BAYESIAN OUTCOME WEIGHTED LEARNING

PRÉ-JOURNÉE STATISTIQUES DU SUD 18/06/24

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OUTLINE

- Contextualization : precision medicine & individualized treatment regimes
- Method : Bayesian Outcome Weighted Learning
- Results : classification & uncertainty quantification







CONTEXTUALIZATION : PRECISION MEDICINE & INDIVIDUALIZE TREATMENT REGIMES



Precision medicine : « The right treatment for the right patient (at the right time) » [1]

- Develop models for personalized decision-making (policy) :
 - Input: patient's unique characteristics
 - **Output:** a treatment recommendation





CONTEXTUALIZATION : PRECISION MEDICINE & INDIVIDUALIZE TREATMENT REGIMES

PATIENT FEATURES $\{s_0, s_1, \dots, s_p\} \in \mathbb{S}$



- features?
 - patients features : gender, age, parentBMI, baselineBMI
 - patient response : reduction in BMI

Application example, determine which of two treatments is more suitable for losing weight according to your proper





One way to solve this problem with machine learning :

- think of it as a two-class classification problem
- with treatment as label
- administered.

Determine a policy : $\pi(A, S) = \mathbb{P}[A \in \{-1, 1\} | S = \{s_0, s_1, \dots, s_p\}]$

Tailored to align with a specified objective : $R \in \mathbb{R}$

and a weighting of individuals by response and propensity of treatments



[3] showed that maximizing the exception part in OWL is equivalent to weighted classification problems where we minimizes the objective function : $Q_{n}^{\text{OWL}}(\beta) = \frac{1}{n} \sum_{i=1}^{n} \frac{R_{i}}{A_{i}\rho + (1 - 1)^{n}}$

[2] OWL method determines an optimal ITR, π^* , by formulating the policy such that: $\pi^* \in \operatorname{argmax}_{\pi} \mathbb{E}\left[\frac{I(A = \pi(X))}{A\rho + (1 - A)/2}R\right]$

$$+ (1 - A_i)/2^{(1 - A_ih(S_i, \beta))} +$$

where $(z)_{+} = max(z,0)$ denotes the hinge loss function and $h(\cdot)$ is the ITR parameterized by β .





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where $(z)_{+} = max(z,0)$ denotes the hinge loss function and $h(\cdot)$ is the ITR parameterized by β .

[4] introduced a penalized variant of OWL that minimizes the objective function: $Q_n^{\mathsf{POWL}}(\beta) = \frac{1}{n} \sum_{i=1}^n \frac{R_i}{A_i \rho + (1 - A_i)/2} (1 - A_i h(S_i, \beta))_+ + \sum_{i=1}^p p_\lambda(|\beta_i|)$ l - **** l'j=1where $p_{\lambda}(\beta)$ is a penalty function and λ is a tuning parameter.

$$+(1 - = A_i)/2^{(1 - A_i n(S_i, p))_+}$$





- It is possible to cast SVM into Bayesian framework [4]
- Why coming back to a statical method?

 - for medical experts



Because Bayesian framework is able to capture and model uncertainty

Uncertainty quantification is a power tool in treatment recommandation





BAYESIAN FRAMEWORK

- > Parametric estimation of β from observed data
- Combines prior information about the parameters with observed data to produce a posterior distribution of the parameters
 - Prior distribution: reflects the initial knowledge or beliefs about the parameter
 - Likelihood: probability of observing the data given the parameter
 - Posterior distribution: Pior + Data ----> Posterior



Pseudo-posterior distribution :

$$p(x \mid a_i, \nu, \alpha) \propto \exp(-Q_n(\beta, \nu, \alpha))$$

$$\propto \exp\left\{\sum_{n=1}^n \frac{r_i}{a_i \rho + (1 - a_i)/2} (1 - a_i h(s_i, \beta))_+\right\} \prod_{j=1}^p p(\beta_j \mid \mu_0, \alpha_j)$$

$$\propto C(\nu, \alpha) L(\alpha \mid \beta) p(\beta \mid \mu_0, \alpha_0^2)$$

Pseudo-likelihood :

$$L_i(a_i | r_i, s_i, \beta) = \exp\left\{-2\frac{r_i}{a_i\rho + (1 - a_i)/2}\max(1 - a_i)\right\}$$

$$= I(a_i = 1) \int_0^\infty \frac{1}{\sqrt{2\pi\lambda_i}} \exp\left\{-\frac{1}{2\lambda_i} \left(\frac{r_i}{\rho}\right)\right\}$$

$$+I(a_i = -1) \int_0^\infty \frac{1}{\sqrt{2\pi\lambda_i}} \exp\left\{-\frac{1}{2\lambda_i}\right\}$$



