

An interacting neuronal network with inhibition: theoretical analysis and perfect simulation

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Workshop : New trends on Hawkes processes

June, 2022



- 1 The model
- 2 Perfect simulation

Introduction

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When a neuron spikes, this spike changes the rate at which other neurons spike. In other words, neurons excite and/or **inhibit** each other.

- 1 The model
 - Description of the model
 - Infinitesimal generator
 - First jump time
- 2 Perfect simulation

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When the neuron i spikes,

- The current state of inhibition of neuron i is replaced by a new value Y^i independently of anything else with distribution F^i .

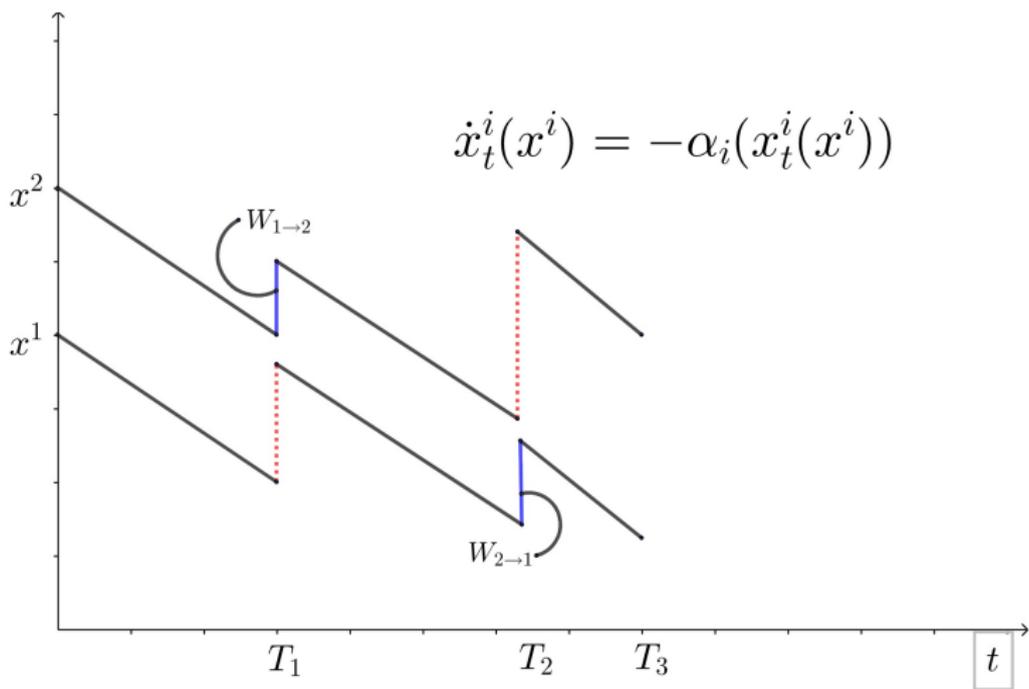
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When the neuron i spikes,

- The current state of inhibition of neuron i is replaced by a new value Y^i independently of anything else with distribution F^i .
- The state of inhibition of any neuron $j \neq i$ is increased by a positive value $W_{i \rightarrow j}$ at time t .

Description of the model



Description of the model

We consider the piecewise deterministic Markov process (PDMP)

$X_t^N = (X_t^{1,N}, \dots, X_t^{N,N}) \in \mathbb{R}_+^N$. For $i \in \{1, \dots, N\}$, the dynamic of $X_t^{i,N}$ is given by:

$$dX_t^{i,N} = -\alpha_i \left(X_{t-}^{i,N} \right) dt + \int_0^\infty \int_0^\infty (y^i - X_{t-}^{i,N}) \mathbf{1}_{\{r \leq \beta_i(X_{t-}^{i,N})\}} M^i(dt, dr, dy^i) \\ + \sum_{j \neq i} W_{j \rightarrow i} \int_0^\infty \int_0^\infty \mathbf{1}_{\{r \leq \beta_j(X_{t-}^{j,N})\}} M^j(dt, dr, dy^j), \quad (1)$$

where M^i is a random Poisson measure with intensity $dt dr F^i(dy)$ and for all i , the M^i are all independent.

