{\mu_3}{\mu}$ ;

 $P_k(x; \mu), k \in \mathbb{N}$ : Charlier polynomials

### Lemma (1)

(i) (orthogonality): For any  $k, \ell \in \mathbb{N}$ 

$$\mathbb{E} P_k(N_1;\mu) P_\ell(N_2;\mu) = \begin{cases} 0, & k \neq \ell, \\ \mu_3^k k!, & k = \ell, \end{cases}$$

(ii) Let  $G_i=G_i(x), x\in\mathbb{N}, i=1,2,$   $\mathrm{E}G_i(N_i)^2<\infty, i=1,2.$  Then

$$\mathrm{E}G_{1}(N_{1})G_{2}(N_{2}) = \sum_{k=0}^{\infty} \frac{c_{G_{1}}(k)c_{G_{2}}(k)}{k!} \mu_{3}^{k}$$

(iii) (Mehler's formula):

$$p(x, y; \mu) = \sum_{k=0}^{\infty} \frac{\mu_3^k}{k!} D_-^k p(x; \mu) D_-^k p(y; \mu)$$

(iii) (Mehler's formula):

$$p(x, y; \mu) = \sum_{k=0}^{\infty} \frac{\mu_3^k}{k!} D_-^k p(x; \mu) D_-^k p(y; \mu)$$
  
=  $p(x; \mu) p(y; \mu) \sum_{k=0}^{\infty} \frac{\mu_3^k}{\mu^{2k}k!} P_k(x; \mu) P_k(y; \mu).$ 

(iii) (Mehler's formula):

$$p(x, y; \mu) = \sum_{k=0}^{\infty} \frac{\mu_3^k}{k!} D_-^k p(x; \mu) D_-^k p(y; \mu)$$
  
=  $p(x; \mu) p(y; \mu) \sum_{k=0}^{\infty} \frac{\mu_3^k}{\mu^{2k} k!} P_k(x; \mu) P_k(y; \mu).$ 

 For μ = 1 (ii) and (iii) provide complete expansions of joint expectations and probabilities in powers μ<sub>3</sub><sup>k</sup> (μ<sub>3</sub> = Corr(N<sub>1</sub>, N<sub>2</sub>))

(iii) (Mehler's formula):

$$p(x, y; \mu) = \sum_{k=0}^{\infty} \frac{\mu_3^k}{k!} D_-^k p(x; \mu) D_-^k p(y; \mu)$$
  
=  $p(x; \mu) p(y; \mu) \sum_{k=0}^{\infty} \frac{\mu_3^k}{\mu^{2k}k!} P_k(x; \mu) P_k(y; \mu).$ 

- For μ = 1 (ii) and (iii) provide complete expansions of joint expectations and probabilities in powers μ<sub>3</sub><sup>k</sup> (μ<sub>3</sub> = Corr(N<sub>1</sub>, N<sub>2</sub>))
- Nearly complete analogy with Hermite expansions in the Gaussian case...

(iii) (Mehler's formula):

$$p(x, y; \mu) = \sum_{k=0}^{\infty} \frac{\mu_3^k}{k!} D_-^k p(x; \mu) D_-^k p(y; \mu)$$
  
=  $p(x; \mu) p(y; \mu) \sum_{k=0}^{\infty} \frac{\mu_3^k}{\mu^{2k}k!} P_k(x; \mu) P_k(y; \mu).$ 

- For μ = 1 (ii) and (iii) provide complete expansions of joint expectations and probabilities in powers μ<sub>3</sub><sup>k</sup> (μ<sub>3</sub> = Corr(N<sub>1</sub>, N<sub>2</sub>))
- Nearly complete analogy with Hermite expansions in the Gaussian case... (how far?)

(iii) (Mehler's formula):

$$p(x, y; \mu) = \sum_{k=0}^{\infty} \frac{\mu_3^k}{k!} D_-^k p(x; \mu) D_-^k p(y; \mu)$$
  
=  $p(x; \mu) p(y; \mu) \sum_{k=0}^{\infty} \frac{\mu_3^k}{\mu^{2k}k!} P_k(x; \mu) P_k(y; \mu).$ 

- For μ = 1 (ii) and (iii) provide complete expansions of joint expectations and probabilities in powers μ<sub>3</sub><sup>k</sup> (μ<sub>3</sub> = Corr(N<sub>1</sub>, N<sub>2</sub>))
- Nearly complete analogy with Hermite expansions in the Gaussian case... (how far?)
- Stationary Markov process  $\{N_t; t = 0, 1, \dots\}$  with marginal Poisson distribution  $P(N_t = x) = p(x; \mu)$  and transition probabilities  $P(N_{t+1} = y | N_t = x) = p(y|x; \mu)$

$$p(y|x;\mu) := rac{p(x,y;\mu)}{p(x;\mu)}, \quad x,y \in \mathbb{N}$$

is Poisson AR(1) or INAR(1) ('Poisson O-U').

(iii) (Mehler's formula):

$$p(x, y; \mu) = \sum_{k=0}^{\infty} \frac{\mu_3^k}{k!} D_-^k p(x; \mu) D_-^k p(y; \mu)$$
  
=  $p(x; \mu) p(y; \mu) \sum_{k=0}^{\infty} \frac{\mu_3^k}{\mu^{2k}k!} P_k(x; \mu) P_k(y; \mu).$ 

- For μ = 1 (ii) and (iii) provide complete expansions of joint expectations and probabilities in powers μ<sub>3</sub><sup>k</sup> (μ<sub>3</sub> = Corr(N<sub>1</sub>, N<sub>2</sub>))
- Nearly complete analogy with Hermite expansions in the Gaussian case... (how far?)
- Stationary Markov process  $\{N_t; t = 0, 1, \dots\}$  with marginal Poisson distribution  $P(N_t = x) = p(x; \mu)$  and transition probabilities  $P(N_{t+1} = y | N_t = x) = p(y|x; \mu)$

$$p(y|x;\mu) := rac{p(x,y;\mu)}{p(x;\mu)}, \quad x,y \in \mathbb{N}$$

is Poisson AR(1) or INAR(1) ('Poisson O-U'). (trawl representation?)

(iii) (Mehler's formula):

$$p(x, y; \mu) = \sum_{k=0}^{\infty} \frac{\mu_3^k}{k!} D_-^k p(x; \mu) D_-^k p(y; \mu)$$
  
=  $p(x; \mu) p(y; \mu) \sum_{k=0}^{\infty} \frac{\mu_3^k}{\mu^{2k}k!} P_k(x; \mu) P_k(y; \mu).$ 

- For μ = 1 (ii) and (iii) provide complete expansions of joint expectations and probabilities in powers μ<sub>3</sub><sup>k</sup> (μ<sub>3</sub> = Corr(N<sub>1</sub>, N<sub>2</sub>))
- Nearly complete analogy with Hermite expansions in the Gaussian case... (how far?)
- Stationary Markov process  $\{N_t; t = 0, 1, \dots\}$  with marginal Poisson distribution  $P(N_t = x) = p(x; \mu)$  and transition probabilities  $P(N_{t+1} = y | N_t = x) = p(y|x; \mu)$

$$p(y|x;\mu) := rac{p(x,y;\mu)}{p(x;\mu)}, \quad x,y \in \mathbb{N}$$

is Poisson AR(1) or INAR(1) ('Poisson O-U'). (trawl representation?) Proof: bivariate generating function.

(iii) (Mehler's formula):

$$p(x, y; \mu) = \sum_{k=0}^{\infty} \frac{\mu_3^k}{k!} D_-^k p(x; \mu) D_-^k p(y; \mu)$$
  
=  $p(x; \mu) p(y; \mu) \sum_{k=0}^{\infty} \frac{\mu_3^k}{\mu^{2k}k!} P_k(x; \mu) P_k(y; \mu).$ 

- For μ = 1 (ii) and (iii) provide complete expansions of joint expectations and probabilities in powers μ<sub>3</sub><sup>k</sup> (μ<sub>3</sub> = Corr(N<sub>1</sub>, N<sub>2</sub>))
- Nearly complete analogy with Hermite expansions in the Gaussian case... (how far?)
- Stationary Markov process  $\{N_t; t = 0, 1, \dots\}$  with marginal Poisson distribution  $P(N_t = x) = p(x; \mu)$  and transition probabilities  $P(N_{t+1} = y | N_t = x) = p(y|x; \mu)$

$$p(y|x;\mu) := rac{p(x,y;\mu)}{p(x;\mu)}, \quad x,y \in \mathbb{N}$$

is Poisson AR(1) or INAR(1) ('Poisson O-U'). (trawl representation?) Proof: bivariate generating function. Particular case of Markov evolutions of non-interacting particle systems with death and immigration:

**Proof of Lemma 1.** (i) Use generating function  $\mathcal{P}(u; x, \mu) = (1 + u)^{x} e^{-u\mu}$  of Charlier polynomials in (16):

