Optimal Control by Franchise and Deductible Amounts in the Classical Risk Model

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Classical Risk Model

- the insurance company has an initial surplus x > 0
- X_t(x) is a surplus at time t ≥ 0 provided that the initial surplus equals x

Premium arrivals

• the insurance company receives premiums with constant intensity c>0

Claim arrivals

- the claim sizes form a sequence $(Y_i)_{i\geq 1}$ of nonnegative i.i.d. random variables with c.d.f. $F(y) = \mathbb{P}[Y_i \leq y]$ and finite expectations μ ; τ_i is the time when the *i*th claim arrives
- the number of claims on the time interval [0, t] is a Poisson process $(N_t)_{t\geq 0}$ with constant intensity $\lambda > 0$; the random variables Y_i , $i \geq 1$, and the process $(N_t)_{t\geq 0}$ are independent
- the total claims on [0, t] equal $\sum_{i=1}^{N_t} Y_i$; we set $\sum_{i=1}^{0} Y_i = 0$ if $N_t = 0$

Classical Risk Model

The surplus of the insurance company at time t equals

$$X_t(x) = x + ct - \sum_{i=1}^{N_t} Y_i, \quad t \ge 0.$$
 (1)

We assume that the net profit condition holds, i.e.

 $c > \lambda \mu$.

The insurance company uses **the expected value principle** for premium calculation, i.e.

$$c = \lambda \mu (1 + \theta),$$

where $\theta > 0$ is a safety loading.

- optimal control by investments: C. Hipp, M. Plum (2000);
 C. Hipp, M. Plum (2003); C. S. Liu, H. Yang (2004); P. Azcue,
 N. Muler (2009)
- optimal control by reinsurance: H. Schmidli (2001); C. Hipp, M. Vogt (2003)
- optimal control by investments and reinsurance: H. Schmidli (2002); M. I. Taksar, C. Markussen (2003); S. D. Promislow, V. R. Young (2005); C. Hipp, M. Taksar (2010)

Definitions

- A franchise is a provision in the insurance policy whereby the insurer does not pay unless damage exceeds the franchise amount.
- A deductible is a provision in the insurance policy whereby the insurer pays any amounts of damage that exceed the deductible amount.

Example 1

The franchise/deductible amount is 10.

- Case 1: the claim size is 5
 - If the franchise is used, then the insurance company pays nothing.
 - If the deductible is used, then the insurance company pays nothing.
- Case 2: the claim size is 100
 - If the franchise is used, then the insurance company pays 100.
 - If the deductible is used, then the insurance company pays 100 10 = 90.

Motivation

- a franchise and a deductible are applied when the insured's losses are relatively small to deter a large number of trivial claims
- a deductible encourages the insured to take more care of the insured property

Optimal Control by the Franchise Amount

Additional assumptions

- the insurance company adjusts the franchise amount d_t at every time $t \ge 0$ on the basis of the information available before time t, i.e. every admissible strategy $(d_t)_{t\ge 0}$ ((d_t) for brevity) of the franchise amount choice is a predictable process w.r.t. the natural filtration generated by $(N_t)_{t\ge 0}$ and $(Y_i)_{i\ge 1}$
- $0 \le d_t \le d_{\max}$, where d_{\max} is the maximum allowed franchise amount such that $0 < F(d_{\max}) < 1$; in particular, if $d_t = 0$, then the franchise is not used at time t
- the safety loading $\theta > 0$ is constant

The premium intensity at time t depends on the franchise amount at this time and it is given by

$$c(d_t) = \lambda(1+\theta) \int_{d_t}^{+\infty} y \,\mathrm{d}F(y).$$

Optimal Control by the Franchise Amount

Let $X_t^{(d_t)}(x)$ be the surplus of the insurance company at time t provided its initial surplus is x and the strategy (d_t) is used. Then

$$X_t^{(d_t)}(x) = x + \int_0^t c(d_s) \, \mathrm{d}s - \sum_{i=1}^{N_t} Y_i \, \mathbb{I}_{\{Y_i > d_{\tau_i}\}}, \quad t \ge 0.$$
 (2)

The ruin time under the admissible strategy (d_t) is defined as

$$\tau^{(d_t)}(x) = \inf \{ t \ge 0 \colon X_t^{(d_t)}(x) < 0 \}.$$

The corresponding infinite-horizon survival probability is given by

$$\varphi^{(d_t)}(x) = \mathbb{P}\big[\tau^{(d_t)}(x) = \infty\big].$$

Our **aim** is to maximize the survival probability over all admissible strategies (d_t) , i.e. to find

$$\varphi^*(x) = \sup_{(d_t)} \varphi^{(d_t)}(x),$$

and show that there exists an optimal strategy (d_t^*) such that $\varphi^*(x) = \varphi^{(d_t^*)}(x)$ for all $x \ge 0$.