Infinitesimal generator

The infinitesimal generator associated with this model is given by:

$$G^N V(x) = - \sum_{i=1}^N \alpha_i(x^i) \frac{\partial}{\partial x^i} V(x) + \sum_{i=1}^N \beta_i(x^i) \int_0^\infty F^i(dy^i) [V(x + e^i y^i - e^i x^i + \sum_{j \neq i} e^j W_{i \rightarrow j}) - V(x)] \quad (2)$$

where V is a smooth function and e^i is the i -th unit vector.

First jump time

Let S_1 be the first jump time of the first neuron to jump. For all $t > 0$,

$$\mathbb{P}(S_1 > t) = \prod_{i=1}^N e^{-\int_0^t \beta_i(x_s^i(x^i)) ds}. \quad (3)$$

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Moreover, if $t < \min_i t_0(x^i)$ where

$$t_0(x^i) := \int_0^{x^i} \frac{dy}{\alpha_i(y)}$$

is the time for the neuron i hit 0 starting from x^i ,

First jump time

We can write

$$\mathbb{P}(S_1 > t) = \prod_{i=1}^N e^{-[\Gamma_i(x^i) - \Gamma_i(x_t^i(x^i))]},$$

with $\Gamma_i(x^i) := \int^{x^i} \gamma_i(y) dy$ and $\gamma_i(x^i) = \beta_i(x^i) / \alpha_i(x^i)$.

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Then, $\mathbb{P}(S_1 = \infty) = 0$ if and only if

Assumption 1

$\Gamma_i(0) = -\infty$ for all $1 \leq i \leq N$

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- We explore the past in order to determine all sets of indices and times which affect the value of neuron i at time 0. The set of all such couples (j, s) will be called **the clan of ancestors of neuron i** . (see Galves and Löcherbach [2], Galves et al. [3])

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- We consider the case where each neuron i has exactly two neighbours so that the neuron i interacts only with the neurons $i + 1$ and $i - 1$.
- T is the time vector.

Algorithm (backward procedure)

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- We simulate , $\forall l \in \mathbb{Z}$, $N_t^{l,S}$ and $N_t^{l,P}$ two Poisson processes with respective intensities β_* and $\beta^* - \beta_*$.

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- The jump times $T_n^{l,S}$ will be considered as times of **sure jumps** (counted by the process $N_t^{l,S}$) and the jump times $T_n^{l,P}$ will be considered as times of **possible jumps** (counted by the process $N_t^{l,P}$)

Algorithm (backward procedure)