 $E\mathcal{P}(u; N_1, \mu)\mathcal{P}(v; N_2, \mu) = e^{-(u+v)\mu}E[(1+u)^{N_1}(1+v)^{N_2}]$ 

$$\begin{split} & \mathrm{E}\mathcal{P}(u; N_1, \mu)\mathcal{P}(v; N_2, \mu) &= \mathrm{e}^{-(u+v)\mu} \mathrm{E}[(1+u)^{N_1}(1+v)^{N_2}] \\ &= \mathrm{e}^{-(u+v)\mu} \mathrm{E}[(1+u)^{M_1}] \mathrm{E}[(1+v)^{M_2}] \mathrm{E}[((1+u)(1+v))^{M_3}] \end{split}$$

$$\begin{split} \mathrm{E}\mathcal{P}(u;N_{1},\mu)\mathcal{P}(v;N_{2},\mu) &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{N_{1}}(1+v)^{N_{2}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{M_{1}}]\mathrm{E}[(1+v)^{M_{2}}]\mathrm{E}[((1+u)(1+v))^{M_{3}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{e}^{(\mu-\mu_{3})u}\mathrm{e}^{((\mu-\mu_{3})v}\mathrm{e}^{((1+u)(1+v)-1)\mu_{3}} \end{split}$$

$$\begin{split} \mathrm{E}\mathcal{P}(u;N_{1},\mu)\mathcal{P}(v;N_{2},\mu) &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{N_{1}}(1+v)^{N_{2}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{M_{1}}]\mathrm{E}[(1+v)^{M_{2}}]\mathrm{E}[((1+u)(1+v))^{M_{3}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{e}^{(\mu-\mu_{3})u}\mathrm{e}^{(\mu-\mu_{3})v}\mathrm{e}^{((1+u)(1+v)-1)\mu_{3}} \\ &= \mathrm{e}^{uv\mu_{3}} = \sum_{k=0}^{\infty}\frac{(uv\mu_{3})^{k}}{k!}. \end{split}$$

$$\begin{split} \mathrm{E}\mathcal{P}(u;N_{1},\mu)\mathcal{P}(v;N_{2},\mu) &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{N_{1}}(1+v)^{N_{2}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{M_{1}}]\mathrm{E}[(1+v)^{M_{2}}]\mathrm{E}[((1+u)(1+v))^{M_{3}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{e}^{(\mu-\mu_{3})u}\mathrm{e}^{(\mu-\mu_{3})v}\mathrm{e}^{((1+u)(1+v)-1)\mu_{3}} \\ &= \mathrm{e}^{uv\mu_{3}} = \sum_{k=0}^{\infty}\frac{(uv\mu_{3})^{k}}{k!}. \end{split}$$

**Proof of Lemma 1.** (i) Use generating function  $\mathcal{P}(u; x, \mu) = (1 + u)^{x} e^{-u\mu}$  of Charlier polynomials in (16):

$$\begin{split} \mathrm{E}\mathcal{P}(u;N_{1},\mu)\mathcal{P}(v;N_{2},\mu) &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{N_{1}}(1+v)^{N_{2}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{M_{1}}]\mathrm{E}[(1+v)^{M_{2}}]\mathrm{E}[((1+u)(1+v))^{M_{3}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{e}^{(\mu-\mu_{3})u}\mathrm{e}^{(\mu-\mu_{3})v}\mathrm{e}^{((1+u)(1+v)-1)\mu_{3}} \\ &= \mathrm{e}^{uv\mu_{3}} = \sum_{k=0}^{\infty}\frac{(uv\mu_{3})^{k}}{k!}. \end{split}$$

On the other hand,

$$\mathbb{E}\mathcal{P}(u; N_1, \mu)\mathcal{P}(v; N_2, \mu) = \sum_{k,\ell=0}^{\infty} \frac{u^k v^\ell}{k!\ell!} \mathbb{E}[\mathcal{P}_k(N_1; \mu)\mathcal{P}_\ell(N_2; \mu)]$$

(i) follows by equating the coefficients of the power series on both sides.

**Proof of Lemma 1.** (i) Use generating function  $\mathcal{P}(u; x, \mu) = (1 + u)^{x} e^{-u\mu}$  of Charlier polynomials in (16):

$$\begin{split} \mathrm{E}\mathcal{P}(u;N_{1},\mu)\mathcal{P}(v;N_{2},\mu) &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{N_{1}}(1+v)^{N_{2}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{M_{1}}]\mathrm{E}[(1+v)^{M_{2}}]\mathrm{E}[((1+u)(1+v))^{M_{3}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{e}^{(\mu-\mu_{3})u}\mathrm{e}^{(\mu-\mu_{3})v}\mathrm{e}^{((1+u)(1+v)-1)\mu_{3}} \\ &= \mathrm{e}^{uv\mu_{3}} = \sum_{k=0}^{\infty}\frac{(uv\mu_{3})^{k}}{k!}. \end{split}$$

On the other hand,

$$\mathbb{E}\mathcal{P}(u; N_1, \mu)\mathcal{P}(v; N_2, \mu) = \sum_{k,\ell=0}^{\infty} \frac{u^k v^\ell}{k!\ell!} \mathbb{E}[\mathcal{P}_k(N_1; \mu)\mathcal{P}_\ell(N_2; \mu)]$$

(i) follows by equating the coefficients of the power series on both sides.(ii) Immediate from (i) and (17).

**Proof of Lemma 1.** (i) Use generating function  $\mathcal{P}(u; x, \mu) = (1 + u)^{x} e^{-u\mu}$  of Charlier polynomials in (16):

$$\begin{split} \mathrm{E}\mathcal{P}(u; N_{1}, \mu)\mathcal{P}(v; N_{2}, \mu) &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{N_{1}}(1+v)^{N_{2}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{M_{1}}]\mathrm{E}[(1+v)^{M_{2}}]\mathrm{E}[((1+u)(1+v))^{M_{3}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{e}^{(\mu-\mu_{3})u}\mathrm{e}^{((\mu-\mu_{3})v}\mathrm{e}^{((1+u)(1+v)-1)\mu_{3}} \\ &= \mathrm{e}^{uv\mu_{3}} = \sum_{k=0}^{\infty}\frac{(uv\mu_{3})^{k}}{k!}. \end{split}$$

On the other hand,

$$\mathbb{E}\mathcal{P}(u; N_1, \mu)\mathcal{P}(v; N_2, \mu) = \sum_{k,\ell=0}^{\infty} \frac{u^k v^\ell}{k!\ell!} \mathbb{E}[P_k(N_1; \mu)P_\ell(N_2; \mu)]$$

(i) follows by equating the coefficients of the power series on both sides.(ii) Immediate from (i) and (17).

(iii) Apply (ii) to  $G_1(x) := \mathbb{I}(x = n), G_2(x) := \mathbb{I}(x = m)$ , for given  $n, m \in \mathbb{N}$ .

**Proof of Lemma 1.** (i) Use generating function  $\mathcal{P}(u; x, \mu) = (1 + u)^{x} e^{-u\mu}$  of Charlier polynomials in (16):

$$\begin{split} \mathrm{E}\mathcal{P}(u; N_{1}, \mu)\mathcal{P}(v; N_{2}, \mu) &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{N_{1}}(1+v)^{N_{2}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{E}[(1+u)^{M_{1}}]\mathrm{E}[(1+v)^{M_{2}}]\mathrm{E}[((1+u)(1+v))^{M_{3}}] \\ &= \mathrm{e}^{-(u+v)\mu}\mathrm{e}^{(\mu-\mu_{3})u}\mathrm{e}^{((\mu-\mu_{3})v}\mathrm{e}^{((1+u)(1+v)-1)\mu_{3}} \\ &= \mathrm{e}^{uv\mu_{3}} = \sum_{k=0}^{\infty}\frac{(uv\mu_{3})^{k}}{k!}. \end{split}$$

On the other hand,

$$\mathbb{E}\mathcal{P}(u; N_1, \mu)\mathcal{P}(v; N_2, \mu) = \sum_{k,\ell=0}^{\infty} \frac{u^k v^\ell}{k!\ell!} \mathbb{E}[\mathcal{P}_k(N_1; \mu)\mathcal{P}_\ell(N_2; \mu)]$$

(i) follows by equating the coefficients of the power series on both sides.(ii) Immediate from (i) and (17).

(iii) Apply (ii) to  $G_1(x) := \mathbb{I}(x = n), G_2(x) := \mathbb{I}(x = m)$ , for given  $n, m \in \mathbb{N}$ . By (19), (16),  $c_{G_1}(k) = \mathbb{E}[D_+^k \mathbb{I}(N_1 = n)] = D_-^k p(n; \mu) = (-1)^k \mu^{-k} P_k(n; \mu) p(n; \mu), c_{G_2}(k) = \mathbb{E}[D_+^k \mathbb{I}(N_2 = m)] = D_-^k p(m; \mu) = (-1)^k \mu^{-k} P_k(m; \mu) p(n; \mu)$ , yielding (iii).

Let  $G_i$ ,  $N_i$ , i = 1, 2 be as in Lemma 1,  $k_C^*(G_i) = Charlier$  rank of  $G_i$ .