The optimal strategy will be a function of the initial surplus only.

Proposition 1

Let the surplus process $(X_t^{(d_t)}(x))_{t\geq 0}$ follow (2) under the above assumptions. If $\varphi^*(x)$ is differentiable on \mathbb{R}_+ , then it satisfies the Hamilton-Jacobi-Bellman equation

$$\sup_{d\in[0,\,d_{\max}]} \left((1+\theta) \int_{d}^{+\infty} y \,\mathrm{d}F(y) \,\left(\varphi^*(x)\right)' + \left(F(d)-1\right) \varphi^*(x) + \int_{d}^{d\vee x} \varphi^*(x-y) \,\mathrm{d}F(y) \right) = 0,$$
(3)

which is equivalent to

$$\left(\varphi^*(x)\right)' = \inf_{d \in [0, d_{\max}]} \left(\frac{\left(1 - F(d)\right)\varphi^*(x) - \int_d^{d \lor x} \varphi^*(x - y) \,\mathrm{d}F(y)}{\left(1 + \theta\right) \int_d^{+\infty} y \,\mathrm{d}F(y)}\right).$$
(4)

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Remark 1

Note that if there exists one solution to (3) or (4), then there exist infinitely many solutions to these equations which differ with a multiplicative constant.

If the random variables Y_i , $i \ge 1$, have a p.d.f. f(y), then there exists the solution G(x) to (4) with $G(0) = \theta/(1+\theta)$, which is nondecreasing and continuously differentiable on \mathbb{R}_+ , and $\theta/(1+\theta) \le \lim_{x\to+\infty} G(x) \le 1$.

The solution G(x) to (4) that satisfies the conditions of Theorem 1 can be found as the limit of the sequence of functions $(G_n(x))_{n\geq 0}$ on \mathbb{R}_+ , where $G_0(x) = \varphi^{(0)}(x)$ is the survival probability provided that $d_t = 0$ for all $t \geq 0$, and

$$G'_n(x) = \inf_{d \in [0, d_{\max}]} \left(\frac{(1 - F(d)) G_{n-1}(x) - \int_d^{d \lor x} G_{n-1}(x - y) dF(y)}{(1 + \theta) \int_d^{+\infty} y dF(y)} \right),$$

$$G_n(0) = \theta / (1 + \theta), \quad n \ge 1.$$

(5)

Let the surplus process $(X_t^{(d_t)}(x))_{t\geq 0}$ follow (2) and G(x) be the solution to (4) that satisfies the conditions of Theorem 1. Then for any $x \geq 0$ and arbitrary admissible strategy (d_t) , we have

$$\varphi^{(d_t)}(x) \le \frac{G(x)}{\lim_{x \to +\infty} G(x)},$$
(6)

and equality in (6) is attained under the strategy $(d_t^*) = (d_t^*(X_{t_-}^{(d_t^*)}(x)))$, where $(d_t^*(x))$ minimizes the right-hand side of (4), i.e.

$$\varphi^*(x) = \varphi^{(d_t^*)}(x) = \frac{G(x)}{\lim_{x \to +\infty} G(x)}.$$

Remark 2

In Theorem 2 we used any solution to (4) that satisfies the conditions of Theorem 1. However, Theorem 2 also implies uniqueness of such a solution. The corresponding strategy (d_t^*) may not be unique in the general case.

Let the surplus process $(X_t^{(d_t)}(x))_{t\geq 0}$ follow (2), the claim sizes be exponentially distributed with mean μ , and $d_{\max} = \mu$. Then the strategy (d_t) with $d_t = 0$ for all $t \geq 0$ is not optimal.

Remark 3

Theorem 3 implies that we can always increase the survival probability adjusting the franchise amount if the claim sizes are exponentially distributed.