- Let $i \in \mathbb{Z}$ fix and

$$T_1 = \inf\{T_1^{i\pm 1,s}, T_1^{i\pm 1,p}, T_1^{i,s}, T_1^{i,p}\}.$$

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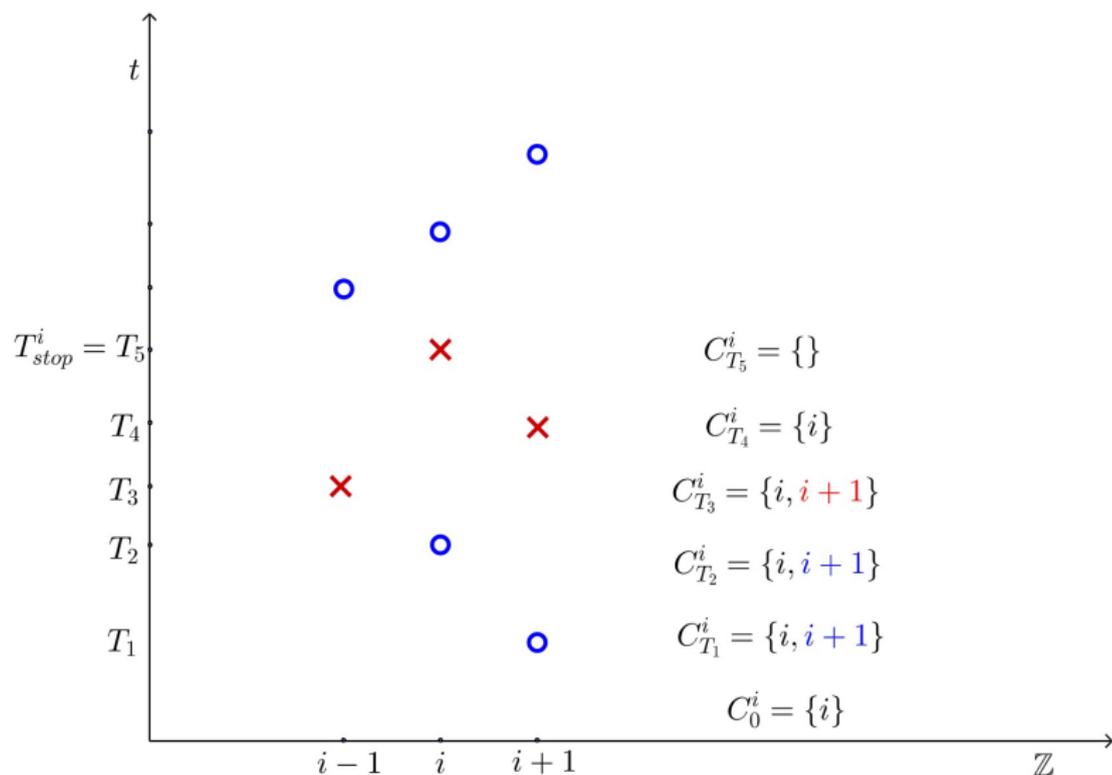
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- If $T_1 = T_1^{i,p}$, we set $C_{T_1}^i = \{i\}$. We put $l_1 = i$.
- If $T_1 = T_1^{i,s}$, we set $C_{T_1}^i = \emptyset$ and we stop the algorithm. We put $l_1 = i$.

Algorithm (backward procedure)



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Suppose T_n is the n -th jump time of $C_{T_n}^i$. We have:

$$T_{n+1} = \inf\{T_m^{j,s}, T_m^{j,p} > T_n : |j - C_{T_n}^i| \leq 1, T_m^{k,s} > T_n, k \in C_{T_n}^i\}.$$

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- If $T_{n+1} = T_m^{j,p}$ we set:

$$\begin{cases} \text{If } j \in C_{T_n}^i, & C_{T_{n+1}}^i = C_{T_n}^i \\ \text{If } j \notin C_{T_n}^i, & C_{T_{n+1}}^i = C_{T_n}^i \cup \{j\} \end{cases}$$

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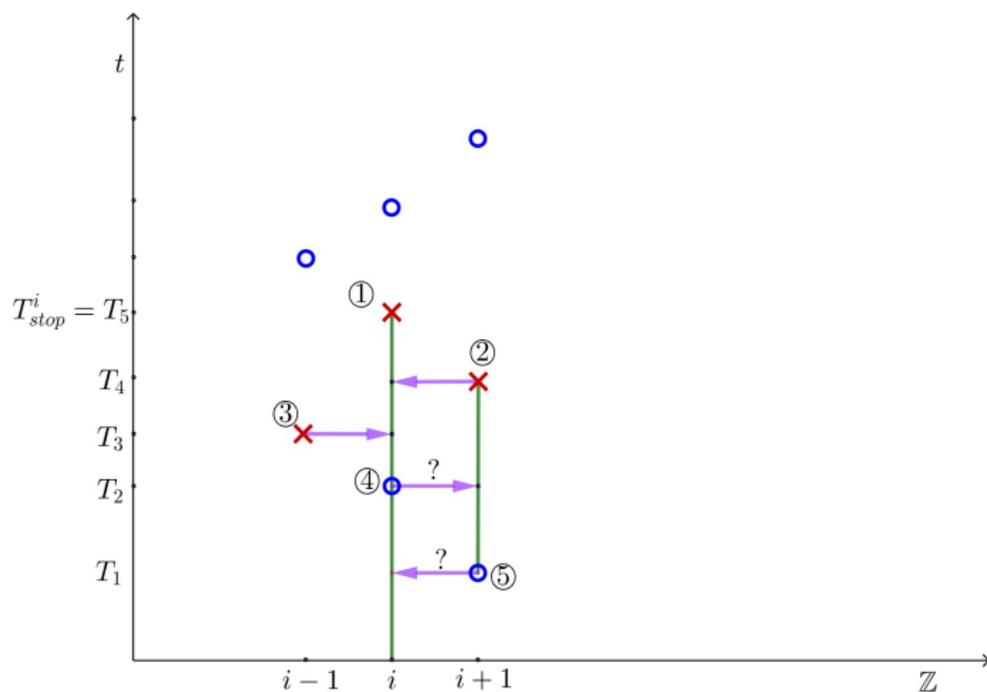
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- If $T_{n+1} = T_m^{k,s}$ we set:

$$\begin{cases} \text{If } k \in C_{T_n}^i, & C_{T_{n+1}}^i = C_{T_n}^i \setminus \{k\} \\ \text{If } k \notin C_{T_n}^i, & C_{T_{n+1}}^i = C_{T_n}^i \end{cases}$$

Forward procedure



Extinction time T_{stop}^i

The procedure will stop at time T_{stop}^i where $T_{stop}^i = \inf\{t : C_t^i = \emptyset\}$.

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Theorem 1

We set $\delta = \frac{\beta_*}{\beta^* - \beta_*}$. There exists a critical value $0 < \delta_c < \infty$ such that:

- if $\delta > \delta_c$, then the extinction time is finite almost surely that is, $\mathbb{P}(\forall i, T_{stop}^i < \infty) = 1$
- if $\delta < \delta_c$, then the extinction time is infinite with a positive probability that is, $\mathbb{P}(\forall i, T_{stop}^i = \infty) > 0$.

Sketch of proof

- We first show that $T_{stop}^i < +\infty$ almost surely for sufficiently large δ .
- We can upper bound $|C_t^i|$ by Z_t almost surely for all $t \geq 0$ where $Z_0 = 1$ and $(Z_t)_{t \geq 0}$ is a branching process.

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- The associated infinitesimal generator of $(Z_t)_{t \geq 0}$ as follows :

$$Af(n) = n[(\beta^* - \beta_*)(f(n+1) - f(n)) + \beta_*(f(n-1) - f(n))].$$

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- If we take, $f(n) = n$, then, for $\delta > 1$, we have

$$Af(n) = -cf(n)$$

where $-c = (\beta^* - \beta_*)(1 - \delta)$.

Sketch of proof

- Assume $x_t = \mathbb{E}(f(Z_t))$ and using the Itô formula, we have:

$$x_t = x_0 + \mathbb{E} \int_0^t Af(Z_s) ds = x_0 - c \int_0^t x_s ds = x_0 e^{-ct}.$$

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- When $t \rightarrow \infty$, we have $x_t \rightarrow 0$.

- If $\delta > 1$, $\mathbb{P}(T_{stop}^i < \infty) \geq \mathbb{P}(\lim_{t \rightarrow \infty} Z_t = 0) = 1$ thus ensuring that $\delta_c \leq 1$.

Sketch of proof

We show that for all $\delta < \delta_c$, $T_{stop}^i = +\infty$ with positive probability.