Let  $G_i, N_i, i = 1, 2$  be as in Lemma 1,  $k_C^*(G_i) = Charlier$  rank of  $G_i$ . Then

$$\begin{aligned} \operatorname{Cov}(G_1(N_1), G_2(N_2)) &= \sum_{k=k_c^*(G_1)\vee k_c^*(G_2)}^{\infty} \frac{c_{G_1}(k)c_{G_2}(k)}{k!} \mu_3^k \\ &= \frac{c_{G_1}(k_c^*)c_{G_2}(k^*)}{k^*!} \mu_3^{k_c^*} + R(k_c^*) \end{aligned}$$

where  $k_{C}^{*} := k_{C}^{*}(G_{1}) \vee k_{C}^{*}(G_{2})$  and

$$|R(k_{C}^{*})| \leq \frac{(\mu_{3}/\mu)^{k_{C}^{*}+1}}{1-(\mu_{3}/\mu)} \prod_{i=1}^{2} E^{1/2} G(N_{i})^{2}.$$
 (22)

Let  $G_i, N_i, i = 1, 2$  be as in Lemma 1,  $k_C^*(G_i) = Charlier rank of G_i$ . Then

$$\begin{aligned} \operatorname{Cov}(G_1(N_1), G_2(N_2)) &= \sum_{k=k_c^*(G_1)\vee k_c^*(G_2)}^{\infty} \frac{c_{G_1}(k)c_{G_2}(k)}{k!} \mu_3^k \\ &= \frac{c_{G_1}(k_c^*)c_{G_2}(k^*)}{k^*!} \mu_3^{k_c^*} + R(k_c^*) \end{aligned}$$

where  $k_{C}^{*} := k_{C}^{*}(G_{1}) \vee k_{C}^{*}(G_{2})$  and

$$|R(k_{C}^{*})| \leq \frac{(\mu_{3}/\mu)^{k_{C}^{*}+1}}{1-(\mu_{3}/\mu)} \prod_{i=1}^{2} E^{1/2} G(N_{i})^{2}.$$
(22)

•  $\operatorname{Cov}(G_1(N_1), G_2(N_2))$  decays as  $\mu_3^{k_C^*}$  when  $\mu_3 \to 0$ 

Let  $G_i, N_i, i = 1, 2$  be as in Lemma 1,  $k_C^*(G_i) = Charlier rank of G_i$ . Then

$$\begin{aligned} \operatorname{Cov}(G_1(N_1), G_2(N_2)) &= \sum_{k=k_c^*(G_1)\vee k_c^*(G_2)}^{\infty} \frac{c_{G_1}(k)c_{G_2}(k)}{k!} \mu_3^k \\ &= \frac{c_{G_1}(k_c^*)c_{G_2}(k^*)}{k^*!} \mu_3^{k_c^*} + R(k_c^*) \end{aligned}$$

where  $k_{C}^{*} := k_{C}^{*}(G_{1}) \vee k_{C}^{*}(G_{2})$  and

$$|R(k_{C}^{*})| \leq \frac{(\mu_{3}/\mu)^{k_{C}^{*}+1}}{1-(\mu_{3}/\mu)} \prod_{i=1}^{2} E^{1/2} G(N_{i})^{2}.$$
(22)

- $\operatorname{Cov}(G_1(N_1), G_2(N_2))$  decays as  $\mu_3^{k_C^*}$  when  $\mu_3 \to 0$
- remainder  $R(k_c^*) = O(\mu_3^{k_c^*+1})$
• Aggregated RG model:

$$X_M(t) = \int_{\mathbb{R}^d \times \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}_M(\mathrm{d} u, \mathrm{d} r)$$

 $\mathcal{N}_M$ : Poisson measure with  $\mathbb{E}\mathcal{N}_M(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = M\mathrm{d}\boldsymbol{u}F(\mathrm{d}\boldsymbol{r})$ 

• Aggregated RG model:

$$X_M(t) = \int_{\mathbb{R}^d \times \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}_M(\mathrm{d} u, \mathrm{d} r)$$

 $\mathcal{N}_M$ : Poisson measure with  $\mathbb{E}\mathcal{N}_M(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = M\mathrm{d}\boldsymbol{u}F(\mathrm{d}\boldsymbol{r})$  $X(\boldsymbol{t}) \equiv X_1(\boldsymbol{t}) \text{ (not aggregated)}$ 

• Aggregated RG model:

$$X_M(t) = \int_{\mathbb{R}^d imes \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}_M(\mathrm{d} u, \mathrm{d} r)$$

 $\mathcal{N}_M$ : Poisson measure with  $\mathbb{E}\mathcal{N}_M(\mathrm{d} u, \mathrm{d} r) = M\mathrm{d} u F(\mathrm{d} r)$  $X(t) \equiv X_1(t)$  (not aggregated)

• Subordinated nonlinear model:

$$Y_M(\boldsymbol{t}) := Gig(rac{X_M(\boldsymbol{t}) - \mathrm{E} X_M(\boldsymbol{t})}{M^{1/2}}ig), \quad \boldsymbol{t} \in \mathbb{R}^d,$$

where  $G(x), x \in \mathbb{R}$  is a given nonlinear function;

• Aggregated RG model:

$$X_M(t) = \int_{\mathbb{R}^d imes \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}_M(\mathrm{d} u, \mathrm{d} r)$$

 $\mathcal{N}_M$ : Poisson measure with  $\mathbb{E}\mathcal{N}_M(\mathrm{d} u, \mathrm{d} r) = M\mathrm{d} u F(\mathrm{d} r)$  $X(t) \equiv X_1(t)$  (not aggregated)

• Subordinated nonlinear model:

$$Y_M(t) := Gig(rac{X_M(t) - \mathbb{E} X_M(t)}{M^{1/2}}ig), \quad t \in \mathbb{R}^d,$$

• Aggregated RG model:

$$X_M(t) = \int_{\mathbb{R}^d imes \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}_M(\mathrm{d} u, \mathrm{d} r)$$

 $\mathcal{N}_M$ : Poisson measure with  $\mathbb{E}\mathcal{N}_M(\mathrm{d} u, \mathrm{d} r) = M\mathrm{d} u F(\mathrm{d} r)$  $X(t) \equiv X_1(t)$  (not aggregated)

• Subordinated nonlinear model:

$$Y_M(\boldsymbol{t}) := Gig(rac{X_M(\boldsymbol{t}) - \mathbb{E} X_M(\boldsymbol{t})}{M^{1/2}}ig), \quad \boldsymbol{t} \in \mathbb{R}^d,$$

where  $G(x), x \in \mathbb{R}$  is a given nonlinear function;  $Y(t) = Y_1(t)$ 

• Re-scaled integrals:  $X_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} X_M(t)\phi(t/\lambda) dt$ ,  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} Y_M(t)\phi(t/\lambda) dt$ ,  $X_{\lambda}(\phi) = X_{\lambda,1}(\phi)$ ,  $Y_{\lambda}(\phi) = Y_{\lambda,1}(\phi)$ 

• Aggregated RG model:

$$X_M(t) = \int_{\mathbb{R}^d imes \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}_M(\mathrm{d} u, \mathrm{d} r)$$

 $\mathcal{N}_M$ : Poisson measure with  $\mathbb{E}\mathcal{N}_M(\mathrm{d} u, \mathrm{d} r) = M\mathrm{d} u F(\mathrm{d} r)$  $X(t) \equiv X_1(t)$  (not aggregated)

Subordinated nonlinear model:

$$Y_M(\boldsymbol{t}) := Gig(rac{X_M(\boldsymbol{t}) - \mathbb{E} X_M(\boldsymbol{t})}{M^{1/2}}ig), \quad \boldsymbol{t} \in \mathbb{R}^d,$$

- Re-scaled integrals:  $X_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} X_M(t)\phi(t/\lambda)dt$ ,  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} Y_M(t)\phi(t/\lambda)dt$ ,  $X_{\lambda}(\phi) = X_{\lambda,1}(\phi)$ ,  $Y_{\lambda}(\phi) = Y_{\lambda,1}(\phi)$
- Problem: limit distribution of Y<sub>λ,M</sub>(φ) and Y<sub>λ</sub>(φ) as λ → ∞ and M = λ<sup>γ</sup> → ∞, for each φ ∈ Φ = L<sup>1</sup>(ℝ<sup>d</sup>) ∩ L<sup>∞</sup>(ℝ<sup>d</sup>)

Aggregated RG model:

$$X_M(t) = \int_{\mathbb{R}^d imes \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}_M(\mathrm{d} u, \mathrm{d} r)$$

 $\mathcal{N}_M$ : Poisson measure with  $\mathbb{E}\mathcal{N}_M(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = M\mathrm{d}\boldsymbol{u}\boldsymbol{F}(\mathrm{d}\boldsymbol{r})$  $X(\boldsymbol{t}) \equiv X_1(\boldsymbol{t})$  (not aggregated)

Subordinated nonlinear model:

$$Y_M(t) := G\left(rac{X_M(t) - \mathbb{E} X_M(t)}{M^{1/2}}
ight), \quad t \in \mathbb{R}^d,$$

- Re-scaled integrals:  $X_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} X_M(t)\phi(t/\lambda) dt$ ,  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} Y_M(t)\phi(t/\lambda) dt$ ,  $X_{\lambda}(\phi) = X_{\lambda,1}(\phi)$ ,  $Y_{\lambda}(\phi) = Y_{\lambda,1}(\phi)$
- Problem: limit distribution of Y<sub>λ,M</sub>(φ) and Y<sub>λ</sub>(φ) as λ → ∞ and M = λ<sup>γ</sup> → ∞, for each φ ∈ Φ = L<sup>1</sup>(ℝ<sup>d</sup>) ∩ L<sup>∞</sup>(ℝ<sup>d</sup>)
- Under Assumption LRD: (≈ P(R > r) = F(r,∞) ~ c<sub>f</sub>r<sup>-α</sup>, r → ∞, α ∈ (1,2)) the limit of *linear* X<sub>λ,M</sub>(φ) and X<sub>λ</sub>(φ) described in Thm 1 [KLNS].