 $a_i s_i^T \beta, 0$ $\sum_{j=1}^{n} + \lambda_i - \frac{r_i}{\rho} a_i s_i^T \beta \right)^2$ $> d\lambda_i$

 $\frac{1}{r_i} \left(\frac{r_i}{1-\rho} + \lambda_i - \frac{r_i}{1-\rho} a_i s_i^T \beta \right)^2 d\lambda_i$



Prior distribution for β

- Normal distribution prior
- **Exponential power prior distribution** :

$$p(\beta_{j} | \nu, \alpha = 1) = \int_{0}^{\infty} \phi(\beta_{j} | 0, \nu^{2} \omega_{j} \sigma_{j}^{2}) \frac{1}{2} e^{-\frac{\omega_{j}}{2}} d\omega_{j}$$

where $p(\omega_{j} | \alpha) \propto \omega_{j}^{-\frac{3}{2}} St_{\alpha/2}^{+}(\omega_{j}^{-1})$ and $St_{\alpha/2}^{+}$ is the density function $\alpha = 1, p(\omega_{j} | \alpha) \sim Exponential(2).$

Spike-and-slab prior distribution :

 $p(\beta_{j} | \gamma_{j}, \nu^{2}) = \gamma_{j} N(0, \nu^{2} \sigma_{j}^{2}) + (1 - \gamma_{j}) \delta_{0}(\beta_{j})$ where $\delta_0(\cdot)$ is the Dirac measure (point mass at 0). The prior on γ_j is given by $p(\gamma_j | \pi) = \pi^{\gamma_j}(1 - \pi)^{1 - \gamma_j}$.

n of a positive stable random variable of index $\alpha/2$. In particular, when



Box 2. Gibbs sampling algorithm for exponential power distribution prior on β Initialize λ , β and ω . Step 1: Draw $\beta^{(g+1)}|\nu, \Lambda^{(g)}, \Omega^{(g)}, \mathbf{r}, \mathbf{a}, \mathbf{x} \sim \mathcal{N}(B_2^{(g)}b_2^{(g)}, B_2^{(g)}).$ Step 2: Draw $\lambda^{-1(g+1)} | \beta^{(g+1)}, \mathbf{r}, \mathbf{a}, \mathbf{x}$ where $\lambda_i^{(g+1)}|\boldsymbol{\beta}^{(g+1)}, \nu, r_i, x_i \sim \mathbb{1}(a_i = 1)\mathcal{GI}$ $+ 1(a_i = -$ **Step 3**: Draw $\omega_j^{-1(g+1)} |\beta_j^{(g+1)}, \nu \sim \mathcal{IG}(\nu \sigma_j |\beta_j|^{-1}, 1)$ Repeat Steps 1, 2, and 3 until the chains converge.

$$\mathcal{IG}\left(\frac{1}{2}, 1, \left(\frac{r_i}{\rho}\right)^2 (1 - a_i \mathbf{x}_i^\top \boldsymbol{\beta}^{(g+1)})^2\right)$$

-1) $\mathcal{GIG}\left(\frac{1}{2}, 1, \left(\frac{r_i}{1 - \rho}\right)^2 (1 - a_i \mathbf{x}_i^\top \boldsymbol{\beta}^{(g+1)})^2\right).$



SIMULATIONS

Patient features: $X_1, ..., X_{10} \sim U([-1,1])$ **Treatment:** $A \sim U([-1,1])$ independently of the prognostic variables with $\mathbb{P}(A = 1) = 1/2$ > Patient response: $R \sim N(1 + 2X_1 + X_2 + 0.5X_3 + (X_1 + X_2)A, 1)$

True optimale value: $I(X_1 + X_2 > 0)$



SIMULATIONS RESULTS

n	OWL	Bayesian OWL Normal Prior	Bayesian OWL Exponential Power Prior	Bayesian OWL Spike and Slab
100	0.24	0.38	0.38	0.39
200	0.18	0.34	0.34	0.34
400	0.13	0.29	0.29	0.30
800	0.10	0.24	0.24	0.26



RESULTS : CLASSIFICATION AND UNCERTAINTY QUANTIFICATION

UNCERTAINTY QUANTIFICATION WITH BAYESIAN OWL EXPONENTIAL PRIOR



Figure 2: Heatmap of uncertainty quantification



CONCLUSION AND PERSPECTIVES

- We have introduced Bayesian formulation for OWL and demonstrated an approach with uncertainty quantification arXiv:2406.11573
- But only examined linear decision rules
- without variable selection



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