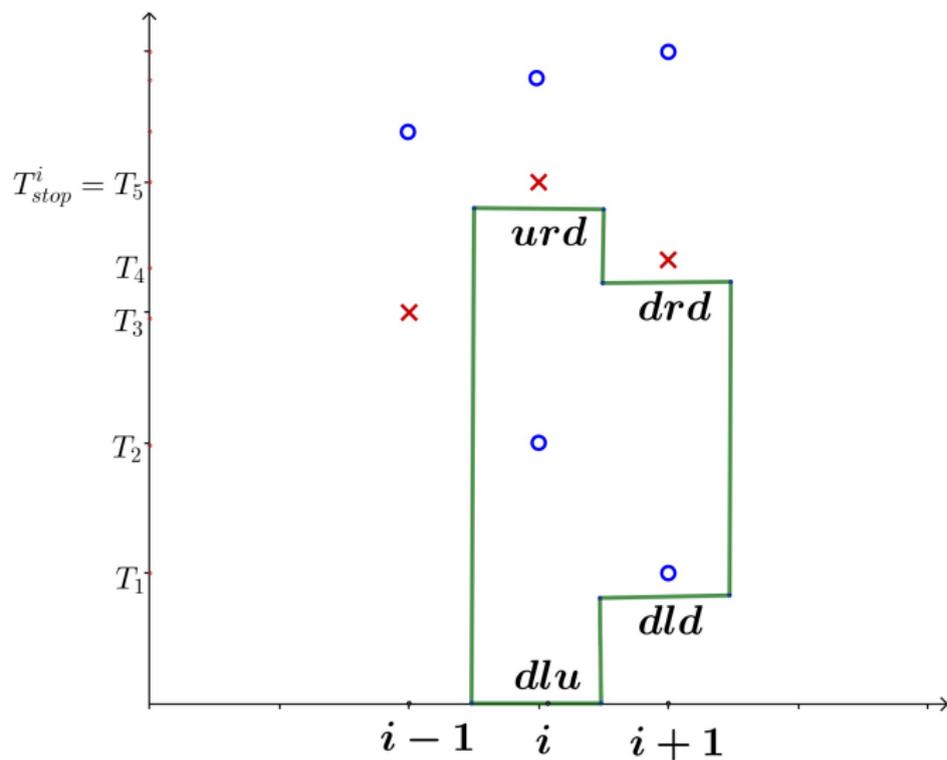
Sketch of proof

We show that for all $\delta < \delta_c$, $T_{stop}^i = +\infty$ with positive probability.

$T_{stop}^i < \infty$ if and only if $\bar{C}^i = \cup_{t \geq 0} C_t^i$ is a finite set.

- We show that $\mathbb{P}(T_{stop}^i < \infty) = \mathbb{P}(|\bar{C}^i| < \infty) < 1$ for sufficiently small values of δ using classical contour techniques. (see Ferrari et al.[1])
- For this, on $|\bar{C}^i| < \infty$, we draw the contour of \bar{C}^i as follow.

Sketch of proof/Contour



Sketch of proof

- Set $N(dlu), N(urd), \dots$ the number of appearances of the different direction vectors.

$$N(dru) = N(urd) - 1 \leq n/2, \quad N(drd) + N(dru) + N(uru) + N(urd) = n.$$

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- Set $N(dlu), N(urd), \dots$ the number of appearances of the different direction vectors.

$$N(dru) = N(urd) - 1 \leq n/2, \quad N(drd) + N(dru) + N(uru) + N(urd) = n.$$

- For $n = 1$, the probability of appearance of a contour of length 4 is equal to $\mathbb{P}(D_1 = urd) \leq \delta$.

Sketch of proof

- For $n = 2$, the probability of appearance of a contour of length 8 is equal to

$$\begin{aligned} & \mathbb{P}(D_1 = ulu, D_2 = urd, D_3 = drd) + \mathbb{P}(D_1 = ulu, D_2 = uru, D_3 = urd) + \\ & \mathbb{P}(D_1 = uru, D_2 = urd, D_3 = dld) + \mathbb{P}(D_1 = urd, D_2 = drd, D_3 = dld) \\ & \leq 4\delta^2. \end{aligned}$$

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- For $n = 2$, the probability of appearance of a contour of length 8 is equal to

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- A very approximate upper bound on the total number of possible triplets is given by $4^{2n} = 16^n$. We get for all $\delta < \frac{1}{(16)^2}$,

$$\mathbb{P}(T_{stop}^i < \infty) \leq \delta + 4\delta^2 + \frac{(16\sqrt{\delta})^3}{1 - 16\sqrt{\delta}} =: \phi(\delta) \rightarrow 0 \text{ when } \delta \rightarrow 0$$

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Thank you for your attention!

Questions?