Aggregated RG model:

$$X_M(t) = \int_{\mathbb{R}^d imes \mathbb{R}_+} \mathbb{I}(t - u \in r^{1/d} \Xi^0) \mathcal{N}_M(\mathrm{d} u, \mathrm{d} r)$$

 $\mathcal{N}_M$ : Poisson measure with  $\mathbb{E}\mathcal{N}_M(\mathrm{d}\boldsymbol{u},\mathrm{d}\boldsymbol{r}) = M\mathrm{d}\boldsymbol{u}\boldsymbol{F}(\mathrm{d}\boldsymbol{r})$  $X(\boldsymbol{t}) \equiv X_1(\boldsymbol{t})$  (not aggregated)

Subordinated nonlinear model:

$$Y_M(t) := G\left(rac{X_M(t) - \mathrm{E} X_M(t)}{M^{1/2}}
ight), \quad t \in \mathbb{R}^d,$$

- Re-scaled integrals:  $X_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} X_M(t)\phi(t/\lambda) dt$ ,  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} Y_M(t)\phi(t/\lambda) dt$ ,  $X_{\lambda}(\phi) = X_{\lambda,1}(\phi)$ ,  $Y_{\lambda}(\phi) = Y_{\lambda,1}(\phi)$
- Problem: limit distribution of Y<sub>λ,M</sub>(φ) and Y<sub>λ</sub>(φ) as λ → ∞ and M = λ<sup>γ</sup> → ∞, for each φ ∈ Φ = L<sup>1</sup>(ℝ<sup>d</sup>) ∩ L<sup>∞</sup>(ℝ<sup>d</sup>)
- Under Assumption LRD: (≈ P(R > r) = F(r,∞) ~ c<sub>f</sub>r<sup>-α</sup>, r → ∞, α ∈ (1,2)) the limit of *linear* X<sub>λ,M</sub>(φ) and X<sub>λ</sub>(φ) described in Thm 1 [KLNS]. Trichotomy of the limit at γ = α − 1

1. If *Hermite rank* of G is 1 then the limits of  $Y_{\lambda,M}(\phi)$  and  $X_{\lambda,M}(\phi)$ ,  $M = \lambda^{\gamma}$  are the same (up to the first Hermite coefficient of G), for any  $\gamma > 0$ 

- 1. If *Hermite rank* of G is 1 then the limits of  $Y_{\lambda,M}(\phi)$  and  $X_{\lambda,M}(\phi)$ ,  $M = \lambda^{\gamma}$  are the same (up to the first Hermite coefficient of G), for any  $\gamma > 0$
- 2. If *Charlier rank* of G is 1 then the limits of  $Y_{\lambda}(\phi)$  and  $X_{\lambda}(\phi)$  are the same (up to the first Charlier coefficient of G)

- 1. If *Hermite rank* of G is 1 then the limits of  $Y_{\lambda,M}(\phi)$  and  $X_{\lambda,M}(\phi)$ ,  $M = \lambda^{\gamma}$  are the same (up to the first Hermite coefficient of G), for any  $\gamma > 0$
- 2. If *Charlier rank* of G is 1 then the limits of  $Y_{\lambda}(\phi)$  and  $X_{\lambda}(\phi)$  are the same (up to the first Charlier coefficient of G)
- For fixed  $\mathbf{t} \in \mathbb{R}^d$ ,  $Y_M(\mathbf{t}) \xrightarrow{d} Z_\mu \sim N(0,\mu)$  as  $M \to \infty$  where  $\mu = \mathrm{E}X(\mathbf{t}) = \mathrm{Var}(X(\mathbf{t})) = Leb_d(\Xi^0)\mathrm{E}R$

- 1. If *Hermite rank* of G is 1 then the limits of  $Y_{\lambda,M}(\phi)$  and  $X_{\lambda,M}(\phi)$ ,  $M = \lambda^{\gamma}$  are the same (up to the first Hermite coefficient of G), for any  $\gamma > 0$
- 2. If *Charlier rank* of G is 1 then the limits of  $Y_{\lambda}(\phi)$  and  $X_{\lambda}(\phi)$  are the same (up to the first Charlier coefficient of G)

• For fixed 
$$\mathbf{t} \in \mathbb{R}^d$$
,  $Y_M(\mathbf{t}) \xrightarrow{d} Z_\mu \sim N(0,\mu)$  as  $M \to \infty$  where  $\mu = \mathrm{E}X(\mathbf{t}) = \mathrm{Var}(X(\mathbf{t})) = Leb_d(\Xi^0)\mathrm{E}R$ 

• Expand  $G(x) = \sum_{k=0}^{\infty} \frac{h_{G,\mu}(k)}{k!} H_k(x;\mu)$  in Hermite polynomials  $H_k(x;\mu)$  with generating function  $\sum_{k=0}^{\infty} (u^k/k!) H_k(x;\mu) = e^{ux - \mu u^2/2}$  and coefficients

$$h_{G,\mu}(k) = \mu^{-k} \mathbb{E}G(Z_{\mu})H_k(Z_{\mu};\mu), \quad k \in \mathbb{N}.$$

- 1. If *Hermite rank* of G is 1 then the limits of  $Y_{\lambda,M}(\phi)$  and  $X_{\lambda,M}(\phi)$ ,  $M = \lambda^{\gamma}$  are the same (up to the first Hermite coefficient of G), for any  $\gamma > 0$
- 2. If *Charlier rank* of G is 1 then the limits of  $Y_{\lambda}(\phi)$  and  $X_{\lambda}(\phi)$  are the same (up to the first Charlier coefficient of G)

• For fixed 
$$\mathbf{t} \in \mathbb{R}^d$$
,  $Y_M(\mathbf{t}) \xrightarrow{d} Z_\mu \sim N(0,\mu)$  as  $M \to \infty$  where  $\mu = \mathrm{E}X(\mathbf{t}) = \mathrm{Var}(X(\mathbf{t})) = Leb_d(\Xi^0)\mathrm{E}R$ 

• Expand  $G(x) = \sum_{k=0}^{\infty} \frac{h_{G,\mu}(k)}{k!} H_k(x;\mu)$  in Hermite polynomials  $H_k(x;\mu)$  with generating function  $\sum_{k=0}^{\infty} (u^k/k!) H_k(x;\mu) = e^{ux - \mu u^2/2}$  and coefficients

$$h_{G,\mu}(k) = \mu^{-k} \mathbb{E}G(Z_{\mu})H_k(Z_{\mu};\mu), \quad k \in \mathbb{N}.$$

•  $h_{G,\mu}(1) = \mu^{-1} \mathbb{E} G(Z_{\mu}) Z_{\mu}$ 

# Theorem (2)

1. Let  $X_M(t)$  satisfy the conditions of Theorem 1,

1. Let  $X_M(\mathbf{t})$  satisfy the conditions of Theorem 1,  $G = G(x), x \in \mathbb{R}$  is an dx-a.e. continuous function such that  $\mathrm{E}G(Y_{\lambda,M}(\mathbf{0}))^2 < \infty \ (\forall M > 0)$  and

 $\lim_{M\to\infty} \operatorname{E} G(Y_{\lambda,M}(\mathbf{0}))^2 = \operatorname{E} G(Z_{\mu})^2 < \infty.$ 

1. Let  $X_M(t)$  satisfy the conditions of Theorem 1,  $G = G(x), x \in \mathbb{R}$  is an dx-a.e. continuous function such that  $\mathrm{E}G(Y_{\lambda,M}(\mathbf{0}))^2 < \infty \ (\forall M > 0)$  and

$$\lim_{M\to\infty} \operatorname{E} G(Y_{\lambda,M}(\mathbf{0}))^2 = \operatorname{E} G(Z_{\mu})^2 < \infty.$$

Let  $M = \lambda^{\gamma}$  for some  $\gamma > 0$ .

1. Let  $X_M(t)$  satisfy the conditions of Theorem 1,  $G = G(x), x \in \mathbb{R}$  is an dx-a.e. continuous function such that  $\mathrm{E}G(Y_{\lambda,M}(\mathbf{0}))^2 < \infty (\forall M > 0)$  and

$$\lim_{M\to\infty} \mathrm{E}G(Y_{\lambda,M}(\mathbf{0}))^2 = \mathrm{E}G(Z_{\mu})^2 < \infty.$$

Let  $M = \lambda^{\gamma}$  for some  $\gamma > 0$ . Then for any  $\phi \in \Phi$  as  $\lambda \to \infty$ 

$$\lambda^{(\gamma/2)-H(\gamma)}(Y_{\lambda,M}(\phi)-\mathrm{E}Y_{\lambda,M}(\phi)) \stackrel{\mathrm{d}}{\longrightarrow} h_{\mathcal{G},\mu}(1) egin{cases} B_lpha(\phi), & \gamma > d(lpha-1),\ L_lpha(\phi), & \gamma < d(lpha-1),\ J_lpha(\phi), & \gamma = d(lpha-1), \end{cases}$$

1. Let  $X_M(t)$  satisfy the conditions of Theorem 1,  $G = G(x), x \in \mathbb{R}$  is an dx-a.e. continuous function such that  $\mathrm{E}G(Y_{\lambda,M}(\mathbf{0}))^2 < \infty (\forall M > 0)$  and

$$\lim_{M\to\infty} \mathrm{E}G(Y_{\lambda,M}(\mathbf{0}))^2 = \mathrm{E}G(Z_{\mu})^2 < \infty.$$

Let  $M = \lambda^{\gamma}$  for some  $\gamma > 0$ . Then for any  $\phi \in \Phi$  as  $\lambda \to \infty$ 

$$\lambda^{(\gamma/2)-H(\gamma)}(Y_{\lambda,M}(\phi)-\mathrm{E}Y_{\lambda,M}(\phi)) \stackrel{\mathrm{d}}{\longrightarrow} h_{\mathcal{G},\mu}(1) egin{cases} B_lpha(\phi), & \gamma > d(lpha-1),\ L_lpha(\phi), & \gamma < d(lpha-1),\ J_lpha(\phi), & \gamma = d(lpha-1), \end{cases}$$

where  $H(\gamma), B_{\alpha}(\phi), L_{\alpha}(\phi), J_{\alpha}(\phi)$  are the same as in Thm 1.