Example 2

If the claim sizes are exponentially distributed with mean $\mu=$ 10, $d_{\rm max}=\mu,$ and $\theta=$ 0.1, then

$$arphi^{(0)}(x) pprox 1 - 0.9090909 \,\mathrm{e}^{-x/110} \,, \quad x \ge 0 \,,$$
 $arphi^*(x) pprox \begin{cases} 0.111048767 \,\mathrm{e}^{x/22} & \mathrm{if} \quad x \le 8.93258, \\ 1 - 0.90382792 \,\mathrm{e}^{x/110} & \mathrm{if} \quad x > 8.93258, \end{cases}$
 $d_t^*(x) = \begin{cases} 10 & \mathrm{if} \quad x \le 8.93258, \\ 0 & \mathrm{if} \quad x > 8.93258. \end{cases}$

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Let the surplus process $(X_t^{(d_t)}(x))_{t\geq 0}$ follow (2) with $d_t \equiv d$ where d > 0 is constant, and the claim sizes be exponentially distributed with mean μ . Then $\varphi^{(d)}(x) = \varphi_{n+1}^{(d)}(x)$ for all $x \in [nd, (n+1)d)$, $n \geq 0$, where

$$\varphi_1^{(d)}(x) = C_{1,1} e^{x/\gamma}, \qquad \varphi_2^{(d)}(x) = \left(C_{2,1} + A_{2,0} x\right) e^{x/\gamma} + C_{2,2} e^{-x/\mu},$$

$$\begin{split} \varphi_{n+1}^{(d)}(x) &= \left(C_{n+1,1} + \sum_{i=0}^{n-1} A_{n+1,i} \; x^{i+1}\right) \mathrm{e}^{x/\gamma} \\ &+ \left(C_{n+1,2} + \sum_{i=0}^{n-2} B_{n+1,i} \; x^{i+1}\right) \mathrm{e}^{-x/\mu} \;, \quad n \ge 2. \end{split}$$

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Here

$$\begin{split} \gamma &= (1+\theta)(\mu+d), \\ C_{1,1} &= \theta/(1+\theta), \\ A_{2,0} &= -\frac{\theta}{(1+\theta)(\gamma+\mu)} e^{-d/\gamma}, \\ C_{2,1} &= \frac{\theta}{1+\theta} \left(1 + \frac{\gamma\mu + d(\gamma+\mu)}{(\gamma+\mu)^2} e^{-d/\gamma} \right) \\ C_{2,2} &= -\frac{\theta\gamma\mu}{(1+\theta)(\gamma+\mu)^2} e^{d/\mu}, \end{split}$$

and other coefficients can be found by recurrent formulas.

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Let conditions of Theorem 4 hold.

If d > 0 and x ∈ [0, min { μ(1+θ)/θ ln (1 + θd/μ(1+θ)), d }], then the correspoding survival probability is less than the classical one.
If d ∈ (0, μ(1+θ) ln(1+θ)/θ) and x is large enough, then the correspoding survival probability is greater than the classical one.

Additional assumptions

- the insurance company adjusts the deductible amount \overline{d}_t at every time $t \ge 0$ on the basis of the information available before time t, i.e. every admissible strategy $(\overline{d}_t)_{t\ge 0}$ $((\overline{d}_t)$ for brevity) of the deductible amount choice is a predictable process w.r.t. the natural filtration generated by $(N_t)_{t\ge 0}$ and $(Y_i)_{i\ge 1}$
- $0 \leq \bar{d}_t \leq \bar{d}_{\max}$, where \bar{d}_{\max} is the maximum allowed deductible amount such that $0 < F(\bar{d}_{\max}) < 1$; in particular, if $\bar{d}_t = 0$, then the deductible is not used at time t
- the safety loading $\theta > 0$ is constant

The premium intensity at time t depends on the deductible amount at this time and it is given by

$$c(\bar{d}_t) = \lambda(1+ heta) \int_{\bar{d}_t}^{+\infty} (y-\bar{d}_t) \,\mathrm{d}F(y).$$

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Let $X_t^{(d_t)}(x)$ be the surplus of the insurance company at time t provided its initial surplus is x and the strategy (\bar{d}_t) is used. Then

$$X_t^{(\bar{d}_t)}(x) = x + \int_0^t c(\bar{d}_s) \, \mathrm{d}s - \sum_{i=1}^{N_t} (Y_i - \bar{d}_{\tau_i})^+. \tag{7}$$

The ruin time under the admissible strategy (\bar{d}_t) is defined as

$$au^{(\bar{d}_t)}(x) = \inf \{ t \ge 0 \colon X_t^{(\bar{d}_t)}(x) < 0 \}.$$

The corresponding infinite-horizon survival probability is given by

$$\varphi^{(\bar{d}_t)}(x) = \mathbb{P}\big[\tau^{(\bar{d}_t)}(x) = \infty\big].$$

Our **aim** is to maximize the survival probability over all admissible strategies (\bar{d}_t) , i.e. to find

$$\varphi^*(x) = \sup_{(\bar{d}_t)} \varphi^{(\bar{d}_t)}(x),$$

and show that there exists an optimal strategy (\bar{d}_t^*) such that $\varphi^*(x) = \varphi^{(\bar{d}_t^*)}(x)$ for all $x \ge 0$.

