

Separable Mixing

P4.1

weighted at random \Rightarrow scalar variables
(despite heterogeneity)

See Section 8.4 of DHB book mentioned on P 2.11

Assume

$$A(\tau, x, \xi) = a(x) g(\tau, \xi)$$

Interpretation: the properties of the two individuals involved have independent influence on the likelihood of transmission.

Consequences: ① The operators K_λ have one-dimensional range spanned by a . Hence

$$R_0 = N \int_{\mathbb{R}} \left(\int_0^\infty g(\tau, \xi) d\tau \right) a(\xi) \Phi(d\xi)$$

$$\text{Euler-Lotka } r = N \int_{\mathbb{R}} \left(\int_0^\infty g(\tau, \xi) e^{-\lambda \tau} d\tau \right) a(\xi) \Phi(d\xi)$$

$\text{sign}(R_0 - 1) = \text{sign } r$ is easy to prove

②

$$s(t, x) = e^{-a(x) w(t)}$$

with w a solution of the

scalar RE

$$w(t) = N \int_0^\infty \left\{ g(\tau, \xi) \right\} \left[1 - e^{-a(\xi) w(t-\tau)} \right] \Phi(d\xi) d\tau$$

with $w(-\infty) = 0$
So the cumulative $F_{w(t)}$ on individuals with trait x equals $a(x) w(t)$

Note that we can discretize the RE for w as described on P 1.10

$$(3) \text{ Final Size : } s(\infty, x) = e^{-\alpha(x)w(\infty)} \quad \text{with} \quad (P4.2)$$

$$w(\infty) = N \int_{\mathbb{R}} \left(\left(\int_0^{\infty} g(\tau, \xi) d\tau \right) \left[1 - e^{-\alpha(\xi)w(\infty)} \right] \Phi(d\xi) \right)$$

Assume

$$g(\tau, \xi) = b(\tau)c(\xi)$$

Define

$$\Psi(w) := \int_{\mathbb{R}} c(\xi) \left(1 - e^{-\alpha(\xi)w} \right) \Phi(d\xi)$$

then the RE for w reads

$$w(t) = N \int_0^{\infty} b(\tau) \Psi(w(t-\tau)) d\tau$$

Gamma Distribution

- J.O. Lloyd-Smith, S.J. Schreiber, P.E. Kopp, W.M. Getz
Superspreading and the effect of individual variation on disease emergence, *Nature* (2005) 438 (7066) : 355–359
- A.S. Novozhilov, On the spread of epidemics in closed heterogeneous populations, *Math. Biosc.* 215 (2008) 177–185
- M.G.M. Gomes & various co-authors, preprints on medRxiv, arXiv
- A.V. Tkachenko et al., Time-dependent heterogeneity leads to transient suppression of the COVID-19 epidemic, not herd immunity, *PNAS* (2021) 118 (17) e2015972118
- J. Neipel et al., Power-Law population heterogeneity governs epidemic waves, *PLoS ONE* (2020) 15 (10) e0239678
- T. Britton, F. Ball, P. Trapman, A mathematical model reveals the influence of population heterogeneity on herd immunity to SARS-CoV2 *Science* (2020) 369 : 846–849

$\Omega = (0, \infty)$ Φ is the Gamma Distribution with

P4.3

mean = 1

so Φ has density $x \mapsto \frac{P}{\Gamma(P)} x^{P-1} e^{-Px}$

and

Variance = $\frac{1}{P}$

So small P corresponds to much heterogeneity!

The Laplace Transform is $\bar{\Phi}(\lambda) = \left(\frac{\lambda}{P} + 1\right)^{-P}$

and $\bar{\Phi}'(\lambda) = -\left(\frac{\lambda}{P} + 1\right)^{-P-1}$, so $\bar{\Phi}'(0) = -1$,

$$\bar{\Phi}''(\lambda) = \frac{P+1}{P} \left(\frac{\lambda}{P} + 1\right)^{-P-2}$$

We discuss / contrast the following three cases

I. no heterogeneity : $a(\xi) = 1$, $c(\xi) = 1$

II. heterogeneity in susceptibility, but not in infectiousness : $a(\xi) = \xi$, $c(\xi) = 1$

III. correlated heterogeneity in susceptibility and infectiousness : $a(\xi) = \xi$, $c(\xi) = \xi$

in particular with respect to the

Effective Reproduction Number

and the

Herd Immunity Threshold

On page 1.3 the Basic Reproduction Number R_0 was defined as the expected number of secondary cases per primary case. It was left implicit that this concerns the start of an outbreak in a fully susceptible host population. On P3.2,3.3 we considered a partly vaccinated population and introduced \tilde{R}_0 as a first example of an Effective Reproduction Number. Now we want to consider an ongoing outbreak and try to quantify the effect of immunity of those that were infected so far. To do so, we simply pretend that the current susceptible subpopulation is the entire population in which the pathogen is introduced and call the corresponding reproduction number $R_{\text{eff}}(t)$. For instance, for the SIR model of page 1.5 we have

$$R_{\text{eff}}(t) = \frac{\beta S(t)}{\alpha} = s(t) R_0 \quad \text{with } s(t) := \frac{S(t)}{N}$$

The Herd Immunity Threshold corresponds, by definition, to $R_{\text{eff}}(t) = 1$ so in the present special case to

$$s(t) = \frac{1}{R_0}$$

(Please keep in mind that upon reaching the HIT during an actual outbreak, the epidemic does not stop, there is an overshoot due to the reservoir of already infected individuals, see the phase plane picture on page 1.5.)