1. Let  $X_M(t)$  satisfy the conditions of Theorem 1,  $G = G(x), x \in \mathbb{R}$  is an dx-a.e. continuous function such that  $\mathrm{E}G(Y_{\lambda,M}(\mathbf{0}))^2 < \infty (\forall M > 0)$  and

$$\lim_{M\to\infty} \mathrm{E}G(Y_{\lambda,M}(\mathbf{0}))^2 = \mathrm{E}G(Z_{\mu})^2 < \infty.$$

Let  $M = \lambda^{\gamma}$  for some  $\gamma > 0$ . Then for any  $\phi \in \Phi$  as  $\lambda \to \infty$ 

$$\lambda^{(\gamma/2)-H(\gamma)}(Y_{\lambda,M}(\phi)-\operatorname{E} Y_{\lambda,M}(\phi)) \stackrel{\mathrm{d}}{\longrightarrow} h_{G,\mu}(1) egin{cases} B_lpha(\phi), & \gamma > d(lpha-1),\ L_lpha(\phi), & \gamma < d(lpha-1),\ J_lpha(\phi), & \gamma = d(lpha-1), \end{cases}$$

where  $H(\gamma)$ ,  $B_{\alpha}(\phi)$ ,  $L_{\alpha}(\phi)$ ,  $J_{\alpha}(\phi)$  are the same as in Thm 1.

2. Let Y(t) = G(X(t)), where X(t) is as in Thm 1 and  $EY(t)^2 < \infty$ . Then for any  $\phi \in \Phi$  as  $\lambda \to \infty$ 

$$\lambda^{-d/lpha}(Y_\lambda(\phi) - \mathrm{E} Y_\lambda(\phi)) \stackrel{\mathrm{d}}{\longrightarrow} c_{\mathcal{G},\mu}(1) L_lpha(\phi),$$

where  $c_{G,\mu}(1) = EG(X(\mathbf{0}))(X(\mathbf{0}) - EX(\mathbf{0}))$  is the first Charlier coefficient of G and  $L_{\alpha}(\phi)$  is the same  $\alpha$ -stable RF as in part 1.

The Boolean model  $\hat{X}(t) = X(t) \wedge 1$  corresponds to Y(t) = G(X(t)) with  $G(x) = x \wedge 1, x \in \mathbb{N}$ .

The Boolean model  $\hat{X}(t) = X(t) \wedge 1$  corresponds to Y(t) = G(X(t)) with  $G(x) = x \wedge 1, x \in \mathbb{N}$ . For  $\phi(x) = \mathbb{I}(x \in A), Y_{\lambda}(\phi) = Leb_d(\mathcal{X} \cap \lambda A) =: \hat{X}_{\lambda}(A)$  (= volume of  $\{X(t) = 1\} \cap \lambda A$ )

The Boolean model  $\hat{X}(t) = X(t) \wedge 1$  corresponds to Y(t) = G(X(t)) with  $G(x) = x \wedge 1, x \in \mathbb{N}$ . For  $\phi(x) = \mathbb{I}(x \in A), \ Y_{\lambda}(\phi) = Leb_d(\mathcal{X} \cap \lambda A) =: \hat{X}_{\lambda}(A) \ (= \text{ volume of } \{X(t) = 1\} \cap \lambda A)$ Charlier coefficients:  $c_{G,\mu}(0) = 1 - e^{-\mu}, c_{G,\mu}(k) = (-1)^{k+1}e^{-\mu} \ (k \ge 1)$ 

The Boolean model  $\hat{X}(t) = X(t) \wedge 1$  corresponds to Y(t) = G(X(t)) with  $G(x) = x \wedge 1, x \in \mathbb{N}$ . For  $\phi(x) = \mathbb{I}(x \in A), \ Y_{\lambda}(\phi) = Leb_d(\mathcal{X} \cap \lambda A) =: \hat{X}_{\lambda}(A) \ (= \text{ volume of } \{X(t) = 1\} \cap \lambda A)$ Charlier coefficients:  $c_{G,\mu}(0) = 1 - e^{-\mu}, c_{G,\mu}(k) = (-1)^{k+1}e^{-\mu} \ (k \ge 1)$ 

## Corollary (1)

Let  $A \subset \mathbb{R}^d$  be a bounded Borel set and X(t) RG model as in Thm 1. Then

$$\lambda^{-d/lpha}(\hat{X}_{\lambda}(\mathcal{A}) - \mathrm{E}\hat{X}_{\lambda}(\mathcal{A})) \stackrel{\mathrm{d}}{\longrightarrow} \mathrm{e}^{-\mu}L_{lpha}(\mathcal{A}), \quad \lambda o \infty$$

where  $L_{\alpha}(A)$  is asymmetric  $\alpha$ -stable r.v. with  $\operatorname{Ee}^{\mathrm{i}\theta L_{\alpha}(A)} = \exp\{-\sigma_{\alpha}|\theta|^{\alpha} \operatorname{Leb}_{d}(A)(1-\mathrm{i}\operatorname{sgn}(\theta)\tan(\pi\alpha/2))\}, \theta \in \mathbb{R}.$ 

$$\mathcal{E}_{M}(\boldsymbol{t}) := \mathrm{e}^{\mathfrak{s}(X_{M}(\boldsymbol{t}) - \mathrm{E}X_{M}(\boldsymbol{t}))/M^{1/2}}, \quad \mathcal{E}_{\lambda,M}(\phi) := \int_{\mathbb{R}^{d}} \phi(\boldsymbol{t}/\lambda) \mathcal{E}_{M}(\boldsymbol{t}) \mathrm{d}\boldsymbol{t}.$$

$$\mathcal{E}_{M}(\boldsymbol{t}) := \mathrm{e}^{\boldsymbol{a}(X_{M}(\boldsymbol{t}) - \mathbb{E}X_{M}(\boldsymbol{t}))/M^{1/2}}, \quad \mathcal{E}_{\lambda,M}(\phi) := \int_{\mathbb{R}^{d}} \phi(\boldsymbol{t}/\lambda) \mathcal{E}_{M}(\boldsymbol{t}) \mathrm{d}\boldsymbol{t}.$$

Particular case of (23) corresponding to  $G(x) = e^{ax}$ . Note  $D^k_+ G(x) = (e^a - 1)^k e^{ax}$ and  $c_{G,\mu}(k) = (e^a - 1)^k e^{(e^a - 1)\mu}$ ,  $k \in \mathbb{N}$ . We also have

$$M^{1/2}c_{G(\cdot/M^{1/2}),\mu M}(1) = \exp\{(e^{a/M^{1/2}} - 1 - (a/M^{1/2}))\mu M\}M^{1/2}(e^{a/M^{1/2}} - 1)$$
  
$$\to ae^{a^2\mu/2} = E[e^{aZ_{\mu}}Z_{\mu}] = h_{G,\mu}(1)$$

$$\mathcal{E}_{M}(\boldsymbol{t}):=\mathrm{e}^{\mathfrak{s}(X_{M}(\boldsymbol{t})-\mathrm{E}X_{M}(\boldsymbol{t}))/M^{1/2}},\quad \mathcal{E}_{\lambda,M}(\phi):=\int_{\mathbb{R}^{d}}\phi(\boldsymbol{t}/\lambda)\mathcal{E}_{M}(\boldsymbol{t})\mathrm{d}\boldsymbol{t}.$$

Particular case of (23) corresponding to  $G(x) = e^{ax}$ . Note  $D^k_+ G(x) = (e^a - 1)^k e^{ax}$ and  $c_{G,\mu}(k) = (e^a - 1)^k e^{(e^a - 1)\mu}$ ,  $k \in \mathbb{N}$ . We also have

$$\begin{split} \mathcal{M}^{1/2} c_{G(\cdot/M^{1/2}),\mu M}(1) &= \exp\{(\mathrm{e}^{a/M^{1/2}} - 1 - (a/M^{1/2}))\mu M\} \mathcal{M}^{1/2}(\mathrm{e}^{a/M^{1/2}} - 1) \\ &\to a \mathrm{e}^{a^2 \mu/2} = \mathrm{E}[\mathrm{e}^{a Z_{\mu}} Z_{\mu}] = h_{G,\mu}(1) \end{split}$$

### Corollary (2)

Let  $X_M(t)$  be as in Thm 1 and  $G(x) = e^{ax}$ . Then

$$\mathcal{E}_{\mathcal{M}}(\boldsymbol{t}) := \mathrm{e}^{\mathfrak{s}(X_{\mathcal{M}}(\boldsymbol{t}) - \mathbb{E} X_{\mathcal{M}}(\boldsymbol{t}))/M^{1/2}}, \quad \mathcal{E}_{\lambda,\mathcal{M}}(\phi) := \int_{\mathbb{R}^d} \phi(\boldsymbol{t}/\lambda) \mathcal{E}_{\mathcal{M}}(\boldsymbol{t}) \mathrm{d}\boldsymbol{t}.$$

Particular case of (23) corresponding to  $G(x) = e^{ax}$ . Note  $D^k_+ G(x) = (e^a - 1)^k e^{ax}$ and  $c_{G,\mu}(k) = (e^a - 1)^k e^{(e^a - 1)\mu}$ ,  $k \in \mathbb{N}$ . We also have

$$\begin{split} \mathcal{M}^{1/2} c_{G(\cdot/M^{1/2}),\mu M}(1) &= \exp\{(\mathrm{e}^{a/M^{1/2}} - 1 - (a/M^{1/2}))\mu M\} \mathcal{M}^{1/2}(\mathrm{e}^{a/M^{1/2}} - 1) \\ &\to a \mathrm{e}^{a^2 \mu/2} = \mathrm{E}[\mathrm{e}^{a Z_{\mu}} Z_{\mu}] = h_{G,\mu}(1) \end{split}$$

## Corollary (2)

Let  $X_M(t)$  be as in Thm 1 and  $G(x) = e^{ax}$ . Then

$$\lambda^{(\gamma/2)-H(\gamma)}(\mathcal{E}_{\lambda,M} - \mathrm{E}\mathcal{E}_{\lambda,M}(\phi)) \xrightarrow{\mathrm{d}} a \mathrm{e}^{a^2\mu/2} \begin{cases} B_{\alpha}(\phi), & \gamma > d(\alpha-1), \\ L_{\alpha}(\phi), & \gamma < d(\alpha-1), \\ J_{\alpha}(\phi), & \gamma = d(\alpha-1), \end{cases}$$

where  $H(\gamma)$ ,  $B_{\alpha}(\phi)$ ,  $L_{\alpha}(\phi)$ ,  $J_{\alpha}(\phi)$  are the same as in Thm 1.