Hamilton-Jacobi-Bellman Equation

Proposition 2

Let the surplus process $(X_t^{(\bar{d}_t)}(x))_{t\geq 0}$ follow (7) under the above assumptions. If $\varphi^*(x)$ is differentiable on \mathbb{R}_+ , then it satisfies the Hamilton-Jacobi-Bellman equation

$$\sup_{\bar{d}\in[0,\bar{d}_{\max}]} \left((1+\theta) \int_{\bar{d}}^{+\infty} (y-\bar{d}) \,\mathrm{d}F(y) \left(\varphi^*(x)\right)' + \left(F(\bar{d})-1\right) \varphi^*(x) + \int_{\bar{d}}^{x+\bar{d}} \varphi^*(x+\bar{d}-y) \,\mathrm{d}F(y) \right) = 0,$$
(8)

which is equivalent to

$$(\varphi^{*}(x))' = \inf_{\bar{d} \in [0, \bar{d}_{\max}]} \left(\frac{(1 - F(\bar{d}))\varphi^{*}(x) - \int_{\bar{d}}^{x + \bar{d}} \varphi^{*}(x + \bar{d} - y) \, \mathrm{d}F(y)}{(1 + \theta) \int_{\bar{d}}^{+\infty} (y - \bar{d}) \, \mathrm{d}F(y)} \right).$$
(9)

If the random variables Y_i , $i \ge 1$, have a p.d.f. f(y), then there exists the solution G(x) to (9) with $G(0) = \theta/(1+\theta)$, which is nondecreasing and continuously differentiable on \mathbb{R}_+ , and $\theta/(1+\theta) \le \lim_{x\to+\infty} G(x) \le 1$.

The solution G(x) to (9) that satisfies the conditions of Theorem 6 can be found as the limit of the sequence of functions $(G_n(x))_{n\geq 0}$ on \mathbb{R}_+ , where $G_0(x) = \varphi^{(0)}(x)$ is the survival probability provided that $d_t = 0$ for all $t \geq 0$, and

$$G'_{n}(x) = \inf_{\bar{d} \in [0, \, \bar{d}_{\max}]} \left(\frac{(1 - F(\bar{d})) \, G_{n-1}(x) - \int_{d}^{x + \bar{d}} \, G_{n-1}(x + \bar{d} - y) \, \mathrm{d}F(y)}{(1 + \theta) \int_{d}^{+\infty} (y - \bar{d}) \, \mathrm{d}F(y)} \right),$$

$$G_{n}(0) = \theta / (1 + \theta), \quad n \ge 1.$$
 (10)

(10)

Let the surplus process $(X_t^{(\bar{d}_t)}(x))_{t\geq 0}$ follow (7) and G(x) be the solution to (9) that satisfies the conditions of Theorem 6. Then for any $x \geq 0$ and arbitrary admissible strategy (\bar{d}_t) , we have

$$\varphi^{(\bar{d}_t)}(x) \le \frac{\mathcal{G}(x)}{\lim_{x \to +\infty} \mathcal{G}(x)},$$
(11)

and equality in (11) is attained under the strategy $(\bar{d}_t^*) = (\bar{d}_t^*(X_{t_-}^{(\bar{d}_t^*)}(x)))$, where $(\bar{d}_t^*(x))$ minimizes the right-hand side of (9), i.e.

$$\varphi^*(x) = \varphi^{(\bar{d}^*_t)}(x) = rac{\mathcal{G}(x)}{\lim_{x \to +\infty} \mathcal{G}(x)}.$$

Let the surplus process $(X_t^{(d_t)}(x))_{t\geq 0}$ follow (7) and the claim sizes be exponentially distributed. Then $\varphi^*(x) = \varphi^{(\bar{d}_t)}(x)$ for every admissible strategy (\bar{d}_t) , i.e. every admissible strategy is optimal.

Remark 4

Theorem 8 implies that we cannot increase the survival probability adjusting the deductible amount for the exponentially distributed claim sizes.

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Thank you very much for your attention!

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