For the RE for w from P4.2 we define

$$\begin{aligned} R_{\text{eff}}(t) &= N \int_0^\infty b(c) dc \Psi'(w(t)) = R_0 \frac{\Psi''(w(t))}{\Psi'(0)} \\ &= R_0 \frac{\int_0^t c(\xi) \alpha(\xi) e^{-\alpha(\xi) w(t)} \Phi(d\xi)}{\int_0^t c(\xi) \alpha(\xi) \Phi(d\xi)} \end{aligned}$$

We consider $R_{\text{eff}} = 1$ as an equation for w P 4.5
 and denote the solution of this equation by \bar{w} .
 The fraction of the population that is still susceptible
 when the HIT is reached is given by

$$\bar{s} = \int_0^{\infty} e^{-a(\xi)\bar{w}} \Phi(d\xi)$$

	$\Psi(w)$	R_0	$R_{\text{eff}}(w)$	\bar{w}	\bar{s}
I	$1 - e^{-w}$	1	$R_0 e^{-w}$	$\ln R_0$	R_0^{-1}
II	$1 - \left(\frac{w}{p} + 1\right)^{-p}$	1	$R_0 \left(\frac{w}{p} + 1\right)^{-p-1}$	$p(R_0^{\frac{1}{p+1}} - 1)$	$R_0^{-1 + \frac{1}{p+1}}$
III	$1 - \left(\frac{w}{p} + 1\right)^{-p-1}$	$1 + \frac{1}{p}$	$R_0 \left(\frac{w}{p} + 1\right)^{-p-2}$	$p(R_0^{\frac{1}{p+2}} - 1)$	$R_0^{-1 + \frac{1}{2p+1}}$

Note that in the "uniform"/homogeneous case I, the HIT for vaccination and for natural immunity(induced by infection) is exactly the same, while for cases II, III and random vaccination, the vaccination HIT R_0^{-1} differs substantially from the HIT for infection induced immunity, especially for small $p \sim$ strong heterogeneity. The reason is, of course, that among the individuals infected so far the highly susceptible individuals are over-represented. This effect plays already a role in the early stage of the outbreak. This can be seen by expanding the

formula for $R_{\text{eff}}(t)$ at the bottom of P4.4 while P4.6
 assuming that $w(t)$ is small and, likewise, expanding
 $s(t) = \int_0^\infty e^{-\alpha(\xi)w(t)} \Phi(d\xi) = 1 - \int_0^\infty \alpha(\xi) \Phi(d\xi) w(t) + \dots$

Eliminating $w(t)$, and using that for all three cases
 $\int_0^\infty \alpha(\xi) \Phi(d\xi) = 1$, we obtain

$$R_{\text{eff}}(t) = R_0 \left(1 - \Theta(1-s(t)) + o(1-s(t)) \right)$$

with

$$\Theta = \frac{\int_0^\infty c(\xi) \alpha^2(\xi) \Phi(d\xi)}{\int_0^\infty c(\xi) \alpha(\xi) \Phi(d\xi)} = \begin{cases} 1 & \text{case I} \\ 1 + \frac{1}{p} & \text{II} \\ 1 + \frac{2}{p} & \text{III} \end{cases}$$

A.V. Tkachenko et al. write: "We named the coefficient Θ the **immunity factor** because it quantifies the effect that a gradual buildup of population immunity has on the spread of an epidemic."

We re-iterate the message of the Exercise on P3.5,3.6:
 if the most susceptible individuals are also the most infectious (in the sense of having many contacts) individuals, heterogeneity has a large impact on R_0 , see second column, third row entry of the matrix scheme.

Assume

$$b(\tau) = \beta e^{-\alpha \tau} \Rightarrow \frac{dw}{dt} = -\alpha w + \beta N \Psi(w)$$

Assume

$$b(\tau) = \beta \frac{\gamma}{\alpha-\gamma} (e^{-\gamma \tau} - e^{-\alpha \tau}) \Rightarrow \begin{aligned} \frac{dw}{dt} &= f \\ \frac{df}{dt} &= \beta N \gamma \Psi(w) \\ &\quad - (\alpha + \gamma) f - \alpha \gamma w \end{aligned}$$

In the paper of J. Neipel e.a., see P4.2, the SIR ODE is related to Covid data of Germany. P4.7

See A. Cori, N.M. Ferguson, C. Fraser, S. Cauchemez

A new framework and software to estimate time-varying reproduction numbers during epidemics

American J. of Epidemiology (2013) 178(9): 1505-1512

for general methodology for estimating R_{eff} from data.

Other relevant references are

G.N. Páez e.a., Alternative strategies for the estimation of a disease's basic reproduction number: a model-

agnostic study, Bull. Math. Biol. (2021) 83: 89

C. Vegvari e.a., Commentary on the use of the reproduction number R during the Covid-19 pandemic

Statistical Methods in Medical Research (2021) 096228022, 110370

F. Vanni e.a., On the use of aggregated human mobility data to estimate the reproduction number

Scientific Reports (Nature) (2021) 11: 23286

H. Nishiura, G. Chowell, The effective reproduction number as a prelude to statistical estimation of time-dependent epidemic trends, pp 103-121 In: G. Chowell e.a., editors, Mathematical and Statistical Estimation Approaches in Epidemiology, Springer, 2009