Burgers' equation with (random) potential initial data:

$$egin{aligned} \partialec v(t,oldsymbol{x})/\partial t+(ec v(t,oldsymbol{x}),
abla)ec v(t,oldsymbol{x})&=&rac{1}{2}\kappa\Deltaec v(t,oldsymbol{x}),\quad t>0,oldsymbol{x}\in\mathbb{R}^d\ ec v(0,oldsymbol{x})&=&-
abla\xi(oldsymbol{x}), \end{aligned}$$

Burgers' equation with (random) potential initial data:

$$egin{aligned} \partialec v(t,oldsymbol{x})/\partial t+(ec v(t,oldsymbol{x}),
abla)ec v(t,oldsymbol{x})&=&rac{1}{2}\kappa\Deltaec v(t,oldsymbol{x}),\quad t>0,oldsymbol{x}\in\mathbb{R}^d\ ec v(0,oldsymbol{x})&=&-
abla\xi(oldsymbol{x}), \end{aligned}$$

•  $\vec{v}(t, \mathbf{x}) = (v_1(t, \mathbf{x}), \cdots, v_d(t, \mathbf{x}))$ :  $\mathbb{R}^d$ -valued function (velocity field),  $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$ 

Burgers' equation with (random) potential initial data:

$$egin{aligned} \partialec v(t,oldsymbol{x})/\partial t+(ec v(t,oldsymbol{x}),
abla)ec v(t,oldsymbol{x})&=&rac{1}{2}\kappa\Deltaec v(t,oldsymbol{x}),\quad t>0,oldsymbol{x}\in\mathbb{R}^d\ ec v(0,oldsymbol{x})&=&-
abla\xi(oldsymbol{x}), \end{aligned}$$

•  $\vec{v}(t, \mathbf{x}) = (v_1(t, \mathbf{x}), \cdots, v_d(t, \mathbf{x}))$ :  $\mathbb{R}^d$ -valued function (velocity field),  $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$ 

Burgers' equation with (random) potential initial data:

$$egin{aligned} \partialec v(t,oldsymbol{x})/\partial t+(ec v(t,oldsymbol{x}),
abla)ec v(t,oldsymbol{x})&=&rac{1}{2}\kappa\Deltaec v(t,oldsymbol{x}),\quad t>0,oldsymbol{x}\in\mathbb{R}^d\ ec v(0,oldsymbol{x})&=&-
abla\xi(oldsymbol{x}), \end{aligned}$$

- $\vec{v}(t, \mathbf{x}) = (v_1(t, \mathbf{x}), \cdots, v_d(t, \mathbf{x}))$ :  $\mathbb{R}^d$ -valued function (velocity field),  $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$
- $\kappa > 0$ : viscosity parameter,  $\Delta = Laplacian$
Burgers' equation with (random) potential initial data:

$$egin{aligned} \partialec v(t,oldsymbol{x})/\partial t+(ec v(t,oldsymbol{x}),
abla)ec v(t,oldsymbol{x})&=&rac{1}{2}\kappa\Deltaec v(t,oldsymbol{x}),\quad t>0,oldsymbol{x}\in\mathbb{R}^d\ ec v(0,oldsymbol{x})&=&-
abla\xi(oldsymbol{x}), \end{aligned}$$

• 
$$\vec{v}(t, \mathbf{x}) = (v_1(t, \mathbf{x}), \dots, v_d(t, \mathbf{x}))$$
:  $\mathbb{R}^d$ -valued function (velocity field),  
 $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$ 

• 
$$\kappa > 0$$
: viscosity parameter,  $\Delta = Laplacian$ 

•  $\xi = \{\xi(\mathbf{x}); \mathbf{x} \in \mathbb{R}^d\}$ : initial scalar (potential) random field (RF);  $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$ 

$$egin{aligned} \partialec v(t,oldsymbol{x})/\partial t+(ec v(t,oldsymbol{x}),
abla)ec v(t,oldsymbol{x})&=&rac{1}{2}\kappa\Deltaec v(t,oldsymbol{x}),\quad t>0,oldsymbol{x}\in\mathbb{R}^d\ ec v(0,oldsymbol{x})&=&-
abla\xi(oldsymbol{x}), \end{aligned}$$

• 
$$\vec{v}(t, \mathbf{x}) = (v_1(t, \mathbf{x}), \dots, v_d(t, \mathbf{x}))$$
:  $\mathbb{R}^d$ -valued function (velocity field),  
 $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$ 

- $\xi = \{\xi(\mathbf{x}); \mathbf{x} \in \mathbb{R}^d\}$ : initial scalar (potential) random field (RF);  $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$
- one of the important equations of mathematical physics [acoustic, astrophysics, cosmology, turbulence]

$$egin{aligned} \partialec v(t,oldsymbol{x})/\partial t+(ec v(t,oldsymbol{x}),
abla)ec v(t,oldsymbol{x})&=&rac{1}{2}\kappa\Deltaec v(t,oldsymbol{x}),\quad t>0,oldsymbol{x}\in\mathbb{R}^d\ ec v(0,oldsymbol{x})&=&-
abla\xi(oldsymbol{x}), \end{aligned}$$

• 
$$\vec{v}(t, \mathbf{x}) = (v_1(t, \mathbf{x}), \dots, v_d(t, \mathbf{x}))$$
:  $\mathbb{R}^d$ -valued function (velocity field),  
 $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$ 

- $\xi = \{\xi(\mathbf{x}); \mathbf{x} \in \mathbb{R}^d\}$ : initial scalar (potential) random field (RF);  $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$
- one of the important equations of mathematical physics [acoustic, astrophysics, cosmology, turbulence]
- nonlinear but explicitly solvable

$$egin{aligned} \partialec v(t,oldsymbol{x})/\partial t+(ec v(t,oldsymbol{x}),
abla)ec v(t,oldsymbol{x})&=&rac{1}{2}\kappa\Deltaec v(t,oldsymbol{x}),\quad t>0,oldsymbol{x}\in\mathbb{R}^d\ ec v(0,oldsymbol{x})&=&-
abla\xi(oldsymbol{x}), \end{aligned}$$

• 
$$\vec{v}(t, \mathbf{x}) = (v_1(t, \mathbf{x}), \dots, v_d(t, \mathbf{x}))$$
:  $\mathbb{R}^d$ -valued function (velocity field),  
 $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$ 

- $\xi = \{\xi(\mathbf{x}); \mathbf{x} \in \mathbb{R}^d\}$ : initial scalar (potential) random field (RF);  $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$
- one of the important equations of mathematical physics [acoustic, astrophysics, cosmology, turbulence]
- nonlinear but explicitly solvable
- solution  $\vec{v}(t, \mathbf{x})$  with random initial data is a (vector-valued) RF

$$egin{aligned} \partialec v(t,oldsymbol{x})/\partial t+(ec v(t,oldsymbol{x}),
abla)ec v(t,oldsymbol{x})&=&rac{1}{2}\kappa\Deltaec v(t,oldsymbol{x}),\quad t>0,oldsymbol{x}\in\mathbb{R}^d\ ec v(0,oldsymbol{x})&=&-
abla\xi(oldsymbol{x}), \end{aligned}$$

• 
$$\vec{v}(t, \mathbf{x}) = (v_1(t, \mathbf{x}), \dots, v_d(t, \mathbf{x}))$$
:  $\mathbb{R}^d$ -valued function (velocity field),  
 $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$ 

- $\kappa > 0$ : viscosity parameter,  $\Delta =$  Laplacian
- $\xi = \{\xi(\mathbf{x}); \mathbf{x} \in \mathbb{R}^d\}$ : initial scalar (potential) random field (RF);  $(\vec{v}(t, \mathbf{x}), \nabla) := \sum_{i=1}^d v_i(t, \mathbf{x}) \partial / \partial x_i$
- one of the important equations of mathematical physics [acoustic, astrophysics, cosmology, turbulence]
- nonlinear but explicitly solvable
- solution  $\vec{v}(t, \mathbf{x})$  with random initial data is a (vector-valued) RF
- behavior of  $\vec{v}(t, \mathbf{x})$  presents considerable physical and mathematical interest and has been extensively studied
- M. Rosenblatt, Ya. Sinai, S. Molchanov, W. Woyczynski, N. Leonenko, ...

• Hopf-Cole substitution:

$$\vec{v}(t, \mathbf{x}) = -\kappa \nabla \log u(t, \mathbf{x}) = -\frac{\kappa \nabla u(t, \mathbf{x})}{u(t, \mathbf{x})}$$

with scalar-valued  $u(t, \mathbf{x})$  satisfying heat equation  $\partial u(t, \mathbf{x})/\partial t = \frac{1}{2}\kappa\Delta u(t, \mathbf{x})$ with the exponential initial condition  $u(0+, \mathbf{x}) = \exp{\{\xi(\mathbf{x})/\kappa\}}, \mathbf{x} \in \mathbb{R}^d$ .

• Hopf-Cole substitution:

$$\vec{v}(t, \mathbf{x}) = -\kappa \nabla \log u(t, \mathbf{x}) = -\frac{\kappa \nabla u(t, \mathbf{x})}{u(t, \mathbf{x})}$$

with scalar-valued  $u(t, \mathbf{x})$  satisfying heat equation  $\partial u(t, \mathbf{x})/\partial t = \frac{1}{2}\kappa\Delta u(t, \mathbf{x})$ with the exponential initial condition  $u(0+, \mathbf{x}) = \exp\{\xi(\mathbf{x})/\kappa\}, \mathbf{x} \in \mathbb{R}^d$ . Explicit representation through heat kernel  $g(t, \mathbf{x}, \mathbf{y}) := (2\pi\kappa t)^{-d/2} \exp\{-\|\mathbf{x} - \mathbf{y}\|^2/2\kappa t\}$ :

Hopf-Cole substitution:

$$\vec{v}(t, \mathbf{x}) = -\kappa \nabla \log u(t, \mathbf{x}) = -\frac{\kappa \nabla u(t, \mathbf{x})}{u(t, \mathbf{x})}$$

with scalar-valued  $u(t, \mathbf{x})$  satisfying heat equation  $\partial u(t, \mathbf{x})/\partial t = \frac{1}{2}\kappa\Delta u(t, \mathbf{x})$ with the exponential initial condition  $u(0+, \mathbf{x}) = \exp\{\xi(\mathbf{x})/\kappa\}, \mathbf{x} \in \mathbb{R}^d$ . Explicit representation through heat kernel  $g(t, \mathbf{x}, \mathbf{y}) := (2\pi\kappa t)^{-d/2} \exp\{-\|\mathbf{x} - \mathbf{y}\|^2/2\kappa t\}$ :  $\vec{v}(t, \mathbf{x}) = -\frac{\kappa \int_{\mathbb{R}^d} \nabla g(t, \mathbf{x}, \mathbf{y}) e^{\xi(\mathbf{y})/\kappa} d\mathbf{y}}{\int_{\mathbf{x} d} g(t, \mathbf{x}, \mathbf{y}) e^{\xi(\mathbf{y})/\kappa} d\mathbf{y}},$ 

Hopf-Cole substitution:

$$\vec{v}(t, \mathbf{x}) = -\kappa \nabla \log u(t, \mathbf{x}) = -\frac{\kappa \nabla u(t, \mathbf{x})}{u(t, \mathbf{x})}$$

with scalar-valued  $u(t, \mathbf{x})$  satisfying heat equation  $\partial u(t, \mathbf{x})/\partial t = \frac{1}{2}\kappa\Delta u(t, \mathbf{x})$ with the exponential initial condition  $u(0+, \mathbf{x}) = \exp\{\xi(\mathbf{x})/\kappa\}, \mathbf{x} \in \mathbb{R}^d$ . Explicit representation through heat kernel  $g(t, \mathbf{x}, \mathbf{y}) := (2\pi\kappa t)^{-d/2} \exp\{-\|\mathbf{x} - \mathbf{y}\|^2/2\kappa t\}$ :  $\vec{v}(t, \mathbf{x}) = -\frac{\kappa \int_{\mathbb{R}^d} \nabla g(t, \mathbf{x}, \mathbf{y}) e^{\xi(\mathbf{y})/\kappa} d\mathbf{y}}{\int_{-\kappa} g(t, \mathbf{x}, \mathbf{y}) e^{\xi(\mathbf{y})/\kappa} d\mathbf{y}},$ 

• Parabolic scaling leads to the RF  $\vec{v}_{\lambda}(t, \mathbf{x}) := \vec{v}(\lambda^2 t, \lambda \mathbf{x})$  written as

$$\vec{v}_{\lambda}(t, \mathbf{x}) = -\frac{\kappa \int_{\mathbb{R}^d} \nabla g(t, \mathbf{x}, \mathbf{y}/\lambda) \mathrm{e}^{\xi(\mathbf{y})/\kappa} \mathrm{d}\mathbf{y}}{\lambda \int_{\mathbb{R}^d} g(t, \mathbf{x}, \mathbf{y}/\lambda) \mathrm{e}^{\xi(\mathbf{y})/\kappa} \mathrm{d}\mathbf{y}}.$$
(23)

Hopf-Cole substitution:

$$\vec{v}(t, \mathbf{x}) = -\kappa \nabla \log u(t, \mathbf{x}) = -\frac{\kappa \nabla u(t, \mathbf{x})}{u(t, \mathbf{x})}$$

with scalar-valued  $u(t, \mathbf{x})$  satisfying heat equation  $\partial u(t, \mathbf{x})/\partial t = \frac{1}{2}\kappa\Delta u(t, \mathbf{x})$ with the exponential initial condition  $u(0+, \mathbf{x}) = \exp\{\xi(\mathbf{x})/\kappa\}, \mathbf{x} \in \mathbb{R}^d$ . Explicit representation through heat kernel  $g(t, \mathbf{x}, \mathbf{y}) := (2\pi\kappa t)^{-d/2} \exp\{-\|\mathbf{x} - \mathbf{y}\|^2/2\kappa t\}$ :  $\vec{v}(t, \mathbf{x}) = -\frac{\kappa \int_{\mathbb{R}^d} \nabla g(t, \mathbf{x}, \mathbf{y}) e^{\xi(\mathbf{y})/\kappa} d\mathbf{y}}{\int_{-\epsilon} g(t, \mathbf{x}, \mathbf{y}) e^{\xi(\mathbf{y})/\kappa} d\mathbf{y}},$ 

• Parabolic scaling leads to the RF  $\vec{v}_{\lambda}(t, \mathbf{x}) := \vec{v}(\lambda^2 t, \lambda \mathbf{x})$  written as

$$\vec{v}_{\lambda}(t, \mathbf{x}) = -\frac{\kappa \int_{\mathbb{R}^d} \nabla g(t, \mathbf{x}, \mathbf{y}/\lambda) \mathrm{e}^{\xi(\mathbf{y})/\kappa} \mathrm{d}\mathbf{y}}{\lambda \int_{\mathbb{R}^d} g(t, \mathbf{x}, \mathbf{y}/\lambda) \mathrm{e}^{\xi(\mathbf{y})/\kappa} \mathrm{d}\mathbf{y}}.$$
(23)

• integrals in numerator and denominator resemble  $Y_{\lambda}(\phi) = \int_{\mathbb{R}^d} G(\xi(\mathbf{y}))\phi(\mathbf{y}/\lambda) d\mathbf{y}$ with  $G(\mathbf{x}) = e^{\mathbf{x}/\kappa}, \phi(\mathbf{y}) = \nabla g(t, \mathbf{x}, \mathbf{y})$  and  $\phi(\mathbf{y}) = g(t, \mathbf{x}, \mathbf{y})$ 

• For  $\kappa > 0$  fixed the limit distribution of  $\vec{v}_{\lambda}(t, \mathbf{x})$  was studied for several models of initial RF  $\xi = \{\xi(\mathbf{y}), \mathbf{y} \in \mathbb{R}^d\}$  with short and long range dependence [Gaussian, Gaussian subordinated, shot-noise, Cox]

 For κ > 0 fixed the limit distribution of v
<sub>λ</sub>(t, x) was studied for several models of initial RF ξ = {ξ(y), y ∈ ℝ<sup>d</sup>} with short and long range dependence [Gaussian, Gaussian subordinated, shot-noise, Cox]

Albeverio, S., Molchanov, S.A. & S.D. (1994) Stratified structure of the Universe and Burgers' equation - a probabilistic approach. Probab. Th. Rel. Fields 100

Funaki, T., S.D. & Woyczynski, W.A. (1995) Gibbs-Cox random fields and Burgers' turbulence. Ann. Appl. Probab. 5

Leonenko, N.N. & Woyczynski, W.A. (1998) Scaling limits of solutions of the heat equation for singular non-Gaussian data. J. Stat. Physics 91

 For κ > 0 fixed the limit distribution of v
<sub>λ</sub>(t, x) was studied for several models of initial RF ξ = {ξ(y), y ∈ ℝ<sup>d</sup>} with short and long range dependence [Gaussian, Gaussian subordinated, shot-noise, Cox]

Albeverio, S., Molchanov, S.A. & S.D. (1994) Stratified structure of the Universe and Burgers' equation - a probabilistic approach. Probab. Th. Rel. Fields 100

Funaki, T., S.D. & Woyczynski, W.A. (1995) Gibbs-Cox random fields and Burgers' turbulence. Ann. Appl. Probab. 5

Leonenko, N.N. & Woyczynski, W.A. (1998) Scaling limits of solutions of the heat equation for singular non-Gaussian data. J. Stat. Physics 91

. . .

Review paper:

S.D. & Woyczynski, W.A. (2003) Limit theorems for the Burgers equation initialized by data with long-range dependence. In: P. Doukhan, G. Oppenheim and M.S. Taqqu (Eds.) Long Range Dependence: Theory and Applications, pp. 507–523. Birkhäuser, Boston.

 For κ > 0 fixed the limit distribution of v
<sub>λ</sub>(t, x) was studied for several models of initial RF ξ = {ξ(y), y ∈ ℝ<sup>d</sup>} with short and long range dependence [Gaussian, Gaussian subordinated, shot-noise, Cox]

Albeverio, S., Molchanov, S.A. & S.D. (1994) Stratified structure of the Universe and Burgers' equation - a probabilistic approach. Probab. Th. Rel. Fields 100

Funaki, T., S.D. & Woyczynski, W.A. (1995) Gibbs-Cox random fields and Burgers' turbulence. Ann. Appl. Probab. 5

Leonenko, N.N. & Woyczynski, W.A. (1998) Scaling limits of solutions of the heat equation for singular non-Gaussian data. J. Stat. Physics 91

. . .

Review paper:

S.D. & Woyczynski, W.A. (2003) Limit theorems for the Burgers equation initialized by data with long-range dependence. In: P. Doukhan, G. Oppenheim and M.S. Taqqu (Eds.) Long Range Dependence: Theory and Applications, pp. 507–523. Birkhäuser, Boston.

• This talk: initial potential RF = aggregated RG model

$$\xi_{\mathcal{M}}(\boldsymbol{y}) := M^{-1/2}(X_{\mathcal{M}}(\boldsymbol{y}) - \mathbf{E}X_{\mathcal{M}}(\boldsymbol{y})), \quad \boldsymbol{y} \in \mathbb{R}^{d},$$
(24)

with intensity  $M = \lambda^{\gamma}$  increasing with  $\lambda$  for some  $\gamma > 0$ 

 For κ > 0 fixed the limit distribution of v
<sub>λ</sub>(t, x) was studied for several models of initial RF ξ = {ξ(y), y ∈ ℝ<sup>d</sup>} with short and long range dependence [Gaussian, Gaussian subordinated, shot-noise, Cox]

Albeverio, S., Molchanov, S.A. & S.D. (1994) Stratified structure of the Universe and Burgers' equation - a probabilistic approach. Probab. Th. Rel. Fields 100

Funaki, T., S.D. & Woyczynski, W.A. (1995) Gibbs-Cox random fields and Burgers' turbulence. Ann. Appl. Probab. 5

Leonenko, N.N. & Woyczynski, W.A. (1998) Scaling limits of solutions of the heat equation for singular non-Gaussian data. J. Stat. Physics 91

. . .

Review paper:

S.D. & Woyczynski, W.A. (2003) Limit theorems for the Burgers equation initialized by data with long-range dependence. In: P. Doukhan, G. Oppenheim and M.S. Taqqu (Eds.) Long Range Dependence: Theory and Applications, pp. 507–523. Birkhäuser, Boston.

• This talk: initial potential RF = aggregated RG model

$$\xi_M(\boldsymbol{y}) := M^{-1/2}(X_M(\boldsymbol{y}) - \mathbf{E}X_M(\boldsymbol{y})), \quad \boldsymbol{y} \in \mathbb{R}^d,$$
(24)

with intensity  $M = \lambda^{\gamma}$  increasing with  $\lambda$  for some  $\gamma > 0$ 

• The meaning of intial condition  $\vec{v}(0+, \mathbf{x}) = -\nabla \xi_M(\mathbf{x})$  ignored

# Corollary

1. Let  $\vec{v}_{\lambda}(t, \mathbf{x})$  be as in (23), (24), with  $X_M, M$  satisfying the conditions of Thm 1.

# Corollary

1. Let  $\vec{v}_{\lambda}(t, \mathbf{x})$  be as in (23), (24), with  $X_M, M$  satisfying the conditions of Thm 1. Then, as  $\lambda \to \infty$ , for any  $\gamma > 0$ 

$$\lambda^{1+d+\frac{\gamma}{2}-H(\gamma)}\vec{v}_{\lambda}(t,\boldsymbol{x}) \xrightarrow{\text{fdd}} \begin{cases} B_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma > d(\alpha-1), \\ L_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma < d(\alpha-1), \\ J_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma = d(\alpha-1), \end{cases}$$
(25)

## Corollary

1. Let  $\vec{v}_{\lambda}(t, \mathbf{x})$  be as in (23), (24), with  $X_M$ , M satisfying the conditions of Thm 1. Then, as  $\lambda \to \infty$ , for any  $\gamma > 0$ 

$$\lambda^{1+d+\frac{\gamma}{2}-H(\gamma)}\vec{v}_{\lambda}(t,\boldsymbol{x}) \xrightarrow{\text{fdd}} \begin{cases} B_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma > d(\alpha-1), \\ L_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma < d(\alpha-1), \\ J_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma = d(\alpha-1), \end{cases}$$
(25)

where  $H(\gamma)$  and the limit RFs are the same as in Thms 1-2.

## Corollary

1. Let  $\vec{v}_{\lambda}(t, \mathbf{x})$  be as in (23), (24), with  $X_M$ , M satisfying the conditions of Thm 1. Then, as  $\lambda \to \infty$ , for any  $\gamma > 0$ 

$$\lambda^{1+d+\frac{\gamma}{2}-H(\gamma)}\vec{v}_{\lambda}(t,\boldsymbol{x}) \xrightarrow{\text{fdd}} \begin{cases} B_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma > d(\alpha-1), \\ L_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma < d(\alpha-1), \\ J_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma = d(\alpha-1), \end{cases}$$
(25)

where  $H(\gamma)$  and the limit RFs are the same as in Thms 1-2.

2.  $\vec{v}_{\lambda}(t, \mathbf{x})$  be as in (23) with  $\xi(\mathbf{y}) = X(\mathbf{y})$  given in (8) (M = 1).

## Corollary

1. Let  $\vec{v}_{\lambda}(t, \mathbf{x})$  be as in (23), (24), with  $X_M$ , M satisfying the conditions of Thm 1. Then, as  $\lambda \to \infty$ , for any  $\gamma > 0$ 

$$\lambda^{1+d+\frac{\gamma}{2}-H(\gamma)}\vec{v}_{\lambda}(t,\boldsymbol{x}) \xrightarrow{\text{fdd}} \begin{cases} B_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma > d(\alpha-1), \\ L_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma < d(\alpha-1), \\ J_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \quad \gamma = d(\alpha-1), \end{cases}$$
(25)

where  $H(\gamma)$  and the limit RFs are the same as in Thms 1-2.

2.  $\vec{v}_{\lambda}(t, \mathbf{x})$  be as in (23) with  $\xi(\mathbf{y}) = X(\mathbf{y})$  given in (8) (M = 1). Then, as  $\lambda \to \infty$ 

$$\lambda^{1+d-\frac{d}{\alpha}}\vec{v}_{\lambda}(t,\boldsymbol{x}) \quad \stackrel{\text{fdd}}{\longrightarrow} \quad \kappa(\mathrm{e}^{1/\kappa}-1)L_{\alpha}(\nabla g(t,\boldsymbol{x},\cdot)), \tag{26}$$

where  $L_{\alpha}$  is  $\alpha$ -stable RF as in part 1.

## 6. Open questions

• Thm 2 yields trivial limits if respective Hermite/Charlier coefficients vanish:  $h_{G,\mu}(1) = 0$  or  $c_{G,\mu}(1) = 0$ . • Thm 2 yields trivial limits if respective Hermite/Charlier coefficients vanish:  $h_{G,\mu}(1) = 0$  or  $c_{G,\mu}(1) = 0$ . Limit distribution of  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} G(X_M(t))\phi(t/\lambda) dt$  in such cases is open • Thm 2 yields trivial limits if respective Hermite/Charlier coefficients vanish:  $h_{G,\mu}(1) = 0$  or  $c_{G,\mu}(1) = 0$ . Limit distribution of  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} G(X_M(t))\phi(t/\lambda) dt$  in such cases is open (Gaussian/Poisson chaos?)

- Thm 2 yields trivial limits if respective Hermite/Charlier coefficients vanish:  $h_{G,\mu}(1) = 0$  or  $c_{G,\mu}(1) = 0$ . Limit distribution of  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} G(X_M(t))\phi(t/\lambda) dt$  in such cases is open (Gaussian/Poisson chaos?)
- 2 Aggregated small-scale limits:  $\lambda \to 0$  together with  $M = \lambda^{\gamma} \to 0$ ?

- **1** Thm 2 yields trivial limits if respective Hermite/Charlier coefficients vanish:  $h_{G,\mu}(1) = 0$  or  $c_{G,\mu}(1) = 0$ . Limit distribution of  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} G(X_M(t))\phi(t/\lambda) dt$  in such cases is open (Gaussian/Poisson chaos?)
- 2 Aggregated small-scale limits:  $\lambda \to 0$  together with  $M = \lambda^{\gamma} \to 0$ ? For linear integrals  $X_{\lambda,M}(\phi)$ : Biermé, H., Estrade, A. & Kaj, I. (2010) Self-similar random fields and rescaled random balls models. J. Theoret. Probab. 23

- **1** Thm 2 yields trivial limits if respective Hermite/Charlier coefficients vanish:  $h_{G,\mu}(1) = 0$  or  $c_{G,\mu}(1) = 0$ . Limit distribution of  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} G(X_M(t))\phi(t/\lambda) dt$  in such cases is open (Gaussian/Poisson chaos?)
- 2 Aggregated small-scale limits:  $\lambda \to 0$  together with  $M = \lambda^{\gamma} \to 0$ ? For linear integrals  $X_{\lambda,M}(\phi)$ : Biermé, H., Estrade, A. & Kaj, I. (2010) Self-similar random fields and rescaled random balls models. J. Theoret. Probab. 23

Nonlinear integrals:  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} G(X_M(t))\phi(t/\lambda) dt$ ?

- **1** Thm 2 yields trivial limits if respective Hermite/Charlier coefficients vanish:  $h_{G,\mu}(1) = 0$  or  $c_{G,\mu}(1) = 0$ . Limit distribution of  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} G(X_M(t))\phi(t/\lambda) dt$  in such cases is open (Gaussian/Poisson chaos?)
- 2 Aggregated small-scale limits:  $\lambda \to 0$  together with  $M = \lambda^{\gamma} \to 0$ ? For linear integrals  $X_{\lambda,M}(\phi)$ : Biermé, H., Estrade, A. & Kaj, I. (2010) Self-similar random fields and rescaled random balls models. J. Theoret. Probab. 23 Nonlinear integrals:  $Y_{\lambda,M}(\phi) = \int_{\mathbb{R}^d} G(X_M(t))\phi(t/\lambda) dt$ ?
- Ox RG model: Poisson grains with random intensity